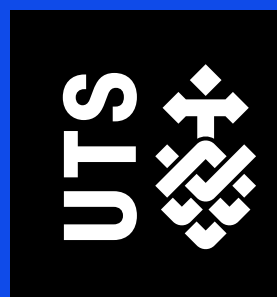


Optimal Model Selection in RDD and Related Settings Using Placebo Zones



Nathan Kettlewell and Peter Siminski (UTS)

ASSA, January 8, 2023

Our 2 Stata Programs

ssc install pzms

<https://econpapers.repec.org/software/bocbocode/s459073.htm>

-pzms- implements our approach. Very easy to use. Only required option is the maximum bandwidth

[pzms sim](#) uses simulations based on the data from any application, to examine likely performance of our approach, compared to alternative approaches <https://sites.google.com/site/nrkettlewell/research>

Outline

- **Key Issues and our Contribution**
- Motivating application – 2 policy changes affecting Learner drivers in NSW
- Theory – show our approach is asymptotically optimal, under restrictive conditions
- Simulations – our approach performs favourably compared to other procedures using stylised and realistic DGPs

RDD Model Selection

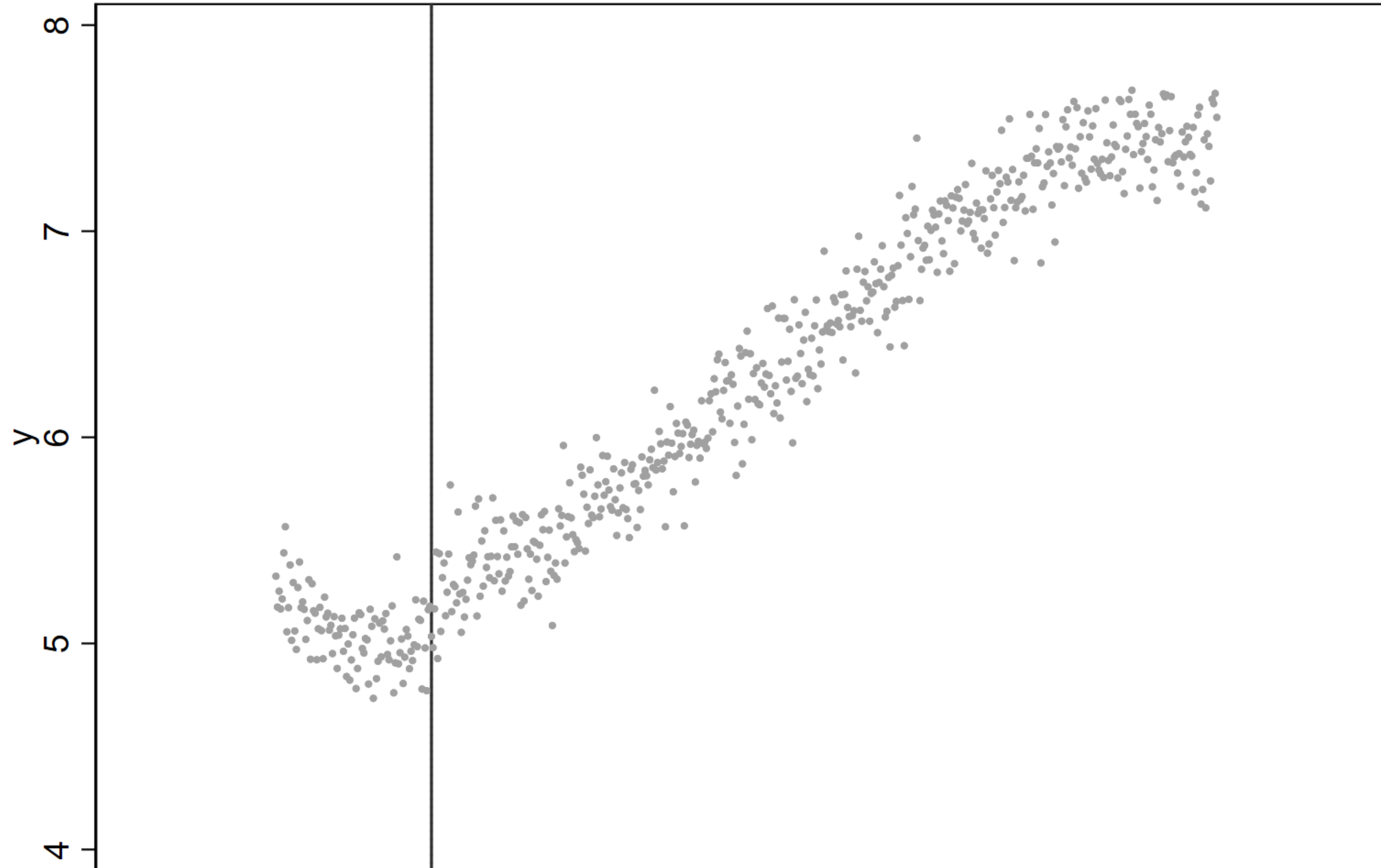
- RDD regarded as a leading quasi-experimental research design
- However, it involves numerous researcher choices e.g. bandwidth, polynomial, kernel, controls etc.
- How to select among the multitudes of potential estimators?
 - Imbens and Kalyanaraman (IK) (2012) and Calonico et al. (CCT) (2014) propose algorithms for BW selection that minimise AMSE of the boundary estimator
 - Pei et al (2021) make a similar suggestion for polynomial order.

Issues with Model Selection Algorithms

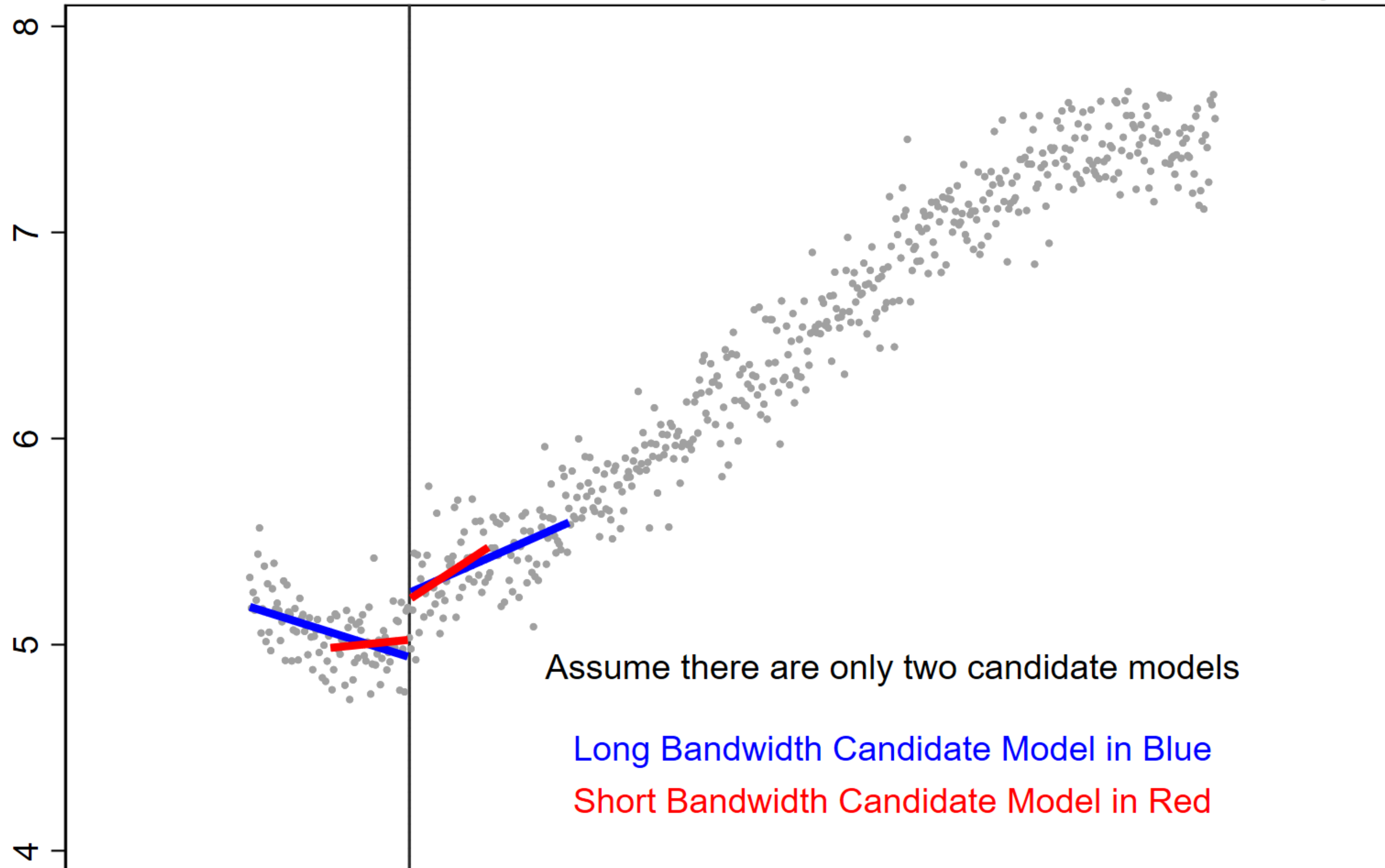
- Existing approaches deal with one choice and hold others constant.
- CCT focussed on inference, not on estimation
- IK/CCT can do poorly in simulations with realistic DGPs (Card et al, 2017).
- While IK/CCT are popular for BW selection, there is no consensus and researchers tend to rely on robustness testing. This may be overly punitive to particular DGPs.

