

Ripple effects from industry defaults

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ABSTRACT

This paper studies early default risk spillovers to small businesses. In this paper we show that default rates among small businesses are significantly higher in the presence of a default on S&P rated debt in an industry which buys their products or in the same industry. Using a new data set on S&P rated debt defaults, small businesses defaults, production process linkages and industry characteristics, we find evidence of negative wealth effects transmitted to small businesses along the production process.

Also, such a ripple effect is mitigated in loan portfolios concentrated into large and highly interconnected industries. We observe that a large number of firms in an industry serves a cushion to default risk transmission just like the wide economic ties offer some benefits of diversification.

IN 2008 THE BIG THREE: General Motors, Chrysler and Ford found themselves on a cliff of financial solvency and seeking financial support from the government. In the highly lever-

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aged and concentrated automotive industry it meant an outbreak of financial distress which propagated onto their suppliers. Just by the end of 2008 GM held off \$ 10 billion of payments to its suppliers for parts which had been delivered (Vlasic and Wayne (2008)). Resulting liquidity shortage forced multiple suppliers into default on their obligations to subcontractors and further weakened the industry's supply chain (Klein (2009)). This is an example of how major corporate credit event (which we call *industry default*) generates for linked firms negative externalities or negative wealth effects. And this affects the creditworthiness of linked firms and can trigger default clustering. On an industry level it adds up to an industry-wide change in default rates¹. Throughout this paper we will use the term *ripple effect* to describe such a response of industry default rate to industry default. The main question we ask is whether an industry default is followed by default clustering (ripple effect) in linked industries. And by linked industries we mean the industries that are linked via supply chain (through customer-supplier relationship as in Cohen and Frazzini (2008)) or via product market (competitors as in Lang and Stulz (1992)).

Our contribution to the existing literature is threefold. First, we derive our results for U.S. small businesses in manufacturing industries for which the empirical evidence for default risk spillovers is scarce. Importantly, private firms are not less susceptible to counterparty risk and liquidity shocks than their more researched corporate peers. But, in general, the measurement of default risk spillovers relies on information on individual counterparty exposures and bilateral links which in small business lending is hindered by both the prohibitive cost and tacit type of information. The information scarcity subjects even a diversified portfolio to volatility of future losses. This paper offers a plausible alternative in which counterparty exposures are modeled on an aggregate level as production process linkages. The proposed alternative feeds only on public data which can facilitate its use at a bank's risk management departments.

As a ripple effect can significantly increase losses on a loan portfolio, its measurement

¹Industry default rate measures the rate at which active and financially sound small businesses default within one year. The default event takes place if a payment is either 90 days past due or is unlikely to be paid.

is of special interest to small business finance providers. According to the FDIC, the US commercial banks' exposure to loans granted to small businesses is significant and amounts in June 2011 to 24.9% of the commercial and industrial loans. This study aims to provide insights into default risk transmission to small businesses from a new data set. The new data set used spans years 2005 to 2011 and combines information on major defaults on S&P rated debt with a panel of small businesses defaults, industry production process linkages and industry characteristics.

Second, the study also provides an original perspective on aspects of portfolio concentration and default risk transmission. In particular, we assume the perspective of small business finance providers which might be concerned with the ripple effect on their concentrated loan portfolios. We examine how magnitude of ripple effects changes with portfolio concentration into large, interconnected and highly concentrated industries. Industry size refers to the number of establishments operating in a given industry, interconnectedness corresponds to the number of bilateral connections between industries and concentration measures the degree of competition between firms and their ability to set price further from marginal cost.

Lastly, we provide novel evidence on Kiyotaki and Moore (2002) *balance sheet contagion*. We analyze here a ripple effect mechanism in which industry default propagates either directly through accounts receivables linking firms along the production process or indirectly through fluctuations in asset prices. It is important to recognize that this mechanism takes place more frequently and is set in motion much in advance of bankruptcy. So far research on default risk transmission focuses on the role of bankruptcy as the event spills default risk to linked industries. But bankruptcies are relatively rare events often anticipated and preceded by defaults, late payments, debt renegotiation and fire sales. A bankruptcy event is therefore a very late indicator of default risk spillovers. Instead, default risk can spread months prior to a bankruptcy and is often set in motion, i.e. by first payment disruption to suppliers. For example, in 2010, out of 50 defaults on S&P rated debt in the U.S. only 15 were caused by bankruptcy events (Chapter 11 filings). To this end, in the spirit of Kiyotaki and Moore (2002) we would like to verify existence of such early default risk spillovers or

ripple effects to linked industries.

We present evidence that a distress in one industry ripples to linked industries. Our results show that default rate among small businesses is higher in the presence of industry default in any industry which buys their products or in the same industry. We find that small businesses in industries with greater size (number of establishments) are subject to lower ripple effect. It means that in sizable industries the damage to an industry's credit worthiness is lower as it is measured relative to the number of establishments. Furthermore, the relationship between interconnectedness (number of bilateral inter-industry connections) and the ripple effect is negative. More inter-industry connections offer more diversification. Thus the ripple is reduced in portfolios of interconnected industries as the counterparty risk is diversified away.

Default clustering can seem to an outside observer as a result of common shocks causing otherwise heterogeneous firms to go simultaneously into financial distress.² Additionally, once initiated this aggregate behavior persists in the economy and ripples through industry sectors. Abstracting from aggregated shocks, as noticed in Horvath (2000) the alternative mechanism which organizes firms' behavior across industries comes naturally from the production process. A large share of commodities is an intermediate input to the production process of a new commodity. We use the production process setting in Figure 1 to illustrate the ripple effect. We talk about customer or supplier ripple which unfolds between two industries linked along a production process.

For example consider industry j which uses the intermediate output of industry i in its own production process of another commodity.³ In this case firms from industry j enter a supplier-customer relationship with firms from industry i which is accompanied by credit chains as in Kiyotaki and Moore (2002, 1997). Suppose an industry default occurs in indus-

²In credit risk modeling such common shocks can be found i.e. in factor models or intensity models. In particular, asymptotic single factor model in Basel II identifies one common risk factor to drive the defaults in the whole economy. Also some intensity models subject firm's default intensity to a change in macroeconomic risk factors. Alternative methods of default clustering in the literature include for example jumps in intensity models (Berndt, Ritchken, and Sun (2010)), Markov chains in which default intensities change at a default of a counterparty (Kraft and Steffensen (2007)) or frailty models in which default clustering is partially explained by an unobserved latent variable driving defaults (Duffie, Eckner, Horel, and Saita (2009)).

try j at time t . Although the involved customers and suppliers are not directly identified, the existence of the linkage along the production process between industries j and i indicates that at least some firms from i enter a direct customer-supplier relationship and are potentially exposed to distress of their counterparties in j . For them the default of firms in j translates into a shock i.e. to their accounts receivables and results in decreased firm value. Consequently the ripple results in an increase in the number of defaults in industry i .

We talk about a competitor ripple which unfolds within the same industry. In this case an industry default can have either negative or positive effect on industry competitors. First the adverse effect, called contagion effect, arises from negative information about industry profit outlooks. Suppose that firm's m investments are correlated with the investments of its competitors. If an industry default occurs due to an adverse shock to competitors investments it also signals a decrease in firm's m investment value. Second the positive effect, called competitive effect, reflects an opportunity to seize new market share that is lost by the distressed competitors, and in consequence to gain market power and to benefit from some form of monopoly (Lang and Stulz (1992)).

II EXAMPLE

In this paper we relate the network of product flows to the pattern with which default risk progresses in the economy. As an example we isolate an industry default and observe whether it is followed by another industry default in vertically linked industries. Figure 2 illustrates such development in a subset of the automotive supplier network. Indeed industry defaults follow here a pattern in which the product flow is a perfect indicator of the sequence in which industries are affected by an industry default. Starting at the top customer - motor vehicle parts - which defaults in the first quarter of 2008 for the first time in a year, the industry defaults occur next in its direct suppliers. Next in line are fabricated metal products which deliver a considerable 7% of its production to motor vehicle parts. With time the default risk ripples further to more distant suppliers as well.

Corresponding image appears in U.S. small businesses operating in those industries. Fig-

Figure 3 illustrates the time series behavior of the private firms operating in the automotive supplier network. To draw the picture in Panel (a) we pull all industries for which customers are shown in Figure 2. In the next step, we compute the small business default rates and set the time window around the distress in the customer industry. We benchmark their behavior to a sample of matched industries which resemble them in all aspects but the industry default in customer industry (the matching procedure is described in detail in section VI A). The general response of small businesses' default rates to industry default in a customer industry is an increase next quarter. Similarly, Panel (b) shows ripple effect for small businesses in industries that buy from the ones in Figure 2 and Panel (c) shows ripple effect for small businesses in the same industries as those in Figure 2. Also, in case of supplier in default the small business rates are always higher than for the matched sample and shows an increase in the default risk one quarter after the industry default in the supplying industry relative to the matched sample.

III RELATED LITERATURE

In an economy with simultaneous borrowing and lending between firms a default on one loan can significantly affect the riskiness of another. Performance of such interlocked loans co-moves with the business cycle and in turbulent times leads to default clustering. Kiyotaki and Moore (2002) discuss a theoretical framework in which local defaults of agents propagate to other sectors in the economy via accounts receivables or via similar assets used as collateral. The accounts receivable mechanism is the subject of numerous studies on the role of supply chain and credit networks in transmission of shocks. For example Raddatz (2010) or Holly and Petrella (2012) presents evidence that a customer-supplier network propagates sectoral or aggregate shocks through the economy. Yet only Wagner, Bode, and Koziol (2011) recognize the importance of market structure in default risk transmission. In their paper a distress of one supplier benefits its competitors as they gain more market power. The collateral mechanism is studied by Acharya, Bharath, and Srinivasan (2007), Benmelech and Bergman (2011) and Kiyotaki and Moore (1997). At the heart of this second mechanism rests a

devaluation of an asset class which if pledged as collateral worsens the ability of a credit-constrained firm to raise more funding and decreases its net worth. As Bernanke and Gertler (1989) point out, such unrelated shocks to a borrower's collateral value and thus its net worth can generate fluctuations in an aggregate economy.

On a portfolio level, both mechanisms of ripple effects can work simultaneously and manifest as default clustering. Empirically it is their net effect that is observed and without any knowledge of collateral prices and redeployability the ripple effect from counterparty risk is virtually undistinguishable from the ripple effect from collateral deterioration. In this paper we study the net effect of those two.

Our study is motivated by the strain of literature examining the role of market structure in the ripple effect which is observed among competing firms in the same industry. An important work by Lang and Stulz (1992) provides empirical evidence for a generally adverse stock price reaction in response to competitor's bankruptcy announcement. This pattern however is reversed for firms in highly concentrated industries with loose credit-constraints. Similar results are shown in Cheng and McDonald (1996) and Hertz et al. (2008). The latter finds significant negative effects which extend beyond the single industry and affects the suppliers and customers industries as well. In addition, a more recent study by Jorion and Zhang (2009) explores the default risk implications for the counterparties of bankrupting firm. For creditors of the distressed firm they find strong evidence of an increase in CDS spreads and a positive probability of failure in the near future. Hertz et al. (2012) discuss changes in loan conditions under which firms obtain their funding around bankruptcy announcement of their industry competitors. However the existing studies focused on the ripple effect of bankruptcies which are late events to capture the balance sheet contagion as described by Kiyotaki and Moore (2002). Also, the aspects of size and production linkages of an industry were missing from the market structure analysis, although getting considerable attention in the banking industry.