Illustration of our proposal

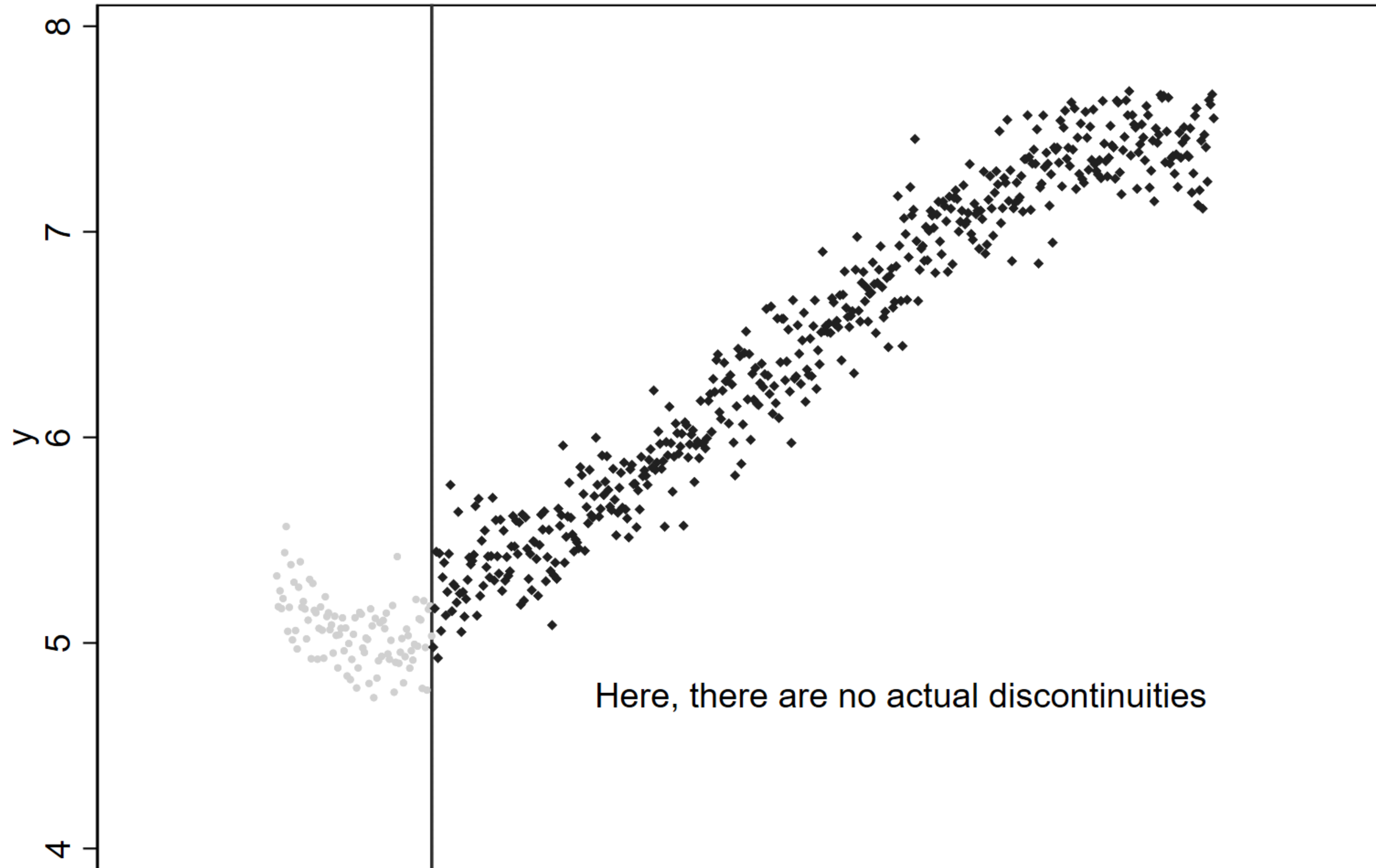
Which Model Should we Use to Estimate the Discontinuity?



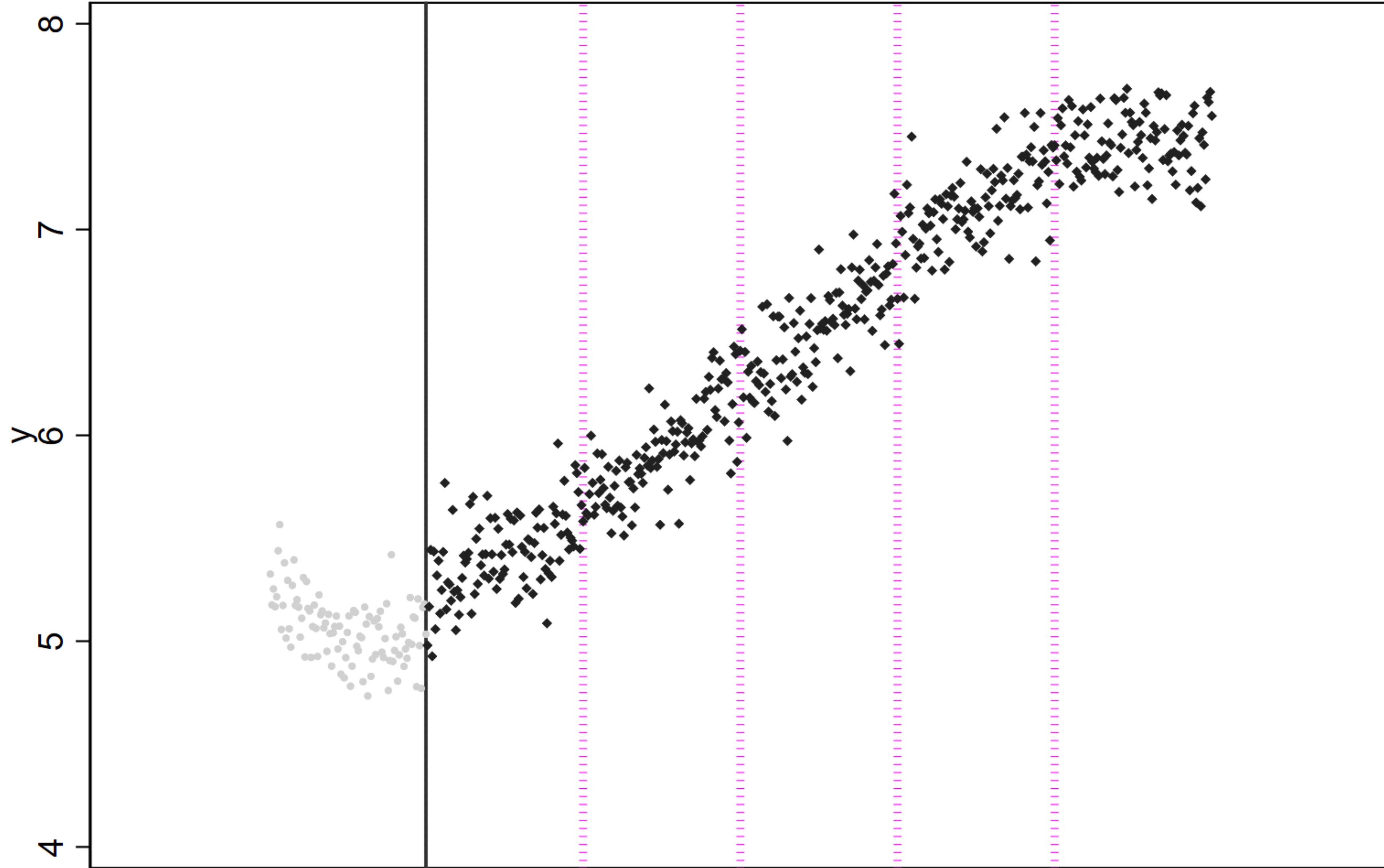
Which Model Should we Use to Estimate the Discontinuity?



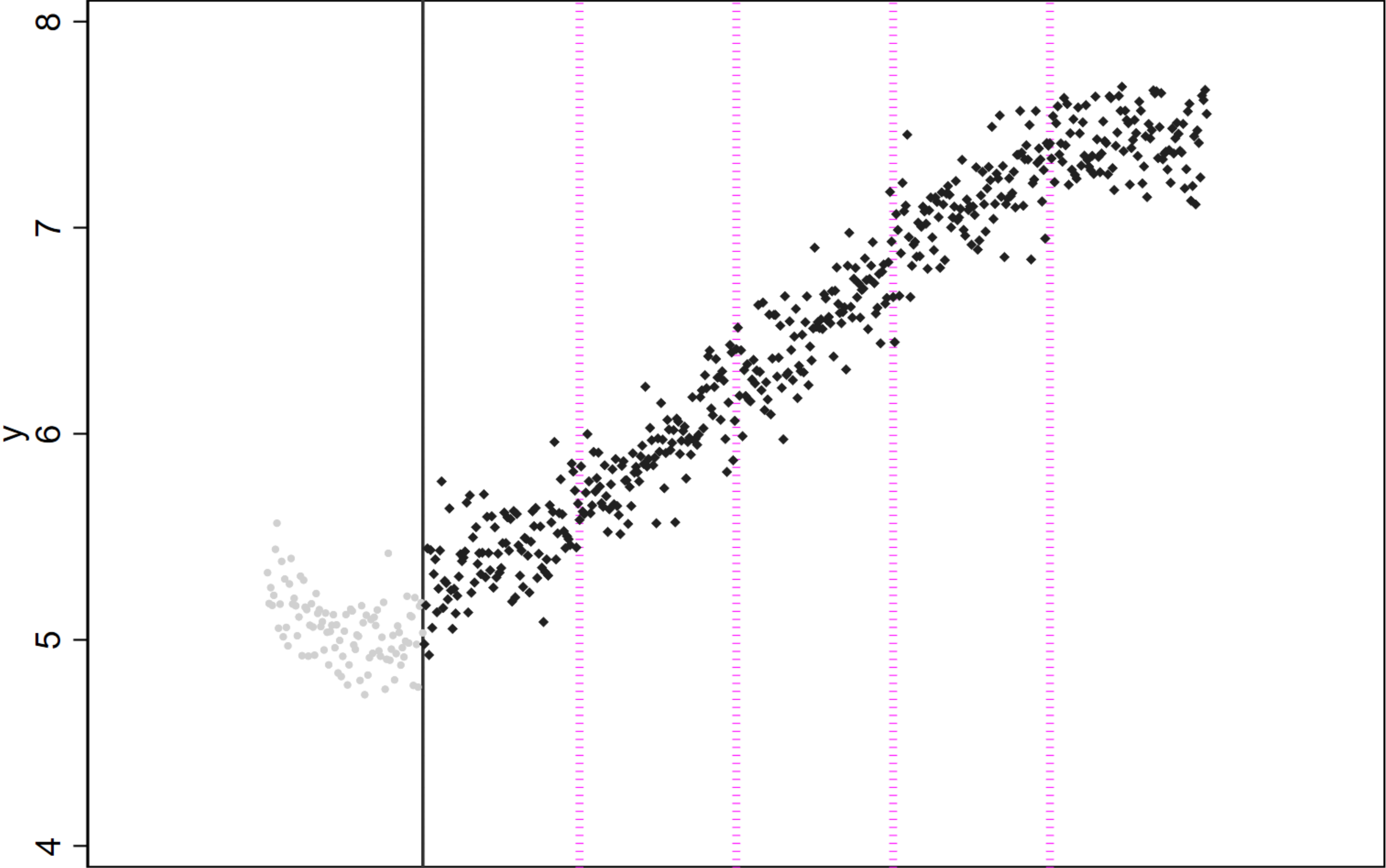
We can test Candidate Models in the Placebo Zone