Thus, although an industry default is an important credit event, to date there is no evidence on whether or how it impacts default rates in small business loan portfolios. Instead

the existing evidence of default risk transmission is limited to outcomes of bankruptcies and only for large public firms. But as Kiyotaki and Moore (1997) notice, the effect of default risk transmission is amplified in an economy with small firms with limited access to capital markets and therefore more credit-constrained. In such an economy the entrepreneur finds herself borrowing from and lending to her suppliers even though she could be credit-constrained herself.

IV EMPIRICAL METHODOLOGY

We test the existence of ripple effect in manufacturing industries using difference-in-difference methodology. To this end we estimate variants of the following specification on industry-quarter observations which include 77 manufacturing industries in 22 quarters. The dependent variable is small business default rate which measures the rate at which financially sound, non-defaulted small businesses go into default within 1 year:

$$\begin{aligned}
 p_{i,t} = & \alpha_{Cu} Dose_{Cu,i,t} \times Post_{Cu,i,t} + \alpha_{Su} Dose_{Su,i,t} \times Post_{Su,i,t} + \\
 & + \alpha_{Co} Dose_{Co,i,t} \times Post_{Co,i,t} + \beta X_{i,t} + \sum_{i=1}^I Industry_i + \sum_{t=1}^T Q_t + \epsilon_{i,t} \quad (1)
 \end{aligned}$$

$$\text{where } Dose_{n,i,t} = \frac{\sum_{j=1}^I Debt_{n,ji,t}}{\sum_{j=1}^I Assets_{n,ji,t}}$$

Subscripts i and t denote industry and quarter respectively. The subscript n corresponds to the treatment type: Cu denotes customer ripple, Su supplier ripple and Co competitor ripple. The variable $Dose$ measures the treatment's intensity. In particular, $Dose_{Cu,i,t}$ is the total amount of debt in default on S&P rated debt in all industries j buying from industry i at time t , divided by the total assets in industries those industries. In other words, it is the amount in default in customer industries standardized by the overall size of customer industries. Similarly, $Dose_{Su,i,t}$ is the amount in default in supplier industries standardized by the overall size of those industries. $Dose_{Co,i,t}$ denotes the amount in default in the same industry standardized by its own size. $Post_{Cu}$, $Post_{Su}$ and $Post_{Co}$ are dummies that take

value of one in the quarter following an industry default in customer, supplier or in the same industry respectively. Matrix X stands for industry level controls. We also include industry and quarter fixed effects. The industry fixed effects subtract any unobserved heterogeneity on industry level. This way we control for any time invariant factors, i.e. infrastructure, supply chain base, etc. In this case the identification of ripple effects comes from the time series variation in small business default rate on an industry level. Also, the quarter fixed effects account for any aggregate comovement in the small business default rate. The variable *Dose* is absorbed by the industry fixed effects as its potential level adds up to the leverage ratio of the customer industries. The variable *Post* is absorbed by the quarter fixed effects.

We expect the interaction terms between *Dose* and *Post* to be positively and significantly associated with the small business default rate. This relationship is expected to be positive because the small business default rate should be higher following a distress in a linked industry (*Post* variable equal one). Also, the more severe is the distress in the linked industry (high level of *Dose*) the higher the small business default rate.

To illustrate the difference-in-difference approach consider the following example. Suppose we are interested in the effect of GM default in the first quarter of 2009 on small businesses default rate in the ‘Engine, turbine, and power transmission equipment manufacturing’ industry that supplies to the ‘Motor vehicle manufacturing’ industry in which GM operates. To this end we would subtract default rate after the GM default from the default rate prior to first quarter of 2009. However, the 2009 GM default overlapped with the recession period which could influence the small business default rate in the ‘Engine, turbine, and power transmission equipment manufacturing’ industry. Therefore, benchmarking the outcome against a ‘control’ industry, i.e. ‘Metalworking machinery manufacturing’, that was not treated by any customer ripple at that time helps to control for general business conditions. In essence, the difference-in-difference approach compares the difference in default rate in ‘Engine, turbine, and power transmission equipment manufacturing’ industry pre and post GM default to the difference in ‘Metalworking machinery manufacturing’ industry pre and post GM default (see also Bertrand and Mullainathan (2003) for other examples

of difference-in-difference approach). Our regression differs from the above example with respect to the fact that we allow more severe industry defaults to be followed by even higher increases in small business default rate.

Since the industry defaults are staggered over time, the regression in (1) will set as ‘control’ the industries that at a given time are not treated by the specific ripple type. The control industries however may include industries that were or will be under ripple effect. In fact all manufacturing industries are treated by a ripple effect at some point in time. Also, if we are interested in the customer ripple, the control industries can be under supplier or competitor ripple. Similar logic applies to supplier and competitor ripple.

To determine the role of industry characteristics in the ripple effect we estimate the following regression:

$$\begin{aligned}
p_{i,t} = & \sum_{n=Cu,Su,Co} \gamma_n Dose_{n,i,t} \times Post_{n,i,t} \times Feature_{i,t} + \sum_{n=Cu,Su,Co} \alpha_n Dose_{n,i,t} \times Post_{n,i,t} \\
& + \sum_{n=Cu,Su,Co} \theta_n Post_{n,i,t} \times Feature_{i,t} + \beta X_{i,t} + \sum_{i=1}^I Industry_i + \sum_{t=1}^T Q_t + \epsilon_{i,t}
\end{aligned} \tag{2}$$

where the *Feature* stands for an industry characteristic of interest, i.e. size which is the number of establishments in an industry, interconnectedness which is the number of overall connections to suppliers and customers and concentration which is the industry markup. The subscript n corresponds to the treatment type: *Cu* denotes customer ripple, *Su* supplier ripple and *Co* competitor ripple. An industry which is smaller, less interconnected and less concentrated is expected to suffer higher ripple effects. The interaction term $Dose \times Feature$ is absorbed by the *Feature* variable.

V DATA

The data is on quarterly frequency with information available on industry level. We are interested to measure the ripple effect for U.S. small businesses in 77 manufacturing industries

in 22 quarters from 2005 q3 to 2010 q4 which amounts to total of 1,694 observations.

A Dependent variable

We adopt the Basel Accords view to compute the small business default rate. It means that default event takes place if a payment occurs either 90 days past due or is unlikely to be paid, i.e. bankruptcy or a credit rating downgrade to default. Here, the small business default rate is a cumulative number and represents a share of financially sound firms that go into default at any point in time within 1 year. In particular, at time t we identify a group of firms in non-defaulted state. We track them over the next four quarters to see if they go into default at any point in time. Then the default rate is the sum of those defaults over the initial number of firms. We repeat this procedure for each quarter.

To that end we conduct an extensive analysis of nearly 240,000 U.S. small businesses per quarter from a new data set provided by Dun & Bradstreet. The data set covers rich quarterly information on firms' actual borrowing and payment behavior, i.e. number and amount of late payments. In addition each record contains information on credit ratings, County Court Judgments, legal pre-failure events, legal form, age, industry or location. The data set spans period from 2005 q2 to 2011 q4⁴ during which the study looks a representative blend of U.S. industries, regions and firm sizes (for more detail on small businesses sample please refer to Bams, Pisa, and Wolff (2012)). The D&B data on small business payment behavior is collected from about 6,000 major firms (both financial and nonfinancial). Table I Panel A summarizes the final sample of U.S. manufacturing small businesses which are exposed to ripple effect from industry defaults. It shows that the number of small businesses per industry ranges from 10 to 37,650.

⁴Sample is limited by the data provided by Dun & Bradstreet. As the computation of small business default rate requires forward looking information on four quarters ahead, we are able to compute small business default rate only up to 2010 q4.

B Independent variables

The independent variables of interest are *Dose* and *Post*. There are three types of treatment, thus there are three variants of Dose and Post variables. An industry can be treated by customer (*Cu*), supplier (*Su*) or competitor (*Co*) ripple. The first variable of interest, *Dose*, measures how severe is any of the three treatments. In particular, $Dose_{Cu,i,t}$ denotes the debt in default in customer industries standardized by the total assets of industries in which the industry default occurred; $Dose_{Su,i,t}$ denotes the intensity of industry default in supplier industries; and $Dose_{Co,i,t}$ denotes the intensity of industry default in its own industry. We standardize it by the total assets of firms in Compustat sample that operate in that industry.

The *Dose* of the treatment is derived from the information on major defaults on S&P rated debt. We focus on major defaults since they have a greater ability to stimulate an industry-wide response (see also Lang and Stulz (1992)). There are at least two reasons to assume it. First, the damage to the existing production relationships increases with size of the default. As a result, a larger number of suppliers is affected and suffers more extensive shock to their accounts receivables. Second, a major default can reveal negative information about the industry competitors if their investments are correlated with the investments of the defaulting firm. This in turn indicates that the industry is in imminent distress. Consequently uninformed customers reduce their demand for intermediate goods and alter industry's credit worthiness.

The major defaults are collected from the public information on U.S. industry defaults from 'Annual Global Corporate Default Study and Rating Transitions' provided by S&P. The data span 2005 q2 to 2010 q4 and include the company name, date and amount on which the default occurred.⁵ Next the industry default data are supplemented by industry classification codes from Thompson One Banker or EDGAR. Subsequently out of 399 defaults on S&P

⁵It is important to note that the industry default comes from a sample that is different from the one containing small businesses for which we compute small business default rate. Since such an industry default is a 'major' default on S&P rated debt, it does not enter the sample of small businesses that we test for presence of ripple effects. In general if a competitor is in distress at time t , it is excluded from the cohort of financially sound firms in non-defaulted state which comprise the base of default rate computation. Therefore the default of the competitor at time t does not have any effect on default rate at the same time but rather it is included in the $t - 1$.

rated debt we retain in our sample 340 which could be matched to a primary NAICS or if unavailable to a primary SIC industry. The coverage of the major defaults is presented in Table I Panel B. If there are multiple major defaults in one industry at the same time, we count it as a one industry default. So there are 255 unique industry defaults of total value \$894,475 million and never less than 2 industry defaults per quarter. Panel A of Figure 4 illustrates the evolution of the industry default events in the final sample of industry defaults with defaults occurring most frequently in 2009 q2 (57 industry defaults) and highest amount in default in 2009 q1 (mil \$12,572.60).