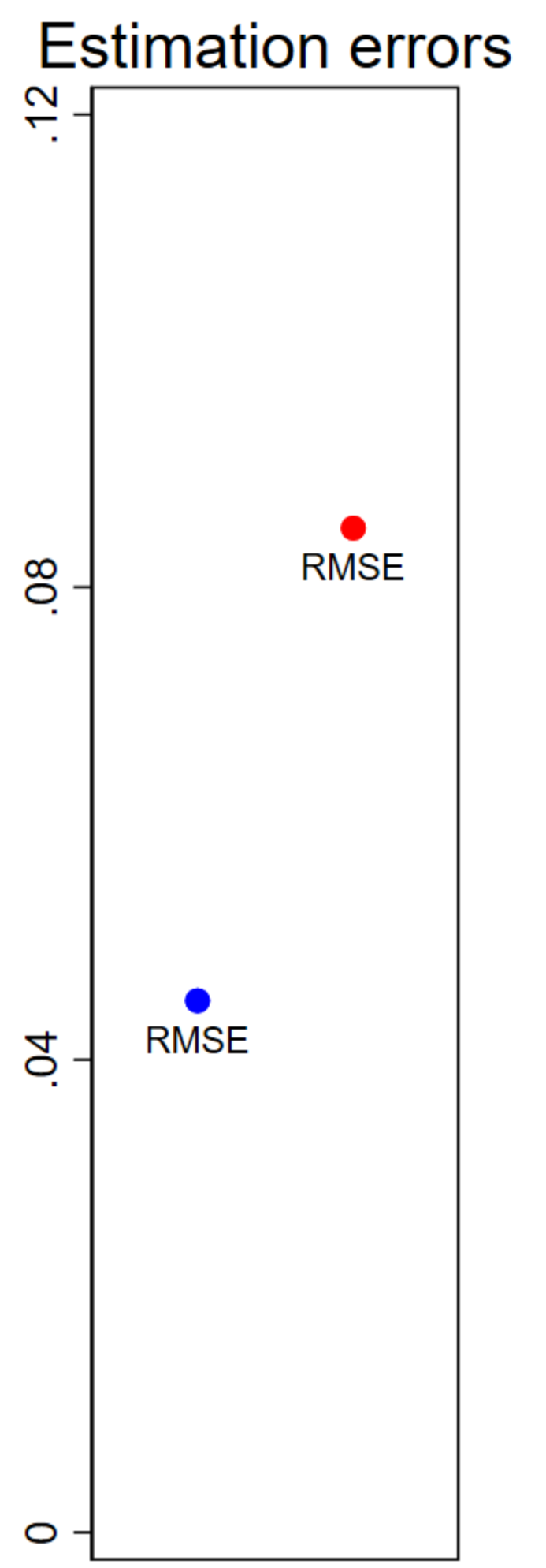
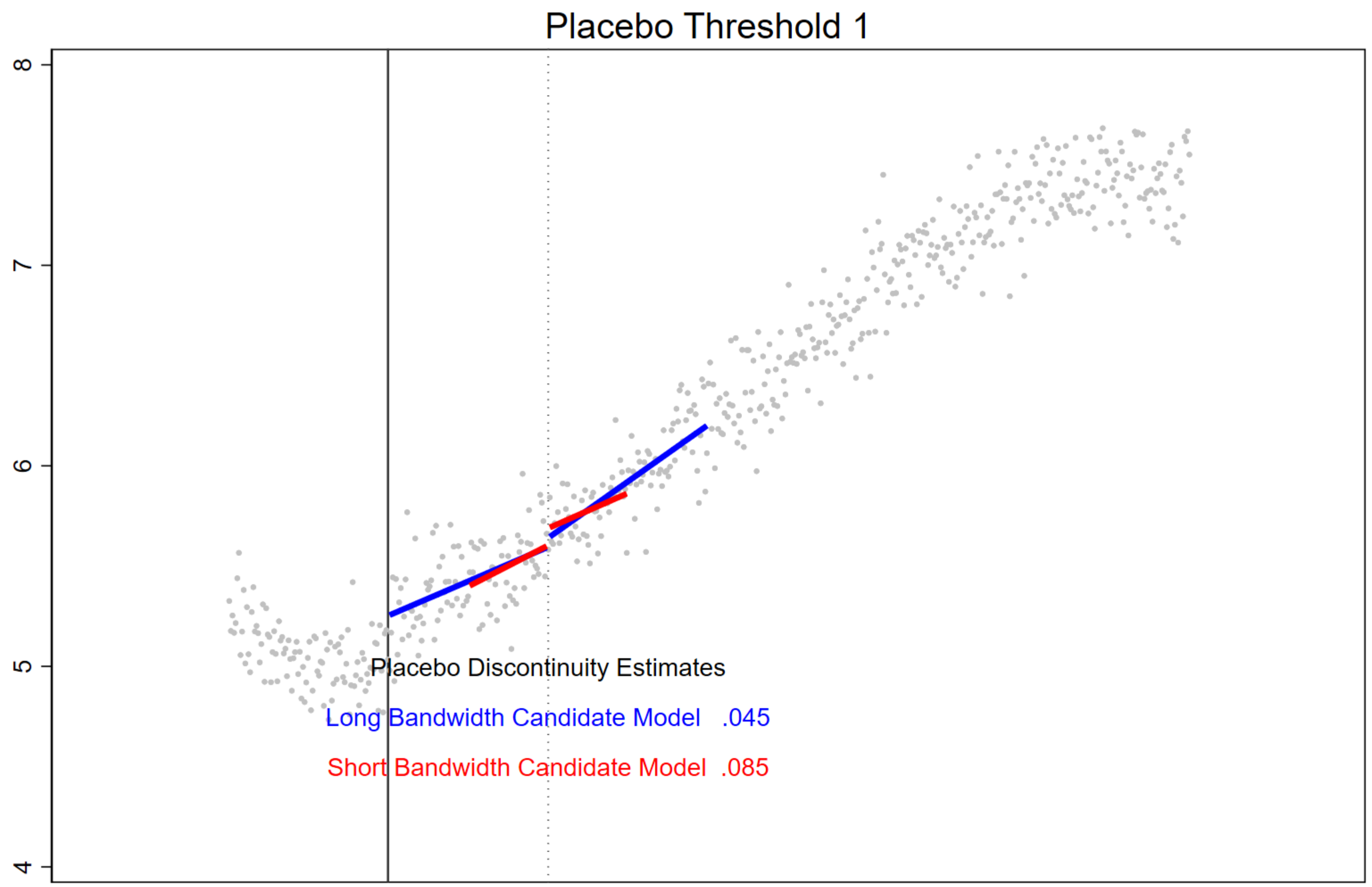


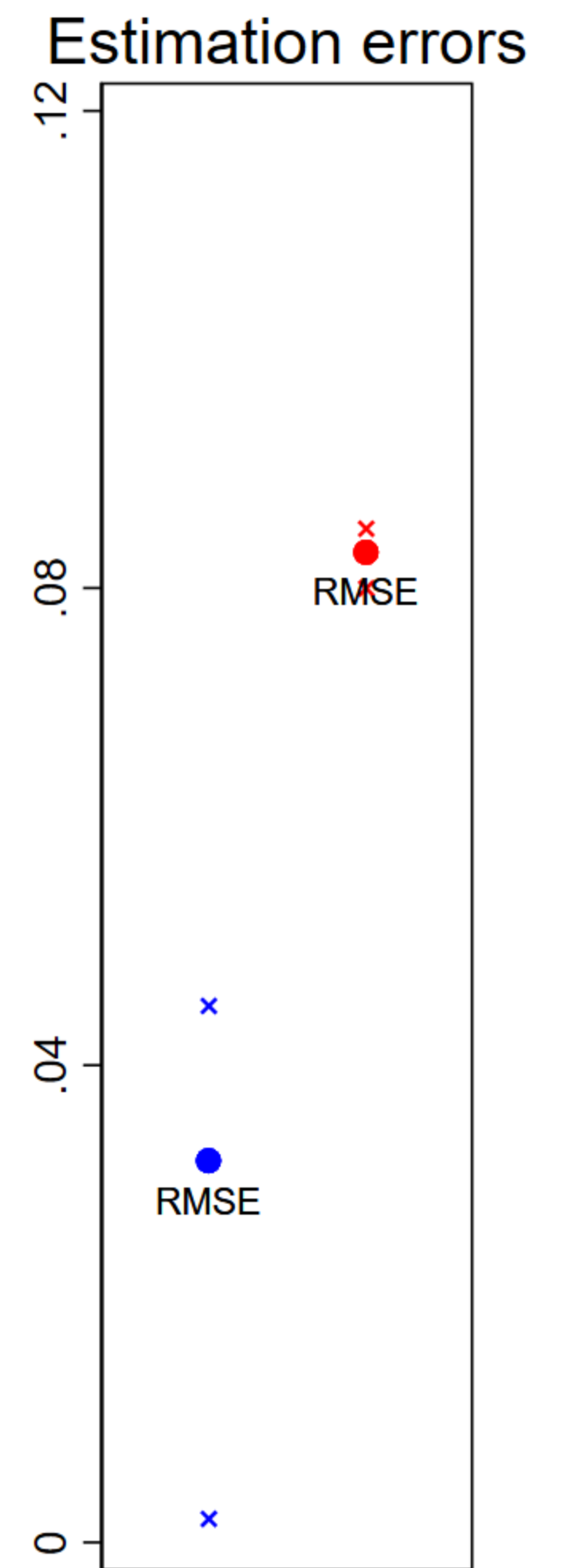
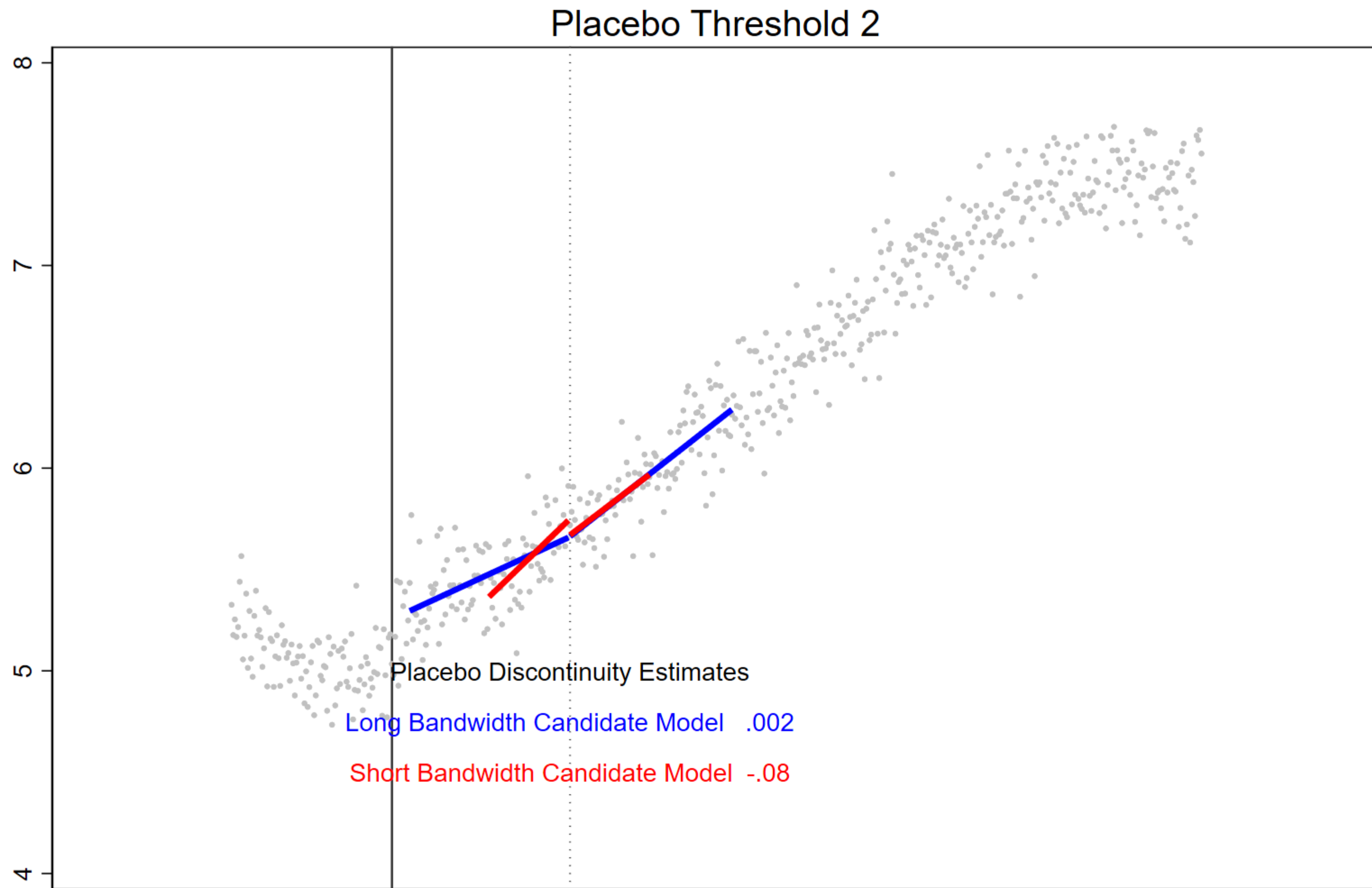
How well does each model estimate 'placebo' discontinuities?