The second independent variable, *Post*, is a dummy variable that takes value of one in quarter following industry default in a linked industry. In principle, any manufacturing industry can be treated by customer, supplier or competitor ripple or by any combination of those. The number of manufacturing industries treated by either type of ripple effect is presented in Panel B of Figure 2. At any point in time there are some industries that are treated by any of the ripples. Panel C of Figure 2 shows the number of industries that are not under any type of treatment. Those industries serve as a control group, i.e. to a pure customer ripple, pure supplier ripple and pure competitor ripple. The number is lowest during the recession when there are only 4 industries not under any treatment. On average during the entire sample period there are 32 industries not under any treatment. Also, all industries are treated at some point in time.

The linkages between industries are determined based on the Make and Use tables of industry Input-Output (IO) accounts which contain the flows of intermediate inputs in the economy. The IO data are provided by U.S. Bureau of Labor Statistics on an annual basis for years 1993-2010 and are derived from the U.S. Bureau of Economic Analysis.⁶ We assume industries are linked if proportion of outputs supplied to or proportion of intermediate products purchased from a given industry is greater than 1% (for more detailed description please refer to Appendix).

⁶The most recent release of detailed IO tables by U.S. Bureau of Economic Analysis dates back to 2002. However our sample covers 2005 q2 to 2011 q4.

C Industry features

We are interested to see whether concentration of small business loan portfolios into large, interconnected or concentrated industries affects the magnitude of ripple effect. To measure the industry size we take the number of establishments from the U.S. Census Bureau County Business Patterns. The annual information is derived from the Census Bureau’s Business Register which is the most comprehensive data set on U.S. business activities. Establishments are defined as single physical locations thus larger firms tend to have more establishments. We aggregate the data into IO industries following the mapping described in Appendix.

Second, the interconnectedness of an industry is computed from U.S. Bureau of Labor Statistics IO data. It is calculated as a sum of all existing inter-industry input-output relationships with IO industries of a value greater than 1%.

Lastly, we measure industry concentration by industry markup which is the price-cost margin in an industry. Industrial organization theory predicts a positive relationship between industry concentration and industry markup. In particular, more concentrated industries are expected to have lesser competition and can set price further from marginal cost. We follow methodology by Allayannis and Ihrig (2001) and Ali, Klasa, and Yeung (2009) and calculate the price-cost margin as:

$$\text{PCM} = \frac{\text{Value of sales} + \Delta\text{Inventories} - \text{Payroll} - \text{Cost of materials}}{\text{Value of sales} + \Delta\text{Inventories}} \quad (3)$$

Given the U.S. Census definition of value added it is equal to $(\text{Value added} - \text{Payroll}) / (\text{Value added} + \text{Cost of materials})$. The annual data used to calculate this measure comes from the U.S. Census Bureau Annual Survey of Manufactures.⁷

D Controls

We collect industry level controls that include and indicator if an industry experiences a major default within one year. We derive it from the ‘Annual Global Corporate Default

⁷We aggregate the data items per IO industry following the NAICS and IO mapping discussed in Appendix.

Study and Rating Transitions' provided by S&P. The share of large firms that employ more than 100 people, share of young firms that are less than 3 years old, and median D&B credit score (CPOINTS)⁸ are also expected to play a role in the small business default rates. For example, young firms have high mortality rate and can be more sensitive to changing business environment. We compute it from the D&B data set. In our analysis we also include an indicator whether the industry has only one customer and an indicator whether the industry has only one supplier. In general, such focused industries are expected to have higher default rates. This information comes from the IO tables. Additionally, to control for demand and supply shocks we include industry's sales and inventories. This information is provided by U.S. Census Bureau Annual Survey of Manufactures.

VI Main Results

In this section particular interest is paid to evolution of industry defaults along the production process and the response they cause in default rates among small businesses. We use the term ripple effect to describe this reaction in small business credit worthiness. This paper distinguishes three types of ripple effects: a customer, supplier and competitor ripple. The last one affects competitors in the defaulting industry. The obtained results are shown to have risk management application in portfolios of loans to small businesses.

A Ripple effects in industry default rates among small manufacturing firms

Our main results are presented in Table II. It shows that default rates among small businesses are significantly higher in quarter following a major default in an industry which buys their products or in the same industry. As expected, the coefficients on the difference-in-difference terms: $Dose_{Cu} \times Post_{Cu}$ and $Dose_{Co} \times Post_{Co}$ are positive and significant. So the more severe the treatment as measured by the amount in default relative to industry's assets, the greater the damage to the small business creditworthiness. The effect is also economically

⁸The credit score predicts firm's likelihood of becoming delinquent during the next one year period. In its computation D&B takes into account payments 90 days past due, relief from creditors or payments not in full. It ranges from 100 to 670 assigning likelihood of delinquency between 2.10-61.50% respectively.

significant and in case of distress in customer industry a one standard deviation increase in the $Dose_{Cu}$ is followed by a 9.8 basis point increase in small business default rate in the supplying industries (regression in column (1)). The effect is even greater if industry default occurs in the same industry. In this case one standard deviation increase in $Dose_{Co}$ is followed by a 12.5 basis points increase in small business default rates in the same industry.

We perform five regressions. In the first regression we include only the difference-in-difference terms together with time and industry fixed effects. In column (2) to (5) of Table II we control if an industry experiences a major default within one year, or for industry's share of large firms, share of young firms and the median credit score. Intuitively industries linked along the production process may share some commonalities which make them sensitive to common shocks. A systematic shock to a group of industries should then be reflected in those controls. For example young firms are among the first to default as they are vulnerable due to new client base and little capital buffer to withstand losses. Also, credit risk measured as median credit score can reflect a common shock if it alters the firms credit worthiness.

Also, holding considerable inventories can work as a cushion in an event of a failure of a supplier. Although failure of a supplier is associated with losses on advanced payments, holding inventories minimizes disruption to the production process and allows firms to continue their production. From this point of view industries with low inventories are more vulnerable to distress in their supplier industry as their production can stop upon a supplier default. This in turn leads to higher volatility of default rates. Apart from the above supply side shock, a common shock can come from demand side i.e. as drop in sales. However, our results are robust to the inclusion of the sales and inventories⁹ variables in column (3). The ripple effect remains valid even 2 quarters after but only for the competitor ripple as is shown in column (4).

Although the potential level of $Dose$ is captured by the industry fixed effects, we would like to address any concern that $Dose$ might not be constant throughout the time by time varying (annual) industry fixed effects. This captures a non-trivial part of the data variability.

⁹Variable sales and inventories are taken from U.S. Census Bureau Annual Survey of Manufactures.

In this case the identification of the ripple effect comes from the intra-annual variability in small business default rates. Results with the time varying industry fixed effects are presented in the column (5) in Table II and the basic message remains unaffected.

Therefore even though industry default is not totally unanticipated, it serves as an indicator about the severity of financial distress in the defaulting firm. Prior to an industry default if the major firm experiences liquidity shortage, it consequently renegotiates or delays its liabilities, i.e. by postponing payments to its suppliers or delivery to its customers. So although a default on S&P rated debt does not affect the small private firms by itself since their direct exposure to this type of debt is rather limited, one have to bear in mind that industry default is merely an indicator of a process which takes place prior to it. In particular credit chains which form along the production to order of most intermediate goods are especially vulnerable to this process. By default this production takes time and the output is client-specific and can be finalized only by the specific supplier. The payment typically cannot be simultaneous with the production process but instead first part is paid up-front to secure the supplier's interests and the rest at the completion to secure the customer's interests. The second payment is therefore a debt repayment and is subject to credit risk (Kiyotaki and Moore (2002)). Also the industry default indicates that the industry is in imminent distress such that a larger number of small businesses can be affected.

As during the recession most industries were under at least one type of treatment, the OLS regression might be sensitive to criticism that it is comparing the pre-recession industries to the ones during recession. To address this issue we use the matching estimation approach in which we focus on the pre-recession (pre December 2007) period. This way we also want to alleviate a concern that our results can be driven by the credit crunch that happened during the recession. A credit crunch can force more small businesses to default on their payments as they are unable to roll over their credit. Although in OLS estimations a credit crunch should be captured by the quarterly fixed effects, the matching estimation approach is a robust, non-parametric approach that can address both concerns.

Ideally, we would like to compare the default rates in an industry under treatment (ripple

effect) to default rates in the same industry had it not received the treatment (ripple effect). Because we are unable to observe the counterfactual we aim to approximate it by another industry that mimics the treated industry in all aspects but the fact that it is treated by some given ripple effect (for discussion on application of matching estimator refer to Malmendier and Tate (2009)). To this end we use the Abadie and Imbens (2007) matching estimation approach and we match exactly on the same quarter industries which were treated by a given ripple to those that were not treated by the same ripple. Matched industries are chosen from all non-treated industries in the same quarter such that they are the closest match based on: all the control we used in regressions in Table II. Additionally, for customer ripple the match is done on: *Dose* during supplier treatment, *Dose* under competitor treatment; for supplier ripple on: *Dose* during customer treatment, *Dose* under competitor treatment; and for competitor ripple on: *Dose* during customer treatment, *Dose* during supplier treatment. By doing this we want to capture the incremental difference in default rates that is due to the specific ripple. So imagine an industry treated by customer and supplier ripple. To measure the customer ripple, it will be matched to another industry that should resemble it in all dimensions but the customer ripple. So, the matched industry should not be treated by customer but should be treated by a supplier ripple.

Table III presents the descriptive statistics of treated industries side by side with the non-treated and matched industries. The difference between treated industries and the entire sample is reported in columns (R-A) and the difference between treated industries and the matched sample is reported in column (R-M). We test if those differences are equal to zero. Among the variables used in the regressions and in the matching, three are significantly different at 1% level between industries treated by customer ripple and all those that are not treated by customer ripple, but only one between the treated and matched. Similarly, Panel B shows that three variables are significantly different at 1% level between industries treated by supplier ripple and all those that are not treated by supplier ripple, but only one between the treated and matched. In case of competitor ripple the treated industries differ with respect to one variable at 1% level of significance while not showing any significant

difference to the matched sample.

The principal ripple effect for U.S. manufacturing small businesses is measured by the average treatment effect in Table IV. The average treatment effect is positive indicating higher default rates among small businesses following an industry default in a linked industry. Economically, one quarter after industry default the difference in default rates between treated and matched industries ranges from 0.57% to 2.62%, depending on ripple type. The most pronounced difference is observed in the pre-recession period after distress in same industry. As expected a distress in a linked industry translates into a significant negative welfare effect for small businesses. It significantly reduces small businesses' credit worthiness. Overall, the production relationships are a strong channel through which negative welfare effects spread and weaken the performance of production partners. Figure 5 illustrates the development of default rates in industries treated by the ripple effect and the matched sample. In most cases, the default rates among small businesses respond by increasing right after the industry default and then tend to converge to the matched sample 3 quarters after.

B Ripple effect and market structure

We continue our analysis in Table V by exploring the role of portfolio concentration on the magnitude of ripple effects. We investigate if large, more interconnected and more concentrated industries are more vulnerable to treatment by ripple effect. We expect that small businesses in industries with greater size (number of establishments) are subject to lower ripple effect. It means that in sizable industries the damage to industry's credit worthiness is lower as it is measured relative to the number of establishments. The damage is therefore contained to a smaller share of firms that suffer a shock to their firm value. This in turn decreases the ripple effect in large industries.