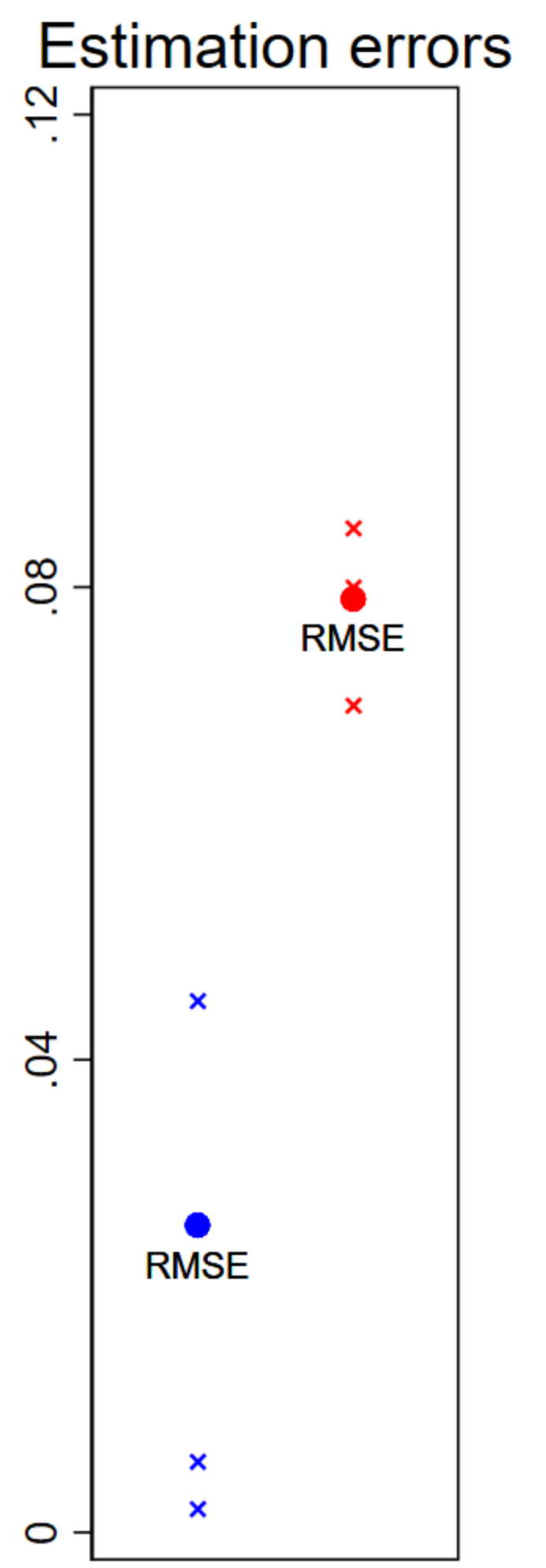
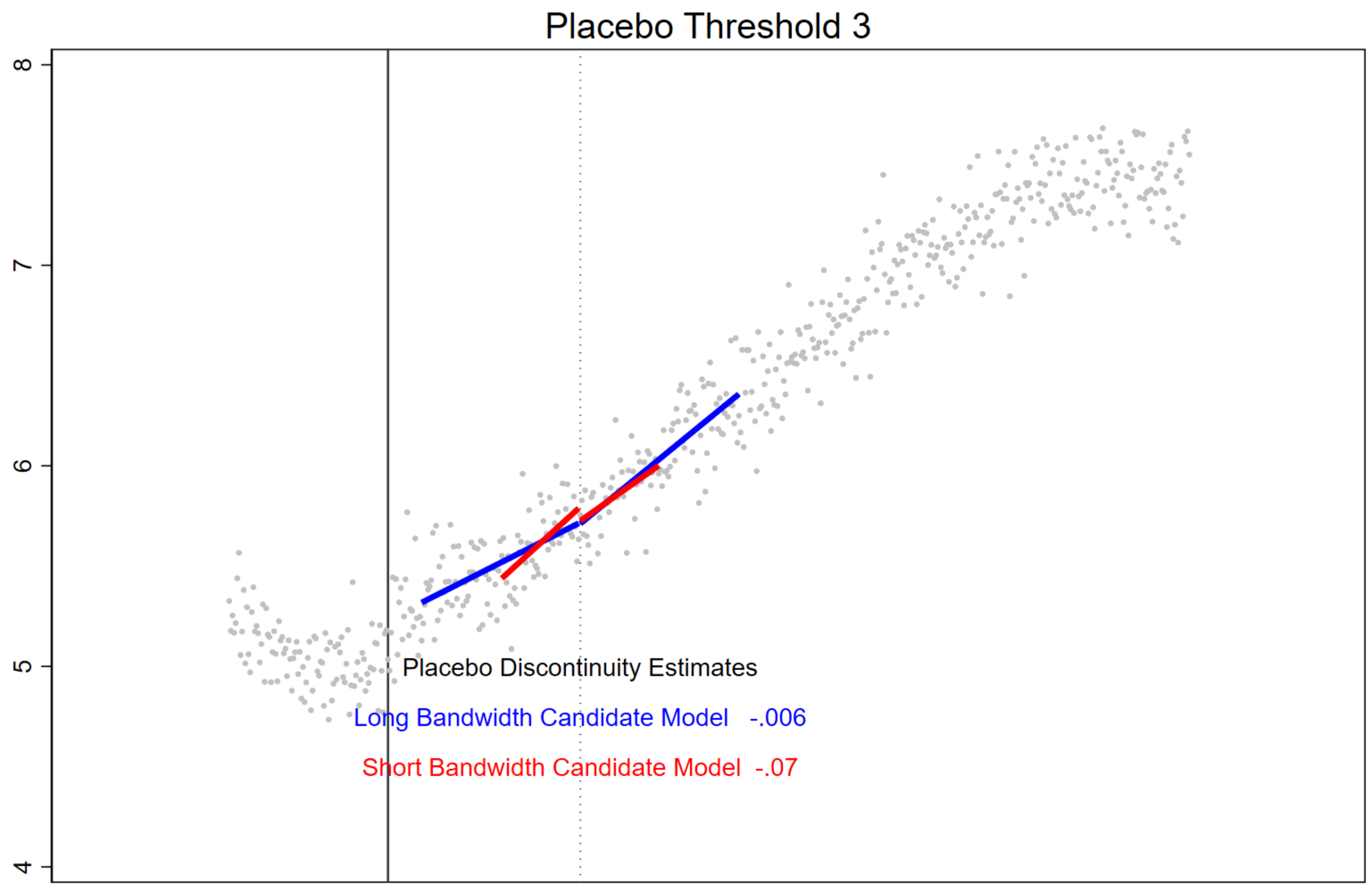


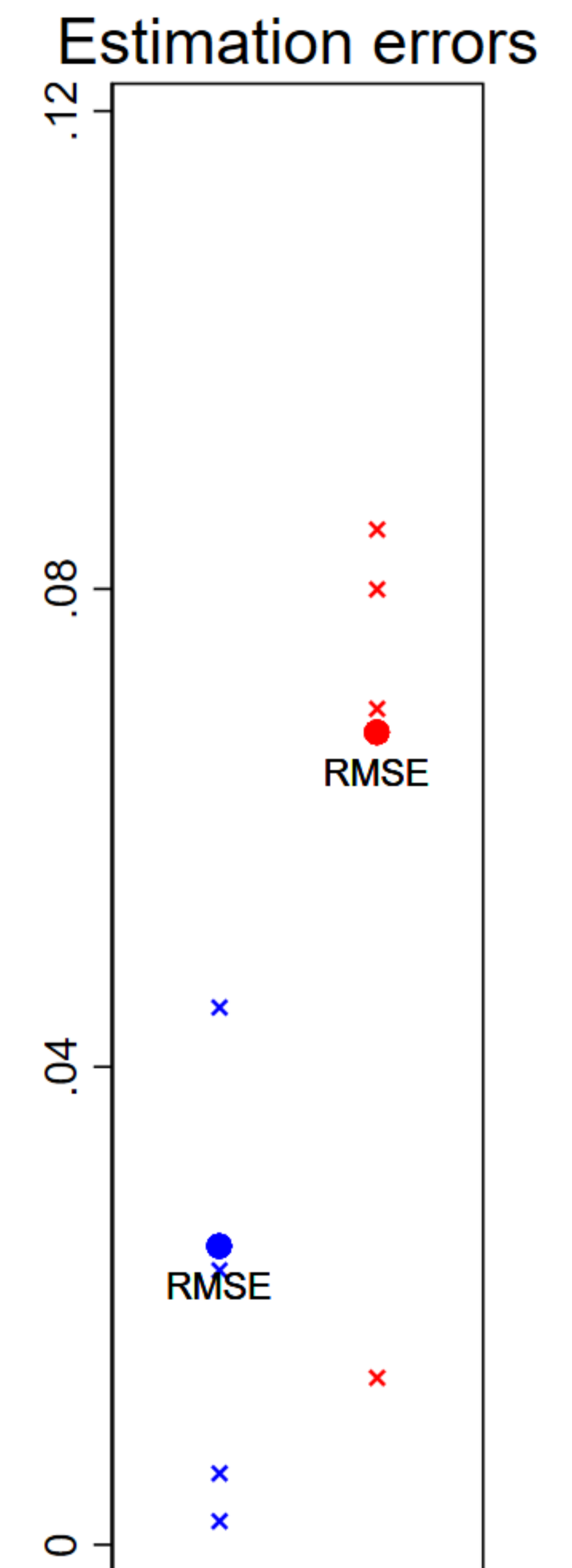
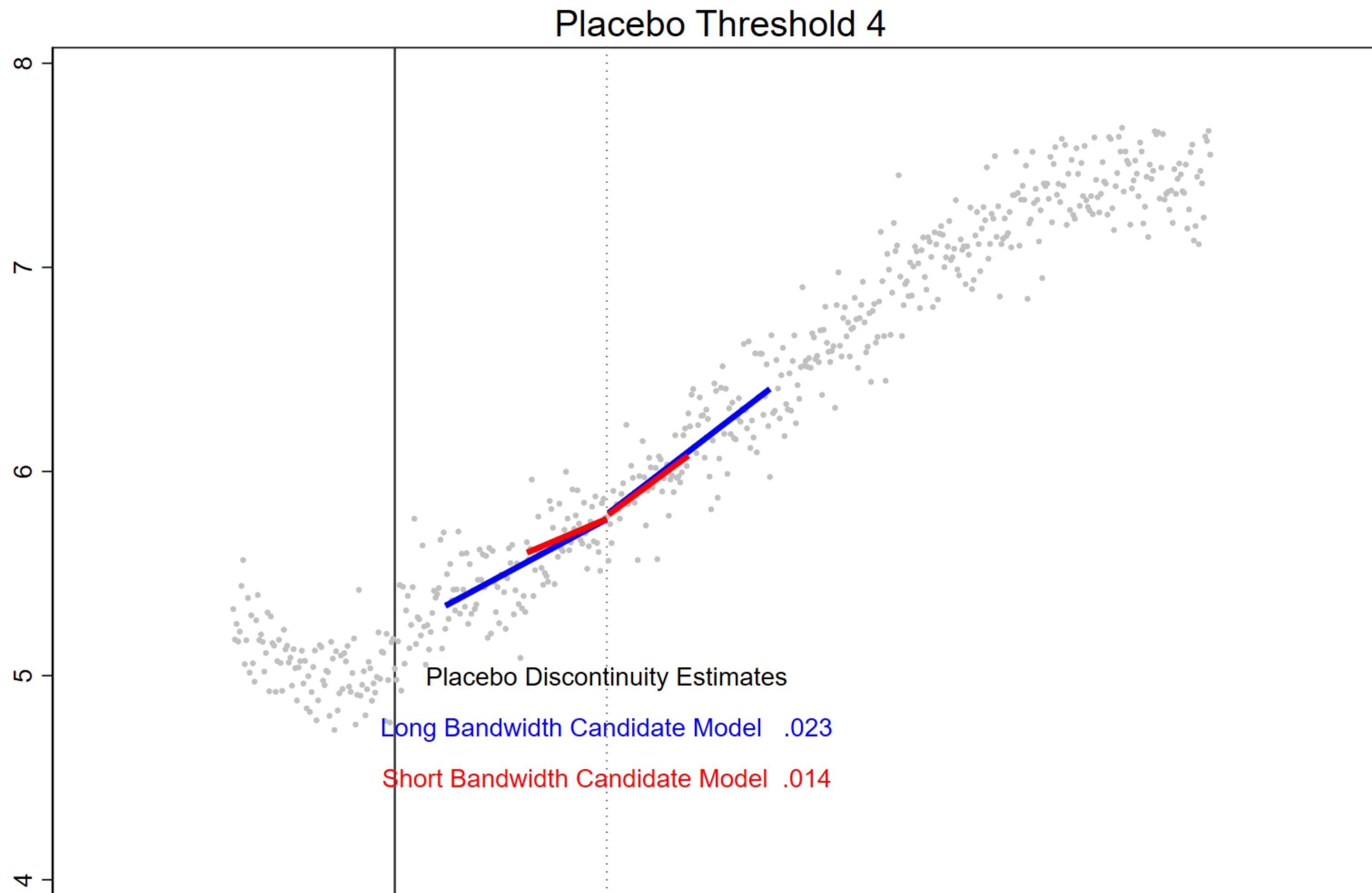
Let's See!