Also, we anticipate a non-linear relationship between ripple effect and interconnectedness. We expect lower ripple effect for more interconnected industries that have greater number of bilateral connections between industries. The more interconnected is the industry the more diverse the economic activity. This potentially allows for diversification of the counterparty

risk. On the other hand, the more interconnected industries are exposed to shocks of various origins. Therefore industries with wide connections serve as a hub transmitting default risk; they become easier infected and at the same time infect its counterparties.

We expect lesser ripple effect in concentrated industries as the firms can have an opportunity to seize new market share that is lost by the distressed competitor. In consequence they are able to gain market power and to benefit from some form of monopoly Lang and Stulz (1992). In sum, the ripple effect is expected to be stronger in small and isolated industries with low concentration.

Table V show the result of regression (2). We include here an interaction term between the difference-in-difference terms and the industry feature as size, interconnectedness and concentration. Column (1) of Table V shows no significant relationship between ripple effect and industry's size. Next, to account for an anticipated nonlinear relationship between interconnectedness and the ripple effect we include an additional interaction term with interconnectedness squared. Column (2) presents the ripple effect for more interconnected industries. We observe no straightforward effect of the interconnectedness on the magnitude of ripple effect, although in general more interconnected industries enjoy lower default rates which reveals some diversification benefits. Importantly, the last column confirms that firms in highly concentrated industries can benefit from distress of other firms. In this case they are able to, i.e. step in and take over market share following a distress in customer industry (positive and significant coefficient on $Feature \times Dose_{Cu} \times Post_{Cu}$). Holding a portfolio of loans to small businesses operating in concentrated industries helps to mitigate counterparty risk and therefore ripple effect. The results show that an average industry (with respect to concentration) experiences a 0.005 basis point increase in default rate following a treatment by an average $Dose_{Cu}$. However, a one standard deviation more concentrated industry actually benefits of an industry default in its customer industry. In this case, we observe a decrease in small business default rates by 1.6 basis points.

In the matching estimator approach we construct three portfolios containing largest, most interconnected or most concentrated industries from the top quintile. We report the resulting

average treatment effects in Table VI. Panel A depicts the results for the pre-recession period and Panel B for the full sample period. In the full sample period and partially before the recession, larger industries treated by the ripple effect respond with a lower ripple effect than the matched sample. Thus the larger is the industry, the lower the relative damage to the production relationships. The damage is contained to a smaller share of firms that suffer a shock. Next, column (2) presents the ripple effect for interconnectedness portfolios. A trend is observed in which the industries with wide connections suffer lower ripple effect than the matched sample. Thus, our results suggest that there are diversification benefits in the more interconnected industries. Column (3) of Table VI shows evidence that during full sample period the ripple effect lessens in highly concentrated industries. It is in line with previous research that reports a positive effect from default in concentrated industries. This pattern is however reversed prior to the recession and can suggest that during this particular time period firms actually experienced contagion in default risk rather than competitive advantages.

C Ripple effect and portfolio loss implication

How does concentration into large, interconnected or concentrated industries relate to the counterparty risk and the ripple effect in portfolios of loans to small businesses? To answer this question we bootstrap small business portfolios from historical data. Each portfolio contains small businesses distributed across 77 manufacturing IO industries proportionally to the historical data. To find the impact of ripple effect on portfolio default distribution we consider two scenarios: one without any ripple effect and second one with single ripple effect.

First, we create the unconditional loss distribution in which we ignore the existence of ripple effect. The unconditional loss distribution is bootstrapped from the historical data in the following way: we randomly draw a quarter for each industry and take the number of defaults and total number of firms that were in that industry at that random quarter. Second, for the distribution with ripple effect we first randomly select a single industry default

from historical data. Then we define the treated industries as linked industries (suppliers, customers or the same industry). For them we take the number of defaults and total number of firms that were in those industries for the following quarter. For the non-treated industries we repeat what we did for the unconditional distribution. We repeat that procedure 100,000 times to obtain distribution presented in Figure 6.

Panel (a) of Figure 6 shows ripple effect from a single industry default for a diversified portfolio including all 77 manufacturing industries. The dashed line shows the distribution of defaults for the unconditional bootstrapping without ripple effects. The results show that in a portfolio without ripple effect the 99.9th percentile of defaults is at 18.17%. In other words, based on this distribution the probability that more than 18.17% firms will default within the next year is less than 0.1%. This type of information can be used in determining the capital requirements or tranching of a portfolio.

This ripple effect has a substantial implication for the portfolio default distribution as shown in Panel (a) of Figure 6. The solid line depicts now the portfolio default distribution with a single ripple effect. Ripple effect from a single industry default shifts the density of the portfolio default distribution to the right and moves some of the mass to the right tail. It is a consequence of increased expected losses and default correlation. The 99.9th percentile of the default distribution increased from 18.17% to 18.21% (which is 4 basis points) after a single industry default. We find that ignoring the ripple effect might lead to an understatement of the portfolio credit risk and thus the required capital.

This non-parametric approach should shed some more light on whether portfolio concentration into large, interconnected and concentrated industries reduces the counterparty risk and ripple effect. Panels (b) to (d) of Figure 6 show default distributions for different sub portfolios which are concentrated in the 20% of larges for (b), most interconnected for (c), and most concentrated industries for (d). Apart from Panel (b), ripple effect is always present, but its magnitude is always smaller than in the diversified portfolio. This gives some scope for risk management in such portfolios.

VII CONCLUDING REMARKS

In this paper we draw the attention to a default risk transmission along the production process. Using a new data set containing information on major defaults on S&P rated debt, small businesses defaults, production process linkages and industry characteristics we present evidence that a distress in one industry ripples to small businesses in linked industries. Our results show that small businesses in industries exposed to distress through product flow experience significant negative wealth effects and suffer higher default risk. We claim that industries linked either by production process or by product market participate in a ripple effect initiated by one of their counterparties.

We derive our results for U.S. small businesses for which the empirical evidence for default risk transmission is scarce. Importantly, private firms are not less vulnerable to counterparty risk and liquidity shocks than their more researched corporate peers. But in general the measurement of default risk transmission relies on information on individual counterparty exposures which in small business lending is hindered by the prohibitive cost of information. This paper offers a plausible alternative in which counterparty exposures are modeled as production process linkages. The proposed alternative feeds only on public data.

We find evidence that ripple effect is hindered in more concentrated industries. There the competitive effect plays a dominant role since the firms are able to benefit from distress of counterparty. Also, we find that small businesses in large industries (measured by number of establishments) are subject to lower ripple effect. The damage is therefore contained to a smaller share of the industry that suffers the shock. In other words, relatively fewer firms suffer a hit to their asset value. Moreover, the relationship between interconnectedness (number of bilateral industry connections) and the ripple effect is negative. We observe that wide economic ties offer some benefits of diversification. Thus the ripple loses strength as the counterparty risk is slowly diversified away.

Appendix A

The IO data cover commodity flows for 195 IO industries. We recode the firm NAICS and SIC codes into one of the 195 IO industries using concordance tables between IO and 2007 NAICS provided by the U.S. Bureau of Labor Statistics. Moreover, the concordance tables between 2007 NAICS, 2002 NAICS and SIC are provided by the U.S. Bureau of Economic Analysis. Our analysis focuses on 77 IO manufacturing industries. In few cases the procedure maps one SIC into few IO industries. In this case we follow Ahern and Harford (2014) and assign a firm from that SIC industry into one of those IO industries at random. It allows us to preserve the behavior of firms in the aggregate in one IO industry while matching the firms to a single IO industry.

To identify the supplier-customer pairs we construct matrixes with commodity flows from the annual Make and Use tables. Following Ahern and Harford (2014) the commodity output matrix $SHARE_{IxK}$ is derived from the make table M_{IxK} and records the proportion of an industry i in production of a commodity k . On the other hand, the u_{ki} element of a use matrix U_{KxI} gives the dollar amount of commodity k used as an intermediate input in production process of industry i . In the next step, the $REVSHARE_{IxI}$ is an industry-by-industry matrix which records the dollar flow from the user industries in columns to the producer industries in rows:

$$REVSHARE = SHARE \times U \tag{A1}$$

Next, the customers' matrix $CUST_{IxI}$ is derived as a proportion of intermediate products produced and supplied by a row industry to its customers. It specifies how much of the outputs of the production process is supplied to a given customer. Analogously, the suppliers' matrix $SUPP_{IxI}$ records the proportion of intermediate products purchased and used by the column industry from its suppliers. In other words it indicates how much of the inputs to the production process comes from a given supplier. A relationship is identified as a customer or supplier relationship if entries of $CUST$ or $SUPP$ are greater than 1%.

References

- Abadie, Alberto, and Guido W. Imbens, 2007, Bias corrected matching estimators for average treatment effects, mimeo Harvard University.
- Acharya, Viral V., Sreedhar T. Bharath, and Anand Srinivasan, 2007, Does industry-wide distress affect defaulted firms? Evidence from creditor recoveries, *Journal of Financial Economics* 85, 787–821.
- Ahern, Kenneth R., and Jarrad Harford, 2014, The importance of industry links in merger waves, *The Journal of Finance* 69, 527–576.
- Ali, Ashiq, Sandy Klasa, and Eric Yeung, 2009, The limitations of industry concentration measures constructed with compustat data: Implications for finance research, *The Review of Financial Studies* 22, 3839–3871.
- Allayannis, George, and Jane Ihrig, 2001, Exposure and markups, *The Review of Financial Studies* 14, 805–835.
- Bams, Dennis, Magdalena Pisa, and Christian Wolff, 2012, Modeling default correlation in a U.S. retail portfolio, CEPR Working Paper No. 9205.
- Benmelech, Efraim, and Nittai K. Bergman, 2011, Bankruptcy and the collateral channel, *The Journal of Finance* 66, 2061–2084.
- Bernanke, Ben, and Mark Gertler, 1989, Agency costs, net worth, and business fluctuations, *The American Economic Review* 79, 14–31.
- Berndt, Antje, Peter Ritchken, and Zhiqiang Sun, 2010, On correlation and default clustering in credit markets, *The Review of Financial Studies* 23, 2680–2729.
- Bertrand, Marianne, and Sendhil Mullainathan, 2003, Enjoying the quiet life? corporate governance and managerial preferences, *Journal of Political Economy* 111, 1043–1075.
- Cheng, Louis T.W., and James E. McDonald, 1996, Industry structure and ripple effects of bankruptcy announcements, *Financial Review* 31, 783–807.
- Cohen, Lauren, and Andera Frazzini, 2008, Economic links and predictable returns, *The Journal of Finance* 63, 1977–2011.
- Duffie, Darrell, Andreas Eckner, Guillaume Horel, and Leandro Saita, 2009, Frailty correlated default, *The Journal of Finance* 64, 2089–2123.
- Hertzel, Michael G., Zhi Li, Micah S. Officer, and Kimberly J. Rodgers, 2008, Inter-firm linkages and the wealth effects of financial distress along the supply chain, *Journal of Financial Economics* 87, 374–387.
- Hertzel, Michael G., and Micah S. Officer, 2012, Industry contagion in loan spreads, *Journal of Financial Economics* 103, 493–506.
- Holly, Sean, and Ivan Petrella, 2012, Factor demand linkages, technology shocks, and the business cycle, *The Review of Economics and Statistics* 94, 948–963.
- Horvath, Michael, 2000, Sectoral shocks and aggregate fluctuations, *Journal of Monetary Economics* 45, 69–106.
- Jorion, Philippe, and Gaiyan Zhang, 2009, Credit contagion from counterparty risk, *The Journal of Finance* 64, 2053–2087.
- Kiyotaki, Nobuhiro, and John Moore, 1997, Credit cycles, *Journal of Political Economy* 105, 211–248.

- , 2002, Balance-sheet contagion, *The American Economic Review* 92, 46–50.
- Klein, Karen E., 2009, Survival advice for auto parts suppliers, *Businessweek*, June 16.
- Kraft, Holger, and Mogens Steffensen, 2007, Bankruptcy, counterparty risk, and contagion, *Review of Finance* 11, 209–252.
- Lang, Larry H. P., and René M. Stulz, 1992, Contagion and competitive intra-industry effects of bankruptcy announcements. An empirical analysis, *Journal of Financial Economics* 32, 45–60.
- Malmendier, Ulrike, and Geoffrey Tate, 2009, Superstar ceos, *The Quarterly Journal of Economics* 124, 1593–1638.
- Raddatz, Claudio, 2010, Credit chains and sectoral comovement: Does the use of trade credit amplify sectoral shocks?, *The Review of Economics and Statistics* 92, 985–1003.
- Vlasic, Bill, and Leslie Wayne, 2008, Auto suppliers share in the anxiety, *The New York Times*, December 12.
- Wagner, Stephan M., Christoph Bode, and Philipp Koziol, 2011, Negative default dependence in supplier networks, *International Journal of Production Economics* 134, 398–406.

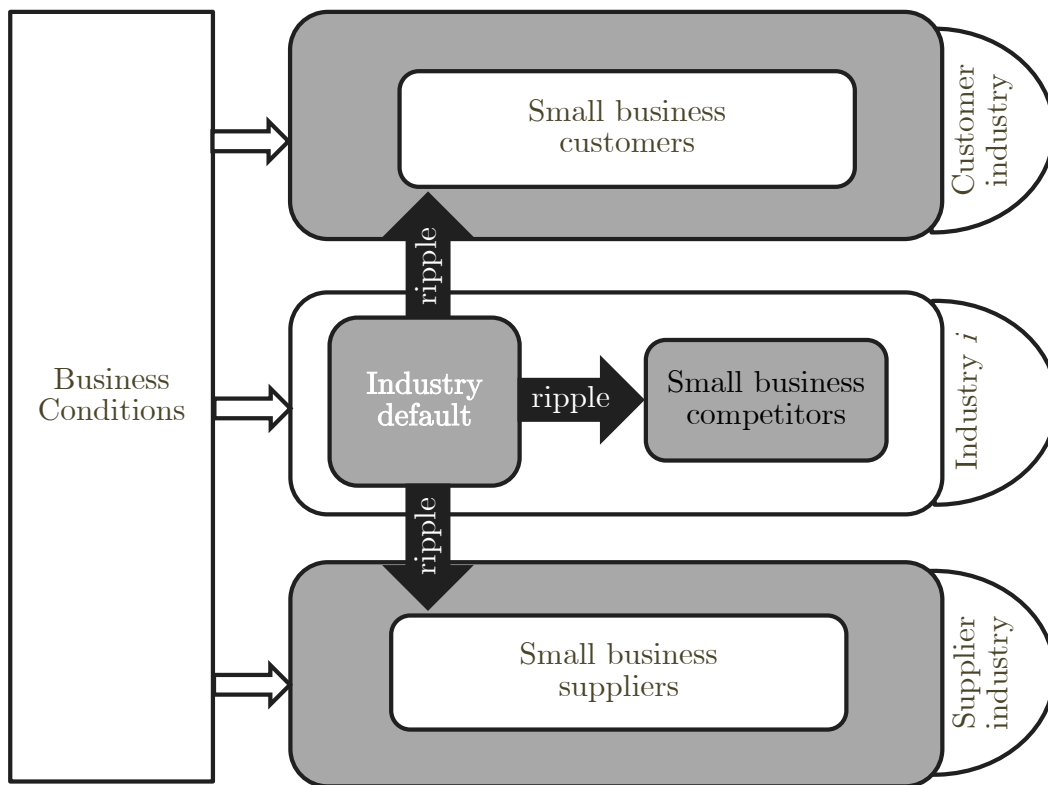


Figure 1: Customer, supplier and competitor ripple effects from an industry default. Industry i awaits intermediate inputs from supplier industry and owes the customer industry to complete products but suffers an industry default.

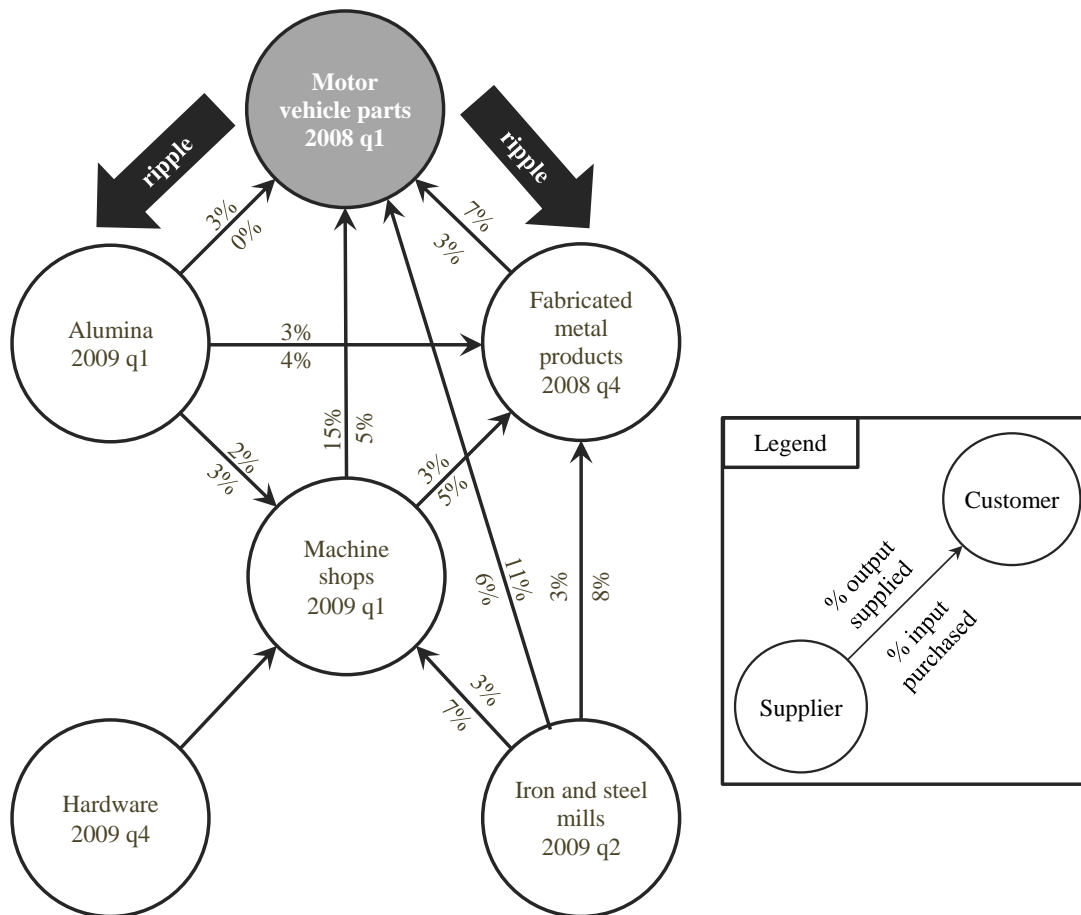


Figure 2: Subset of the automotive supplier network and industry default. The figure presents a supplier network given by the U.S. Bureau of Labor Statistics Input Output tables. The arrows indicate product flows. The quarters in the circles denote quarters in which first industry default occurred as of 2007 q1.

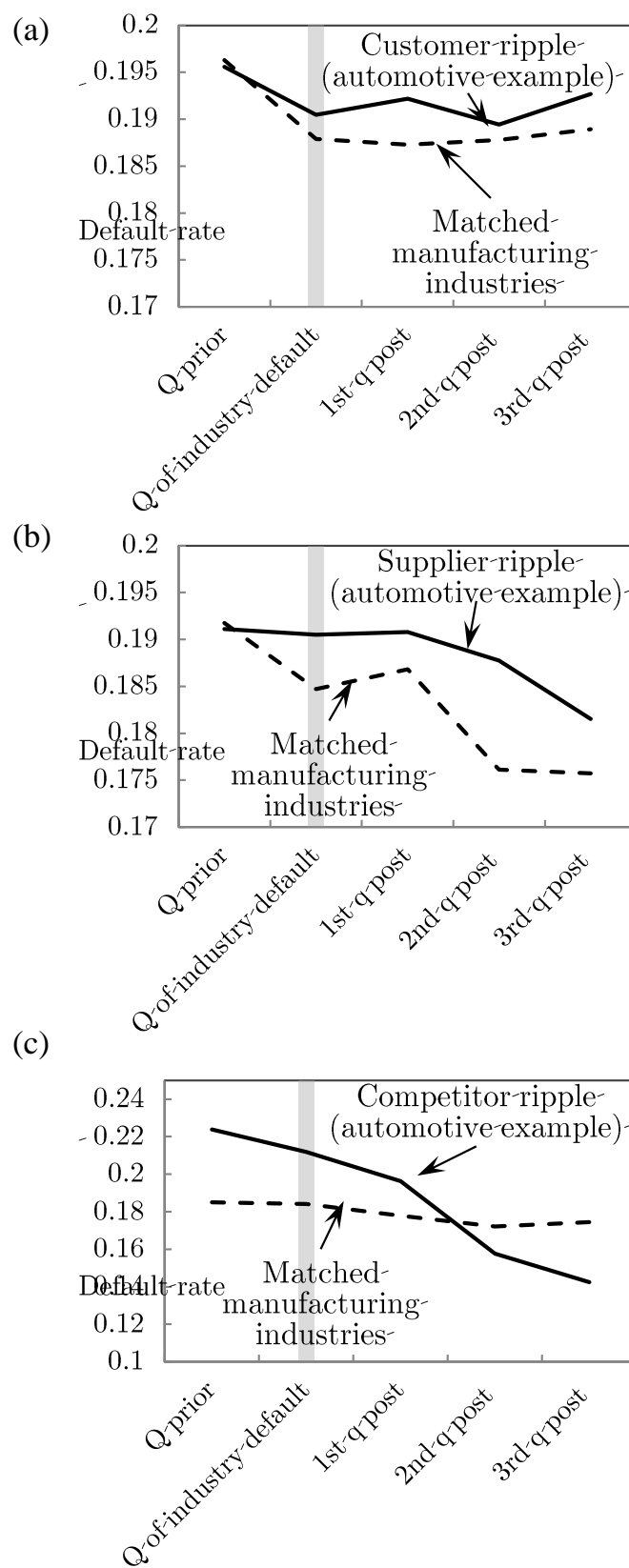


Figure 3: Default rates among small firms in the automotive supplier network. The figure presents default rates in U.S. small firms in the industries related to automotive industries around industry defaults displayed in Figure 3.