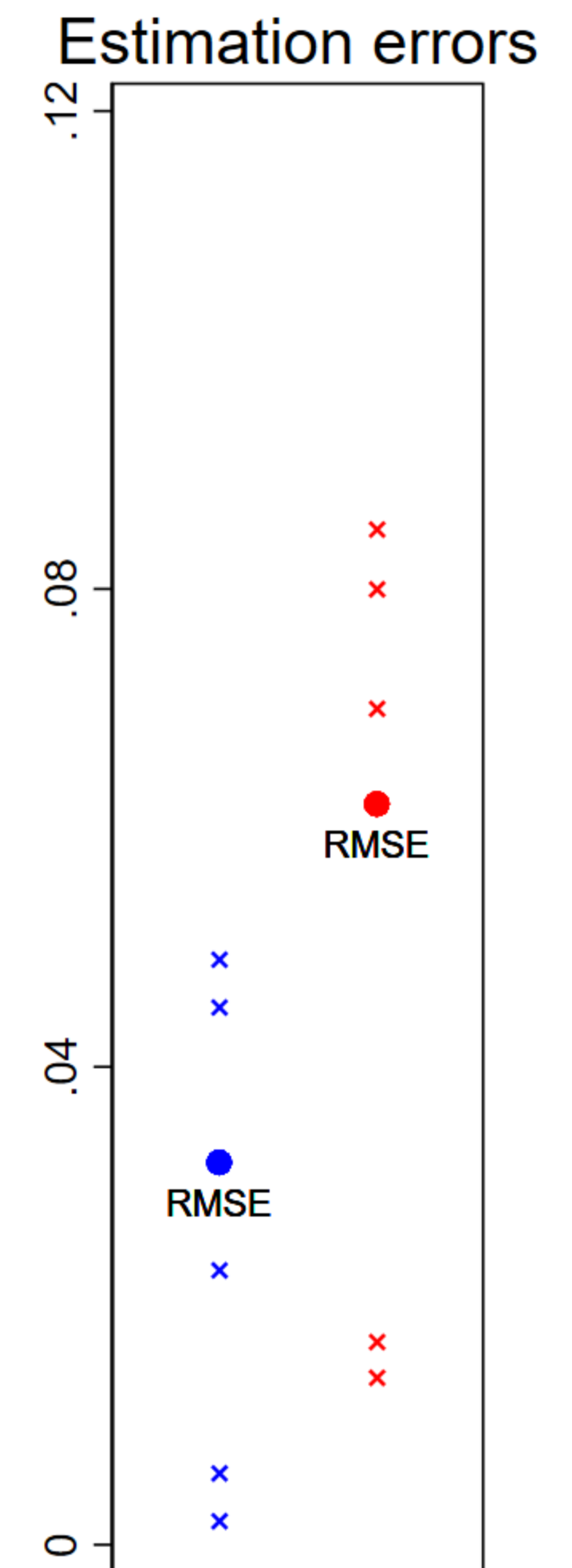
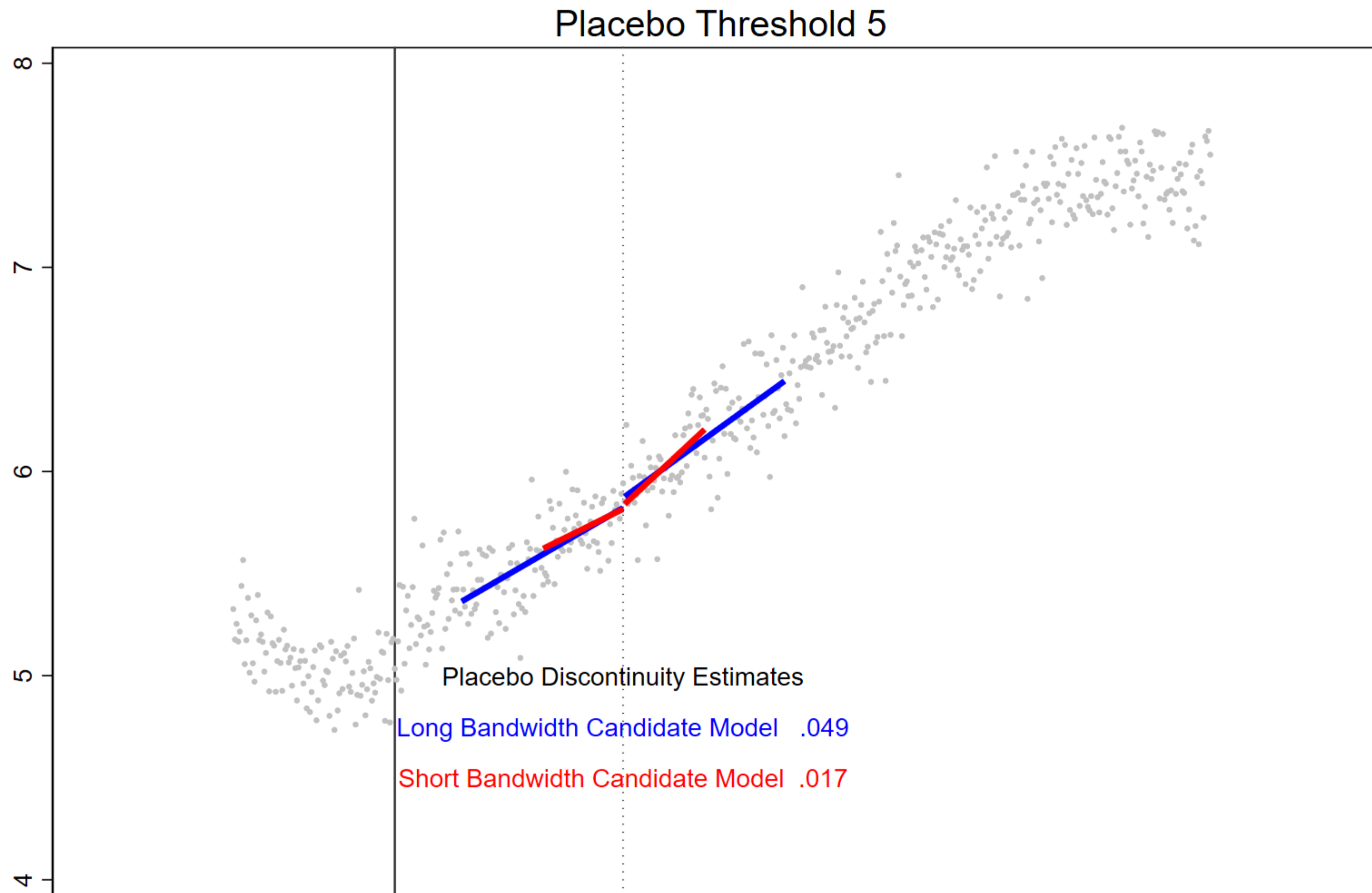