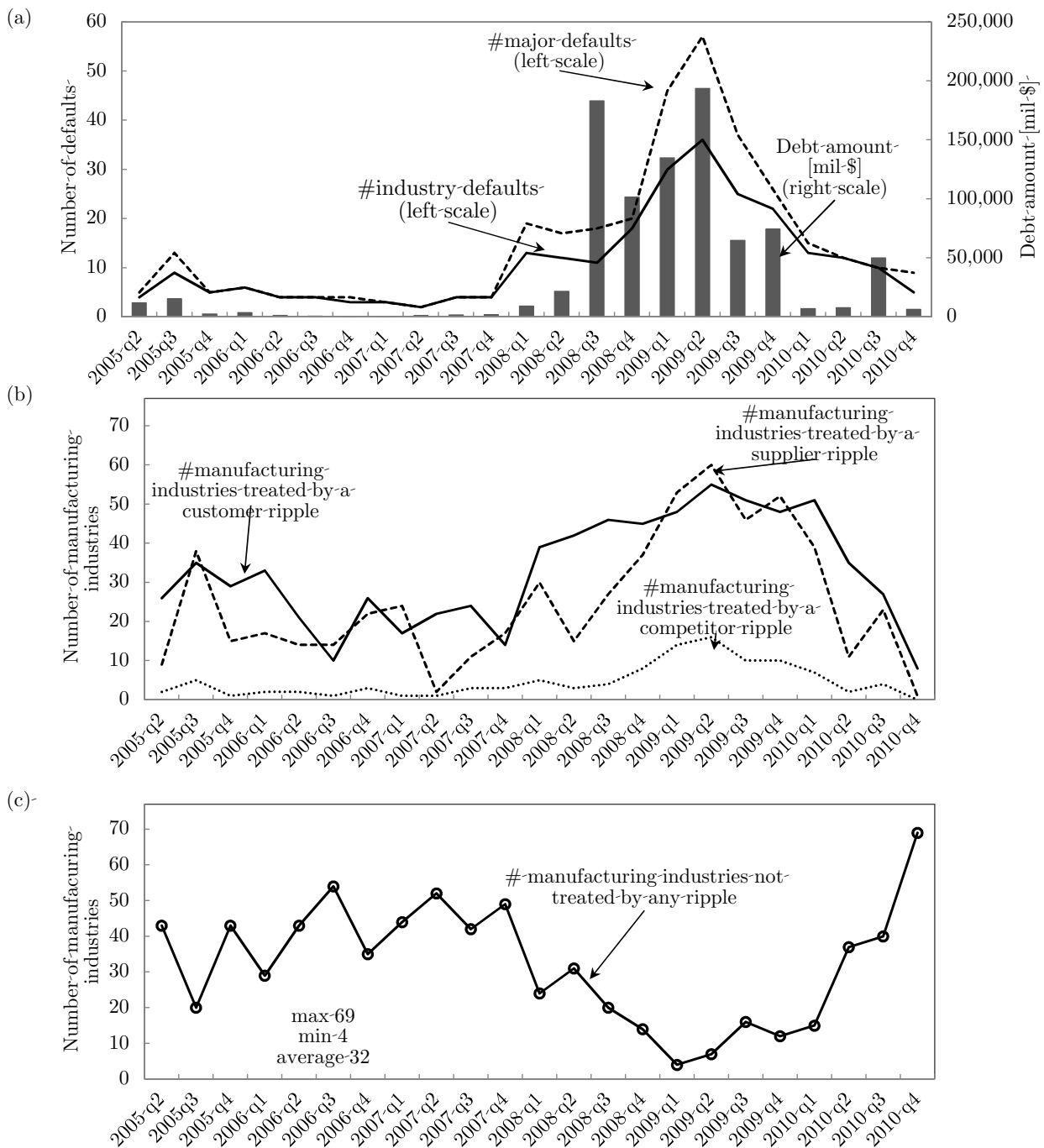
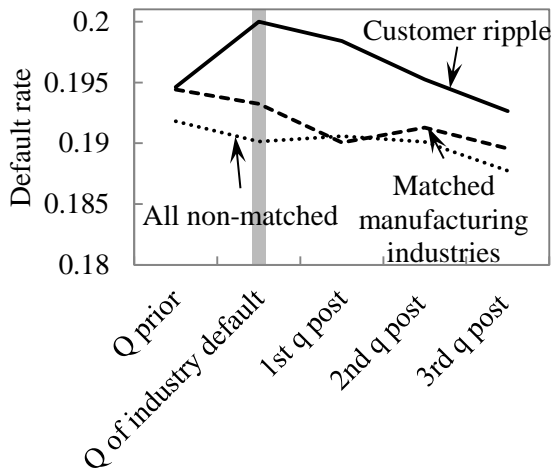


Figure 4: Industry defaults and major defaults. Panel (a) presents time series pattern in industry defaults, number of major defaults on the U.S. S&P rated debt and the debt amount on which the major defaults occurred from 2005 q2 to 2010 q4. Some of the major defaults in one industry fall on the same quarter, so there are 254 unique industry defaults compared to 340 major defaults. Out of the 255 unique industry defaults, 107 occurred in manufacturing industries. Panel (b) presents the number of manufacturing industries that were treated by customer, supplier or competitor ripple and Panel (c) that were not. There are 77 manufacturing industries in total.

(a) Pre-recession



(b) Whole sample

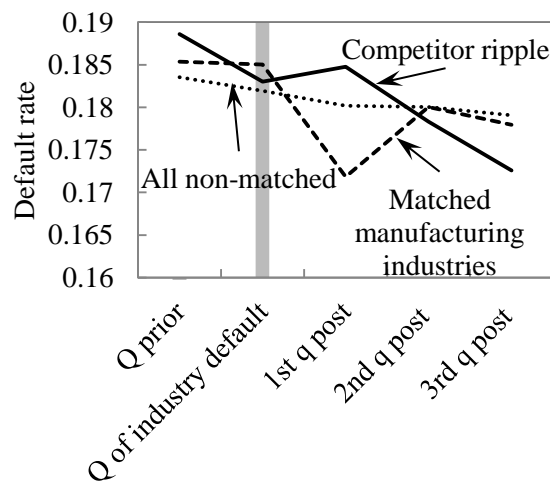
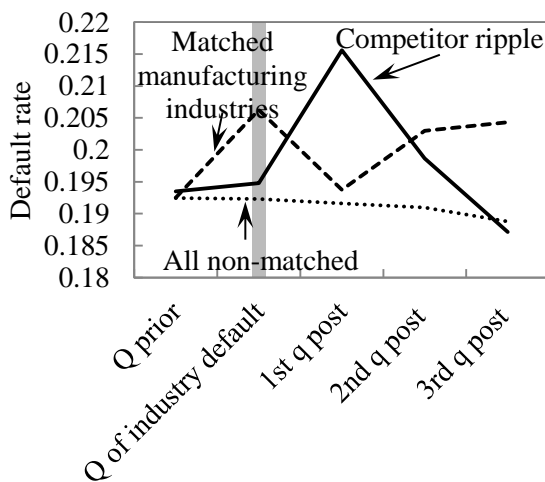
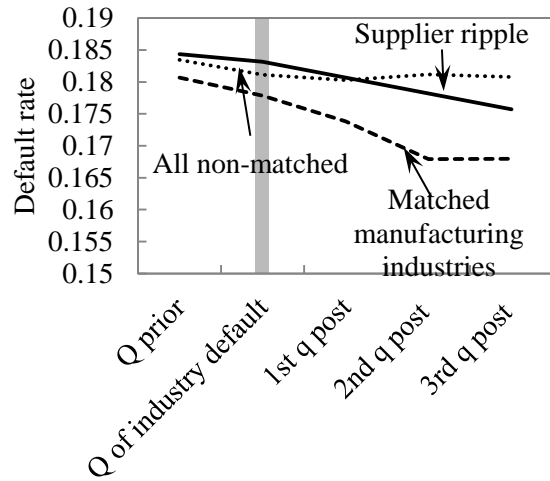
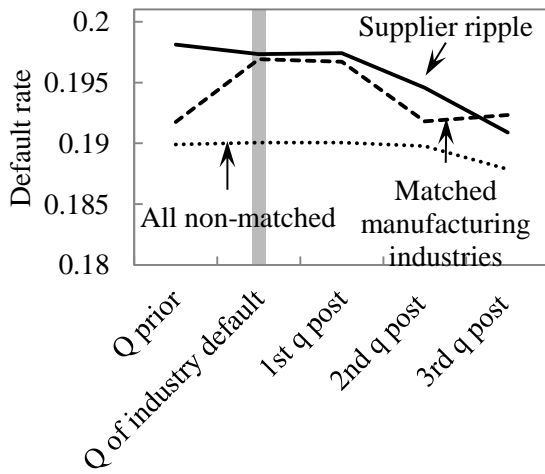
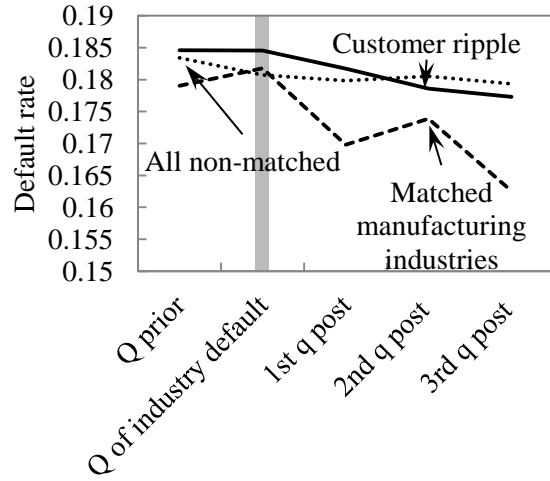


Figure 5: Default rates among small firms in the manufacturing industries. The figure presents default rates in U.S. small manufacturing firms around industry defaults in their customer or supplier industries or in their own industry.