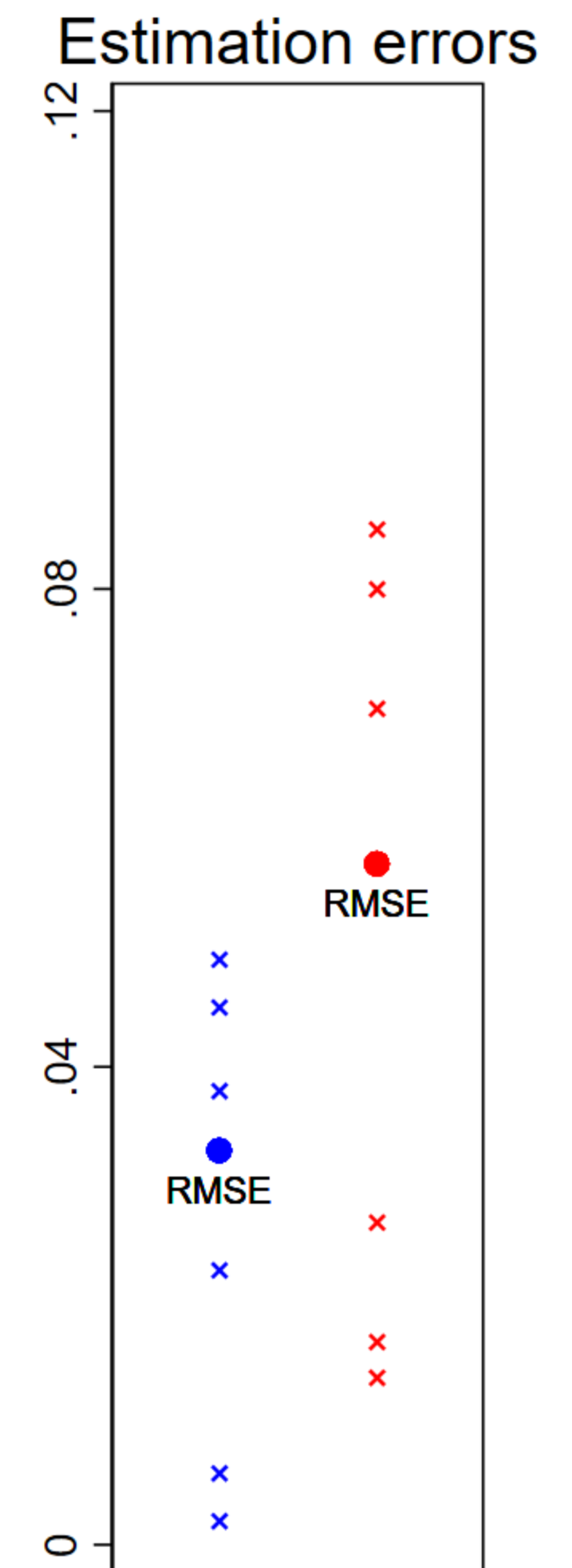
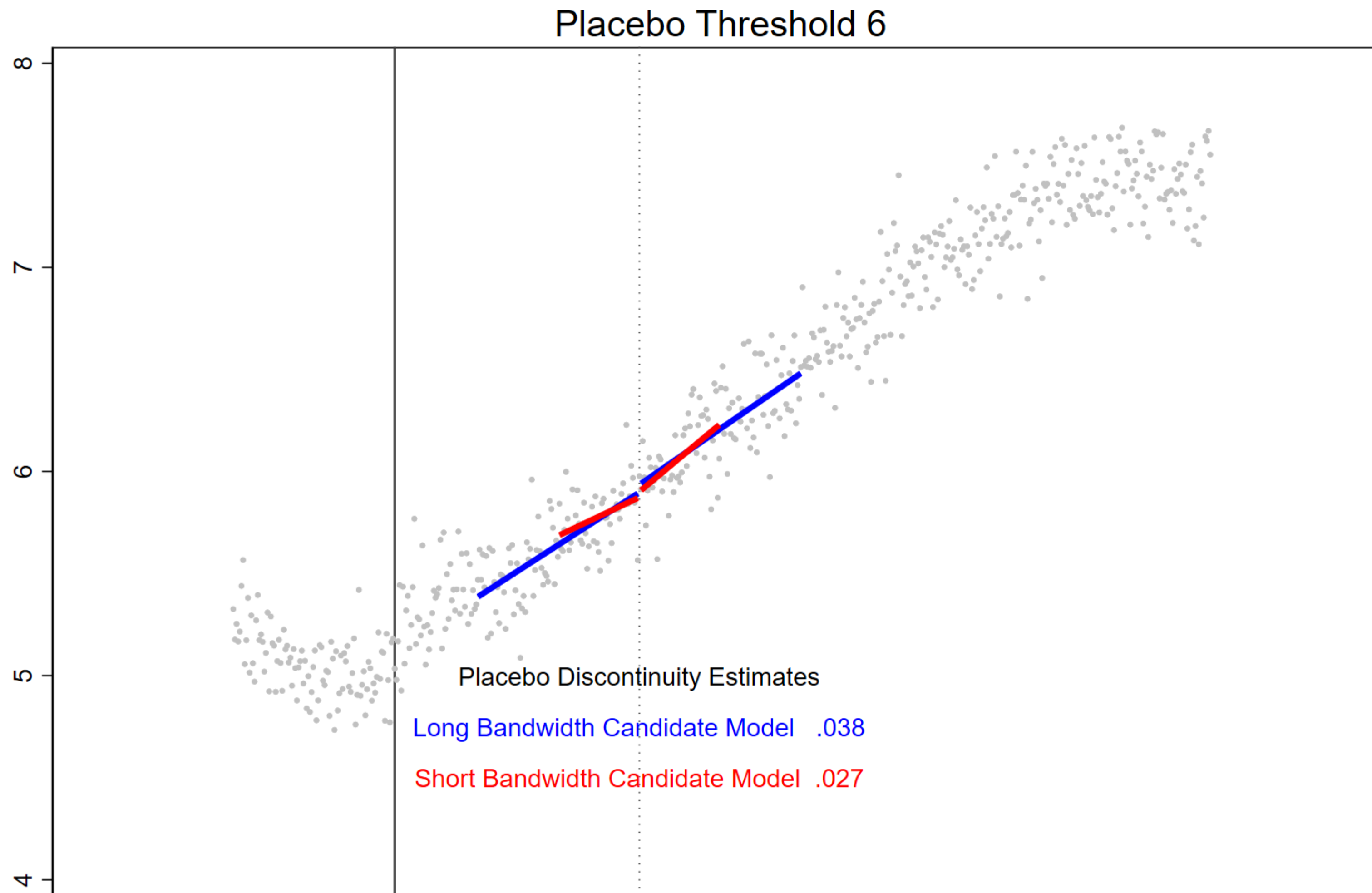


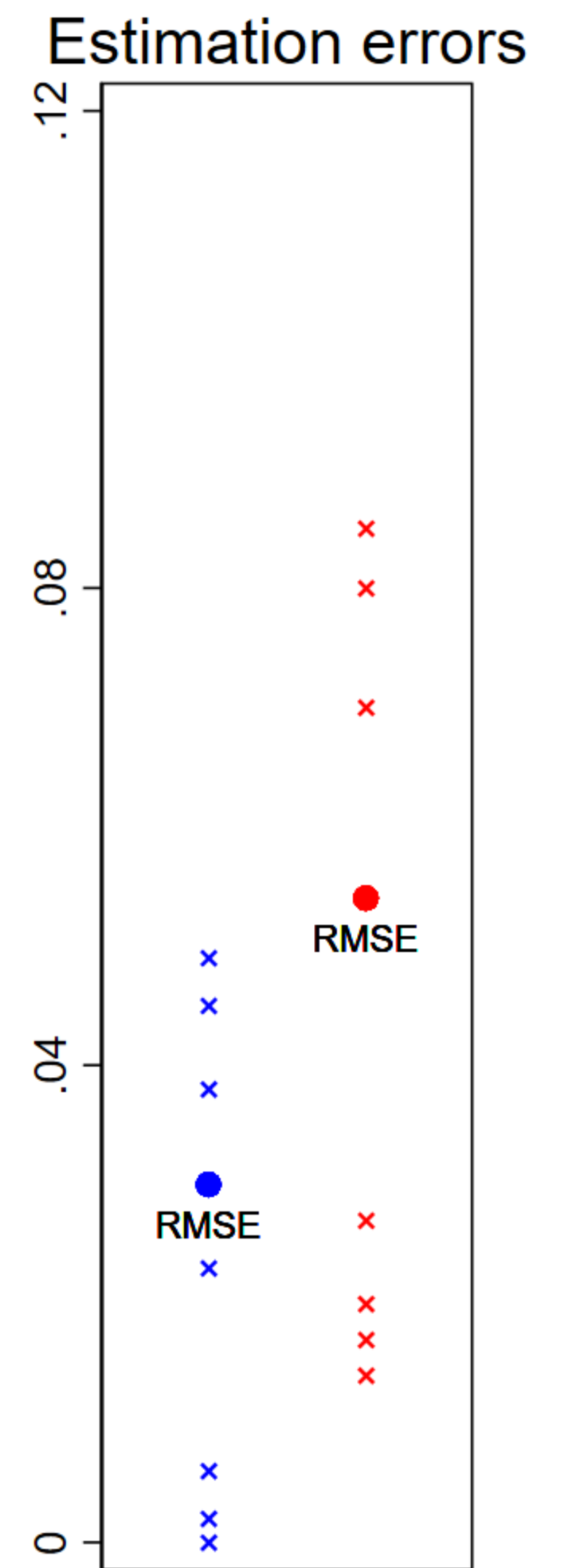
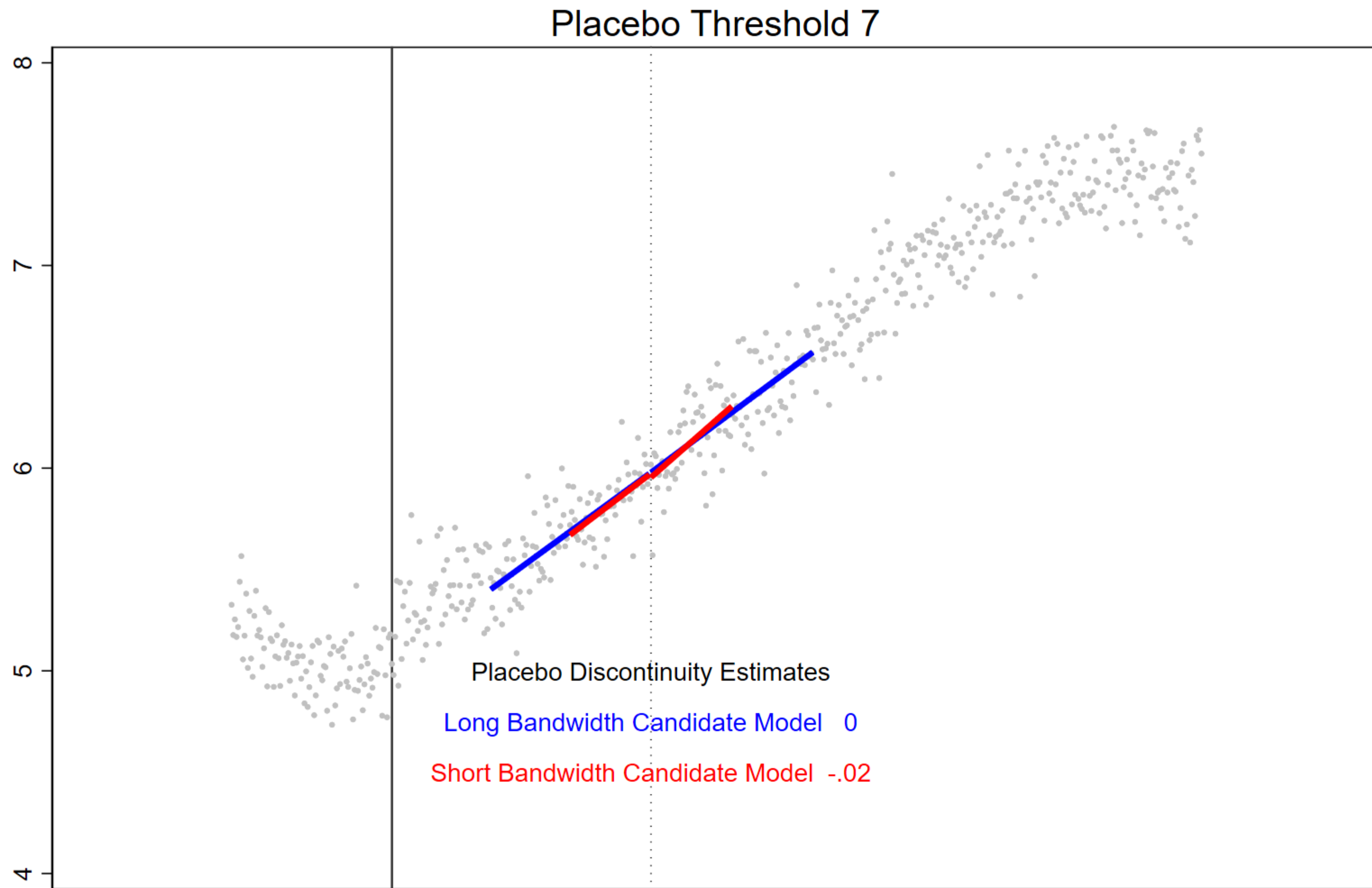