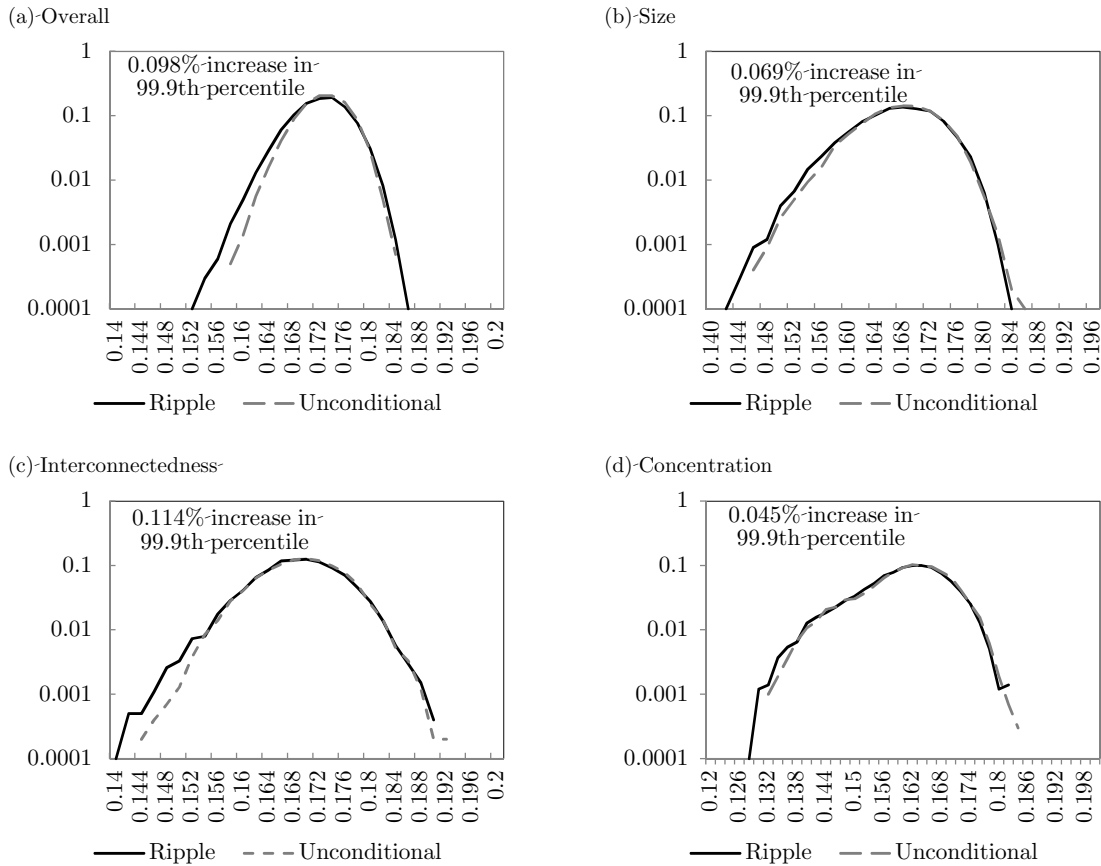


Figure 6: Portfolio default distribution. Ripple effect from a single industry default shifts the density of the portfolio default distribution to the right. It is a consequence of increased expected losses and default correlation. Panel (a) shows ripple effect from a single industry default for a diversified portfolio. Panel (b) shows ripple effect from a single industry default for a portfolio concentrated in large industries. Panel (c) shows ripple effect from a single industry default for a portfolio concentrated in interconnected industries. Panel (d) shows ripple effect from a single industry default for a portfolio concentrated in concentrated industries. The unconditional distribution is given by dashed line and with ripple effects is given by solid line. The distributions are bootstrapped from the historical data in the following way: for the unconditional distribution we randomly draw a quarter for each industry and take the number of defaults and total number of firms that were in that industry at that random quarter. For the distribution with ripple effect we first randomly select an industry default from historical data. Then we define the treated industries as the linked industries (suppliers, customers or the same industry). For them we take the number of defaults and total number of firms that were in those industries for the following quarter. For the non-treated industries we repeat what we did for the unconditional distribution. We repeat that 100.000 times to obtain distribution. Graphs are on a logarithmic scale.

Table I
Summary statistics

The sample runs from 2005 q3 to 2010 q4 and includes 340 major defaults on the S&P rated debt with complete information on industry association. Some of the major defaults in one industry fall on the same quarter, so there are 255 unique industry defaults. The table reports the total amount on which industry default occurred and describes manufacturing industries in the U.S. The industry interconnectedness is the total number of input-output relations as derived from U.S. Bureau of Labor Statistics IO data. The input-output relationships are only those relationships in which either CUST or SUPP take value greater than 1%.

	N	Mean	SD	Min	Max
<i>Panel A: Manufacturing industries characteristics</i>					
Coverage of the small businesses	1,694	1407.836	3617.539	10	37,650.00
Default rate of small businesses (%)	1,694	17.991	5.103	0	40
Dose customer x Post (%)	1,694	0.983	7.576	0	173.905
Dose supplier x Post (%)	1,694	1.923	17.465	0	391.792
Dose competitor x Post (%)	1,694	0.284	4.424	0	173.905
Major default within 1 Y	1,694	0.143	0.351	0	1
Share large firms (%)	1,694	4.425	3.745	0	31.818
Share young firms (%)	1,694	0.982	1.653	0	13.158
Median credit score	1,694	486.419	14.06	461	560
Single customer industry	1,694	0.065	0.246	0	1
Single supplying industry	1,694	0.078	0.268	0	1
Sales [mil]	1,694	64.8	79.7	2.892	773
Inventories [mil]	1,694	6.463	6.853	0.31	51
Industry size	1,694	4,156.37	5,512.68	101	34,385.00
Industry interconnectedness	1,694	30.679	7.267	8	46
Industry concentration	1,694	0.338	0.103	0.109	0.843
<i>Panel B: Major defaults</i>					
Debt amount [mil \$] per major default	340	2,630.81	10,438.96	0	144,426.20

Table II
Ripple effects on industry default rates among small manufacturing firms

This table shows pooled OLS regression estimates (%) based on an industry-quarter observations form manufacturing industries. The dependent variable is small business default rate which measures the rate at which active and financially sound small businesses default within one year. Regression (4) assumes that industry is treated for two quarters after the shock rather than for one. Regression (5) includes the year-times-industry fixed effects. The figures in square brackets represent a percentage change in small business default rate to a one standard deviation change in a given covariate. Standard errors are calculated by clustering at industry level and are reported in parenthesis. Significance is denoted by * at the 90% level, ** at the 95% level and *** at 99% level.

Dependent variable	<i>Post</i> =1 for 1 Q after	<i>Post</i> =1 for 1 Q after	<i>Post</i> =1 for 1 Q after	<i>Post</i> =1 for 2 Qs after	<i>Post</i> =1 for 1 Q after
	Default rate	Default rate	Default rate	Default rate	Default rate
	(1)	(2)	(3)	(4)	(5)
<i>Dose</i> _{<i>Cu</i>} × <i>Post</i> _{<i>Cu</i>}	1.290** (0.586) [0.100]	1.312* (0.711) [0.100]	1.343* (0.718) [0.100]	-0.744 (1.054) [-0.070]	2.519* (1.500) [0.190]
<i>Dose</i> _{<i>Su</i>} × <i>Post</i> _{<i>Su</i>}	-0.271 (0.597) [-0.050]	-0.192 (0.590) [-0.030]	-0.162 (0.583) [-0.030]	-0.432 (0.506) [-0.110]	0.346 (0.508) [0.060]
<i>Dose</i> _{<i>Co</i>} × <i>Post</i> _{<i>Co</i>}	2.833*** (0.774) [0.130]	2.999*** (0.860) [0.130]	2.914*** (0.867) [0.130]	2.387*** (0.510) [0.150]	1.709* (1.004) [0.080]
Major default within 1 Y		0.418 (0.391)	0.371 (0.396)	0.448 (0.389)	0.628 (0.576)
Share large firms		19.097*** (6.320)	18.883*** (6.348)	18.814*** (6.208)	24.838*** (7.763)
Share young firms		-3.515 (17.103)	-0.248 (16.592)	-3.512 (17.073)	-40.254 (25.583)
Median credit score		-0.076 (0.054)	-0.073 (0.054)	-0.077 (0.053)	-0.09 (0.063)
Single customer industry		-0.276 (0.932)	7.845 (6.968)	-0.25 (0.918)	-15.447*** (2.274)
Single supplying industry		7.584*** (0.960)	-3.649 (6.152)	7.625*** (0.944)	-0.139 (2.340)
Sales			-0.008 (0.010)		
Inventories			0.196 (0.199)		
Quarter F.E.	Yes	Yes	Yes	Yes	No
Industry F.E.	Yes	Yes	Yes	Yes	No
Year x Industry F.E.	No	No	No	No	Yes
# Industries	77	77	77	77	77
# Q	22	22	23	22	22
<i>R</i> ²	0.383	0.392	0.394	0.392	0.641
<i>N</i>	1,694	1,694	1,694	1,694	1,694

Table III
Summary statistics for industries receiving ripple in the pre-recession period

The table shows descriptive statistics for manufacturing industries treated by ripple effect, all manufacturing industries not treated by ripple effect and a control (matched industries). Each treated industry is matched to one non-treated manufacturing industry. The matched industries are chosen from all non-treated industries in the same quarter such that they are the closest match based on: major default within one year, share of large firms, share of young firms, median credit score, whether the industry has only one customer, whether the industry has only one supplier, sales and inventories. Additionally, for customer ripple the match is done on: dose during supplier treatment, dose under competitor treatment; for supplier ripple on: dose during customer treatment, dose under competitor treatment; and for competitor ripple on: dose during customer treatment, dose during supplier treatment. We allow for heteroscedasticity in standard errors (4 matches). The sample runs from 2005 q3 to 2007 q3 and includes 77 industries in 9 quarters. Panel A compares industries under treatment by a customer ripple (R) with those that are intact by any customer ripple (A) and with the matched sample (M). Panel B does the same for industries under treatment by supplier ripple and Panel C for industries under treatment by competitor ripple. The column (R-A) reports the two-sample t-test for difference in means between the treated industries and all non-treated. The column (R-M) reports the two-sample t-test for difference in means between the treated industries (R) and the matched industries (M). Standard errors are in parenthesis. Significance is denoted by * at the 90% level, ** at the 95% level and *** at 99% level.