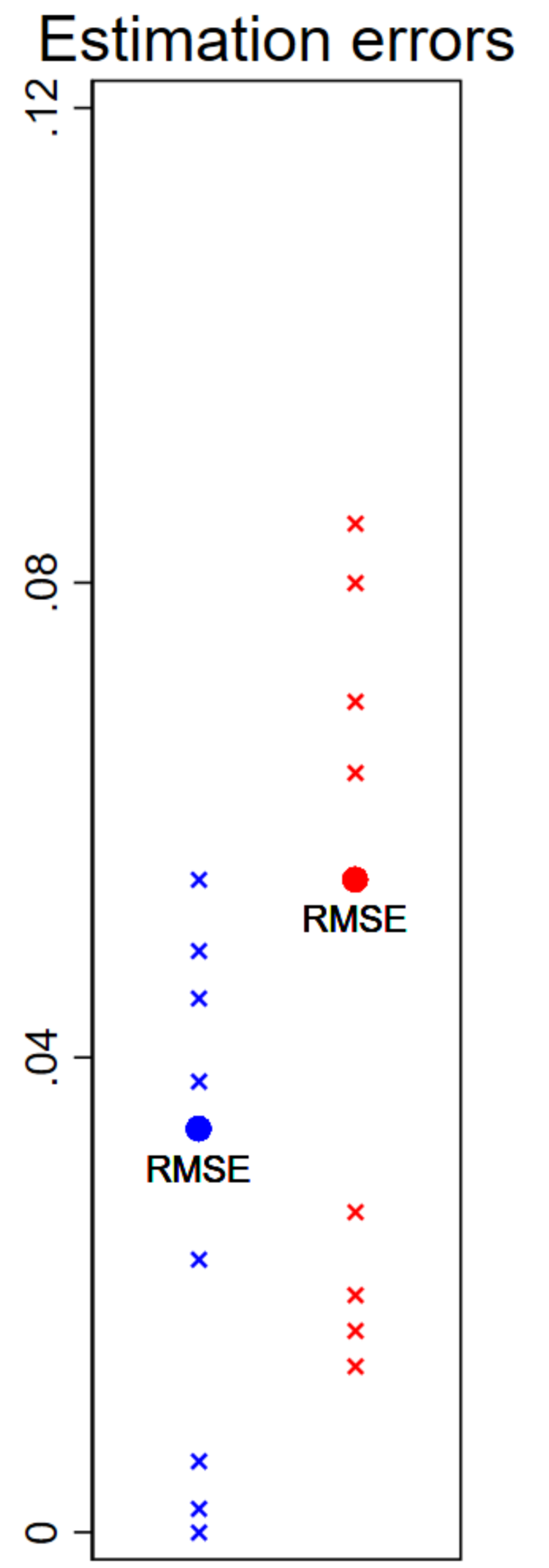
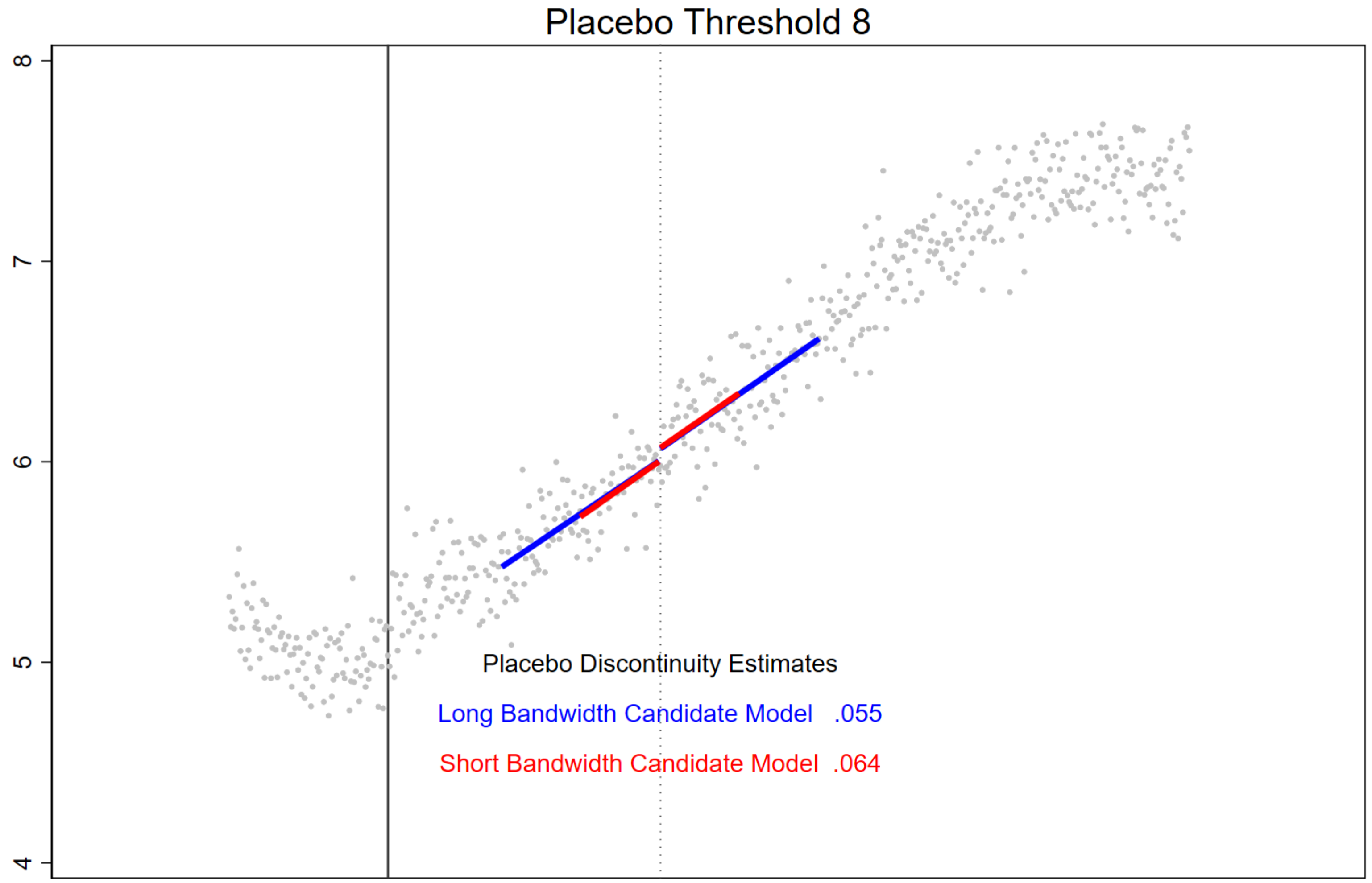


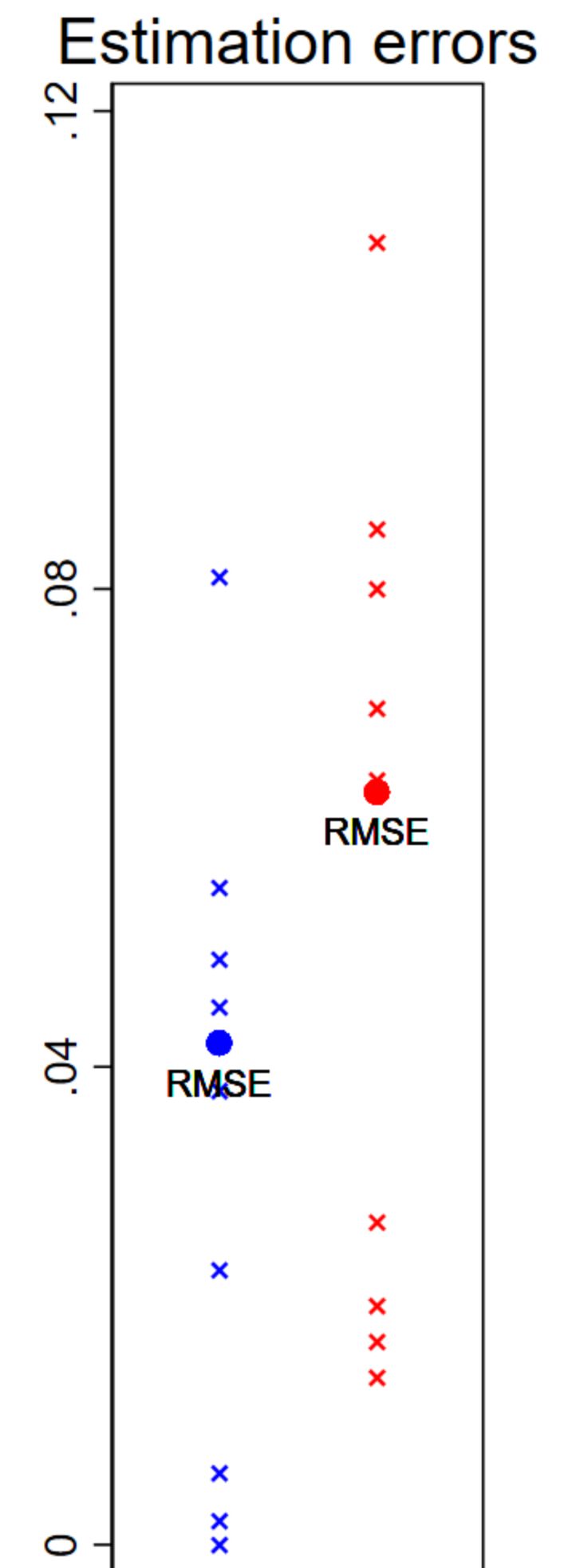
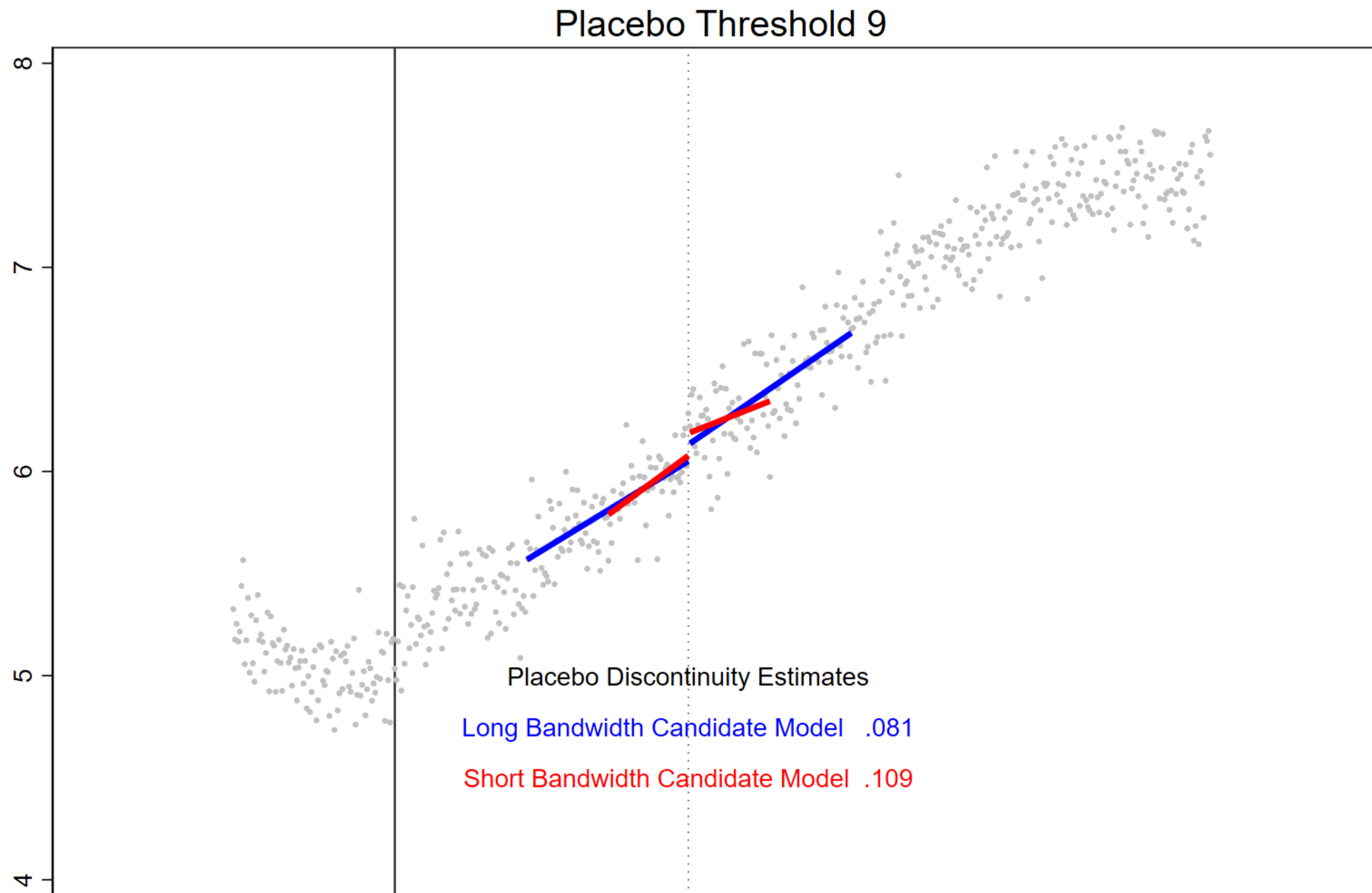