	Industries treated by customer ripple (R)		All industries not treated by customer ripple (A)		Matched industries (M)		Difference in means	
	N	Mean	N	Mean	N	Mean	(R-A)	(R-M)
	Default rate $p_{i,t}$ (%)	155	19.841 (0.397)	615	19.058 (0.195)	105	19.149 (0.387)	0.783 (0.437)
Dose supplier (%)	155	0.565 (0.139)	538	0.506 (0.079)	105	0.482 (0.163)	0.059 (0.166)	0.083 (0.216)
Dose competitor (%)	155	0.046 (0.028)	538	0.040 (0.023)	105	0.032 (0.022)	0.006 (0.045)	0.014 (0.039)
Major default within 1 Y	155	0.123 (0.026)	615	0.068 (0.010)	105	0.114 (0.031)	0.054** (0.024)	0.008 (0.041)
Large firms share (%)	155	4.440 (0.272)	615	4.858 (0.169)	105	4.580 (0.313)	-0.418 (0.364)	-0.140 (0.419)
Young firms share (%)	155	2.383 (0.163)	615	2.368 (0.086)	105	2.445 (0.186)	0.015 (0.191)	-0.062 (0.250)
Median credit score	155	491.855 (1.647)	615	493.329 (0.818)	105	490.724 (1.942)	-1.474 (1.827)	1.131 (2.561)
Sole customer	155	0.000 (0.000)	615	0.081 (0.011)	105	0.000 (0.000)	-0.081*** (0.022)	0.000 (0.000)
Sole supplier	155	0.103 (0.025)	615	0.072 (0.010)	105	0.086 (0.027)	0.032 (0.024)	0.018 (0.037)
Sales	155	95.080 (5.978)	615	57.372 (2.951)	105	61.585 (4.005)	37.709*** (6.600)	33.495*** (7.978)
Inventories	155	8.683 (0.537)	615	5.352 (0.216)	105	6.941 (0.534)	3.331*** (0.507)	1.742** (0.787)

Continued

Panel B: Supplier ripple

	Industries treated by supplier ripple (R)		All industries not treated by supplier ripple (A)		Matched industries (M)		Difference in means	
	N	Mean	N	Mean	N	Mean	(R-A)	(R-M)
	Default rate $p_{i,t}$ (%)	219	19.741 (0.328)	551	19.006 (0.207)	148	19.646 (0.351)	0.735* (0.388)
Dose supplier (%)	219	0.143 (0.035)	474	0.193 (0.036)	148	0.164 (0.046)	-0.05 (0.058)	-0.021 (0.057)
Dose competitor (%)	219	0.024 (0.015)	474	0.049 (0.026)	148	0.022 (0.021)	-0.025 (0.040)	0.002 (0.025)
Major default within 1 Y	219	0.110 (0.021)	551	0.067 (0.011)	148	0.061 (0.020)	0.042** (0.022)	0.049 (0.030)
Large firms share (%)	219	4.621 (0.247)	551	4.835 (0.179)	148	4.176 (0.243)	-0.213 (0.324)	0.446 (0.361)
Young firms share (%)	219	2.294 (0.141)	551	2.402 (0.091)	148	2.336 (0.160)	-0.107 (0.169)	-0.042 (0.216)
Median credit score	219	492.790 (1.367)	551	493.129 (0.868)	148	490.193 (1.544)	-0.339 (1.624)	2.597 (2.092)
Sole customer	219	0.023 (0.010)	551	0.082 (0.012)	148	0.034 (0.015)	-0.059** (0.020)	-0.011 (0.017)
Sole supplier	219	0.009 (0.006)	551	0.105 (0.013)	148	0.027 (0.013)	-0.096*** (0.021)	-0.018 (0.013)
Sales	219	86.338 (7.079)	551	56.467 (2.427)	148	59.669 (3.874)	29.871*** (5.892)	26.669*** (9.185)
Inventories	219	7.651 (0.407)	551	5.376 (0.237)	148	6.508 (0.477)	2.275*** (0.456)	1.143* (0.632)

Panel C: Competitor ripple

	Industries treated by competitor ripple (R)		All industries not treated by competitor ripple (A)		Matched industries (M)		Difference in means	
	N	Mean	N	Mean	N	Mean	(R-A)	(R-M)
	Default rate $p_{i,t}$ (%)	18	21.562 (1.030)	752	19.159 (0.178)	18	19.375 (1.116)	2.403** (1.159)
Dose supplier (%)	18	0.208 (0.160)	675	0.178 (0.028)	18	0.150 (0.057)	0.052 (0.170)	-0.023 (0.096)
Dose competitor (%)	18	0.127 (0.077)	675	0.527 (0.071)	18	0.453 (0.235)	-0.32 (0.434)	-0.245 (0.284)
Major default within 1 Y	18	0.278 (0.109)	752	0.074 (0.010)	18	0.278 (0.109)	0.203*** (0.064)	0.000 (0.154)
Large firms share (%)	18	5.563 (1.445)	752	4.755 (0.146)	18	4.820 (0.703)	0.808 (0.967)	0.743 (1.607)
Young firms share (%)	18	2.455 (0.457)	752	2.369 (0.077)	18	2.507 (0.467)	0.086 (0.506)	-0.052 (0.654)
Median credit score	18	493.472 (4.686)	752	493.022 (0.742)	18	494.250 (5.212)	0.45 (4.850)	-0.778 (7.009)
Sole customer	18	0.000 (0.000)	752	0.066 (0.009)	18	0.000 (0.000)	-0.066 (0.059)	0.000 (0.000)
Sole supplier	18	0.000 (0.000)	752	0.080 (0.010)	18	0.000 (0.000)	-0.08 (0.064)	0.000 (0.000)
Sales	18	101.625 (17.863)	752	64.085 (2.725)	18	81.796 (14.600)	37.54** (17.832)	19.829 (23.070)
Inventories	18	8.670 (1.278)	752	5.960 (0.211)	18	6.911 (1.039)	2.711** (1.379)	1.759 (1.647)

Table IV
Ripple effects in the pre-recession period - matching estimator approach

The table reports the Abadie and Imbens (2007) bias-corrected average treatment effect matching estimator (ATT) for manufacturing industries treated by the ripple effect. Each treated industry is matched to one non-treated manufacturing industry. The matched sample (control) is chosen from all non-treated industries in the same quarter such that it is the closes match based on major default within one year, share of large firms, share of young firms, median credit score, whether the industry has only one customer and whether the industry has only one supplier, sales and inventories. Additionally, for customer ripple the match is done on: dose during supplier treatment, dose under competitor treatment; for supplier ripple on: dose during customer treatment, dose under competitor treatment; and for competitor ripple on: dose during customer treatment, dose during supplier treatment. We allow for heteroscedasticity in standard errors (4 matches). The sample runs from 2005 q3 to 2007 q3 and includes 77 industries in 9 quarters. Standard errors are in parenthesis. Significance is denoted by * at the 90% level, ** at the 95% level and *** at 99% level.

Average treatment effect on small business default rate		
	Pre-recession 2005 q3 to 2007 q3 (1)	All sample period 2005 q3 to 2010 q4 (2)
Customer ripple ATT (%)	0.615 (0.588)	0.786** (0.350)
Supplier ripple ATT (%)	0.739* (0.437)	0.575** (0.268)
Competitor ripple ATT (%)	2.623** (1.226)	0.966 (0.637)

Table V
Industry features and ripple effects on industry default rates among small manufacturing firms

This table shows pooled OLS regression results (%) based on an industry-quarter observations for manufacturing industries. The dependent variable is small business default rate which measures the rate at which active and financially sound small businesses default within one year. All regressions contain controls as in Table II, that is: major default within one year, share large firms, share young firms, median credit score, single customer industry, single supplying industry, sales, and inventories. Standard errors are calculated by clustering at industry level and are reported in parenthesis. Significance is denoted by * at the 90% level, ** at the 95% level and *** at 99% level.

Feature	Size	Inter-connectedness	Concentration
	(1)	(2)	(3)
$Dose_{Cu} \times Post_{Cu}$	1.519 (1.227)	-102.127 (69.128)	5.265*** (1.168)
$Dose_{Su} \times Post_{Su}$	0.414 (0.932)	-2.369 (8.626)	1.382 (1.992)
$Dose_{Co} \times Post_{Co}$	7.201 (7.416)	-147.948 (242.097)	14.405 (25.835)
$Feature \times Dose_{Cu} \times Post_{Cu}$	0.000 (0.000)	6.065 (3.881)	-15.543*** (4.310)
$Feature \times Dose_{Su} \times Post_{Su}$	0.000 (0.000)	0.107 (0.713)	-3.961 (5.506)
$Feature \times Dose_{Co} \times Post_{Co}$	0.000 (0.000)	9.198 (12.931)	-35.772 (78.540)
$Feature$	0.000 (0.000)	-0.909** (0.400)	7.340 (7.068)
$Feature \times Post_{Cu}$	0.000 (0.000)	-0.013 (0.046)	0.214 (0.787)
$Feature \times Post_{Su}$	0.000 (0.000)	-0.016 (0.056)	0.124 (0.956)
$Feature \times Post_{Co}$	0.000 (0.000)	0.039 (0.100)	0.509 (1.444)
$Feature^2 \times Dose_{Cu} \times Post_{Cu}$		-0.086 (0.053)	
$Feature^2 \times Dose_{Su} \times Post_{Su}$		0.000 (0.014)	
$Feature^2 \times Dose_{Co} \times Post_{Co}$		-0.131 (0.169)	
$Feature^2$		0.014** (0.006)	
$Feature^2 \times Post_{Cu}$		0.000 (0.001)	
$Feature^2 \times Post_{Su}$		0.001 (0.002)	
$Feature^2 \times Post_{Co}$		-0.001 (0.003)	
Controls	Yes	Yes	Yes
Quarter F.E.	Yes	Yes	Yes
Industry F.E.	Yes	Yes	Yes
# Industries	77	77	77
# Q	22	22	22
R^2	0.394	0.399	0.395
N	1,694	1,694	1,694

Table VI
Industry features and ripple effects in the pre-recession period - matching estimator approach

The table reports the Abadie and Imbens (2007) bias-corrected average treatment effect matching estimator (ATT) for small business treated by the ripple effect. The ‘high feature’ sub-portfolios contain industries in the top quintile of a given feature. Each treated industry is matched to one non-treated manufacturing industry. The matched sample (control) is chosen from all non-treated industries in the same quarter such that it is the closes match based on: major default within one year, share of large firms, share of young firms, median credit score, whether the industry has only one customer and whether the industry has only one supplier, sales and inventories. Additionally, for customer ripple the match is done exactly on $Post_{Cu}$ and continuously on: dose during supplier treatment, dose under competitor treatment; for supplier ripple exactly on $Post_{Su}$ and continuously on: dose during customer treatment, dose under competitor treatment; and for competitor ripple exactly on $Post_{Co}$ and continuously on: dose during customer treatment, dose during supplier treatment. We allow for heteroscedasticity in standard errors (4 matches). The sample runs from 2005 q3 to 2007 q3 and includes 77 industries in 9 quarters. Standard errors are in parenthesis. Significance is denoted by * at the 90% level, ** at the 95% level and *** at 99% level.

Average treatment effect on small business default rate			
	High size (1)	High interconnectedness (2)	High concentration (3)
<i>Panel A: Pre-recession period</i>			
Customer ripple ATT (%)	1.477** (0.606)	-1.995** (0.954)	-0.499 (0.850)
Supplier ripple ATT (%)	-2.549*** (0.410)	-4.008*** (0.541)	2.055** (0.816)
Competitor ripple ATT (%)	-0.990 (1.782)	-1.193 (0.999)	insufficient observations
<i>Panel B: All sample</i>			
Customer ripple ATT (%)	0.107 (0.323)	0.662 (0.638)	-1.102** (0.469)
Supplier ripple ATT (%)	-1.793*** (0.272)	0.143 (0.433)	-0.384 (0.473)
Competitor ripple ATT (%)	0.069 (0.820)	0.760 (0.746)	-0.945 (2.201)