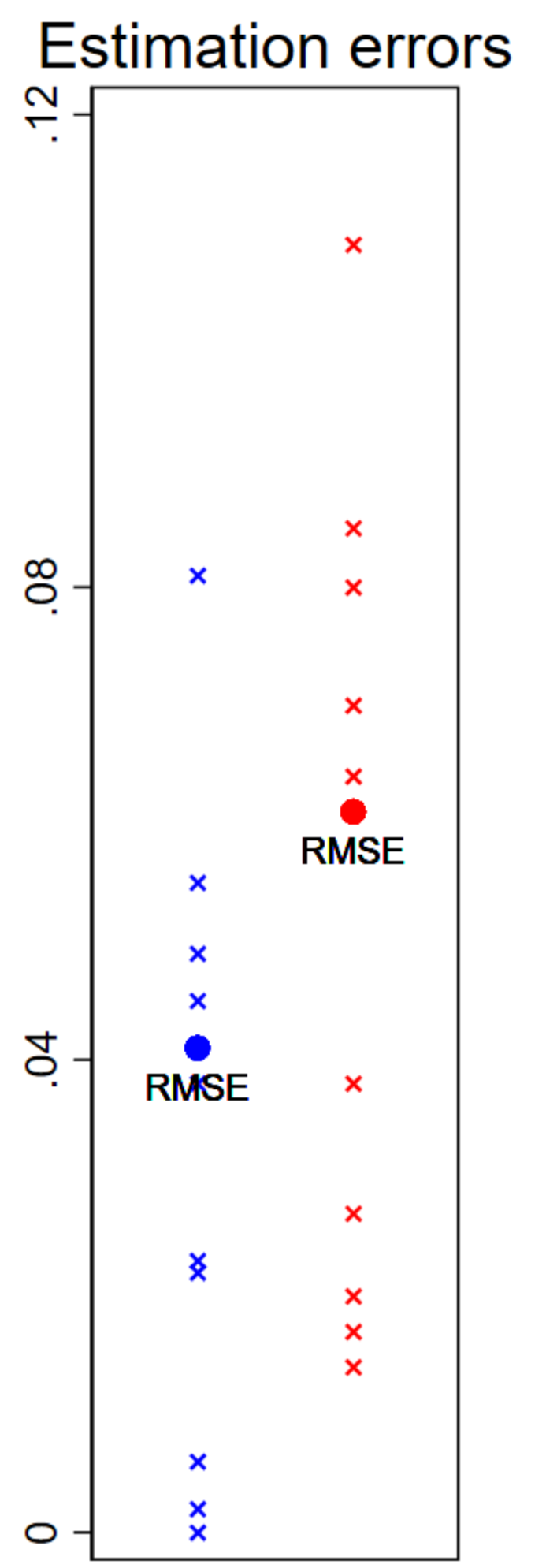
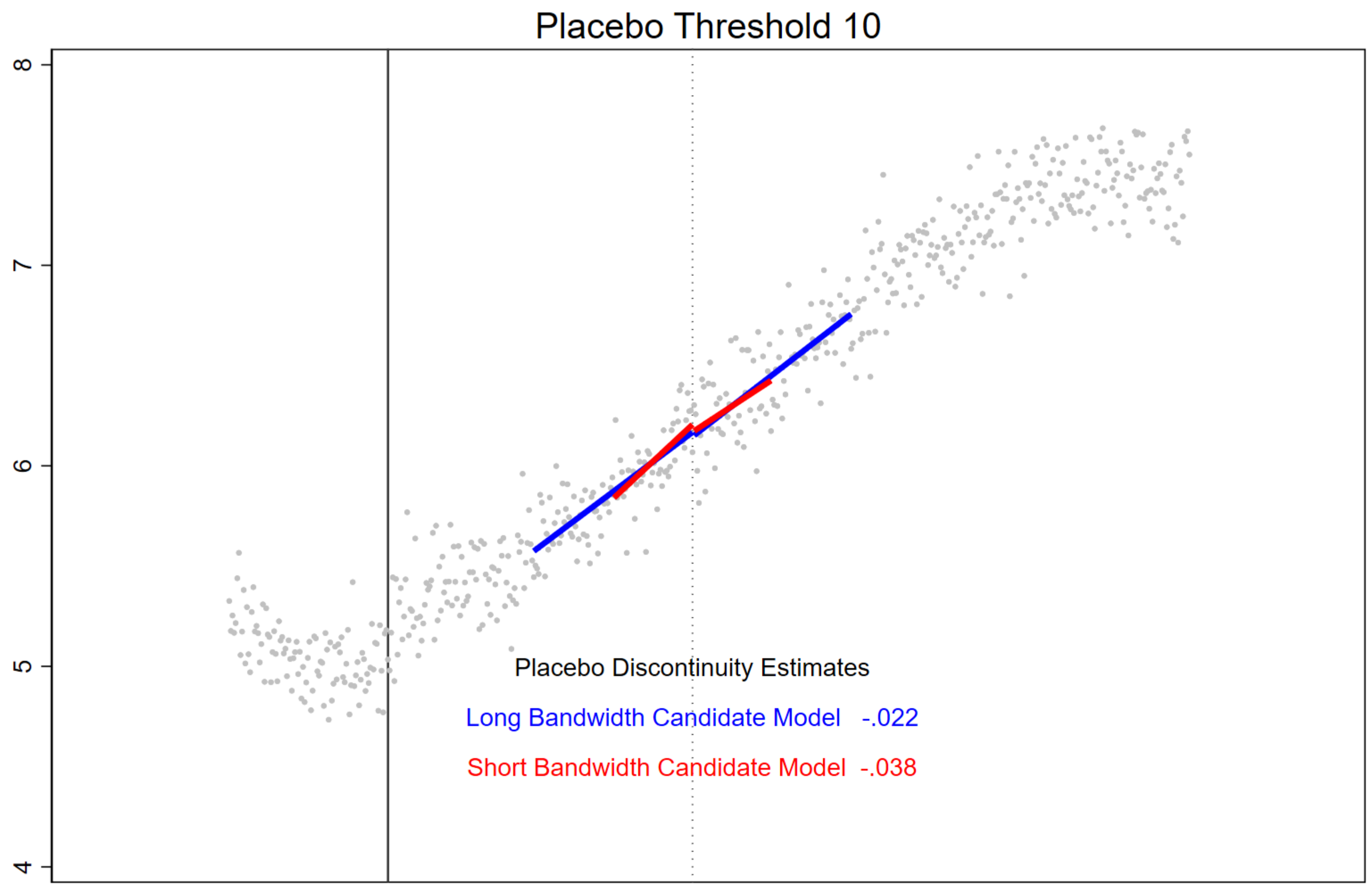


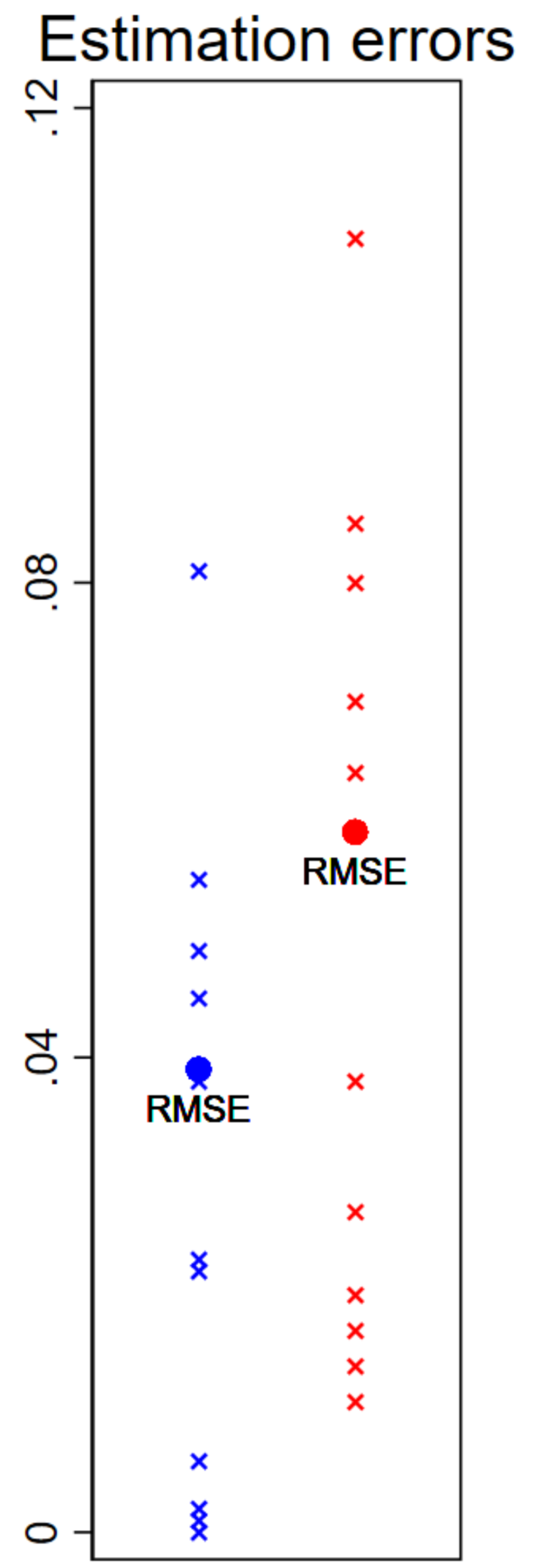
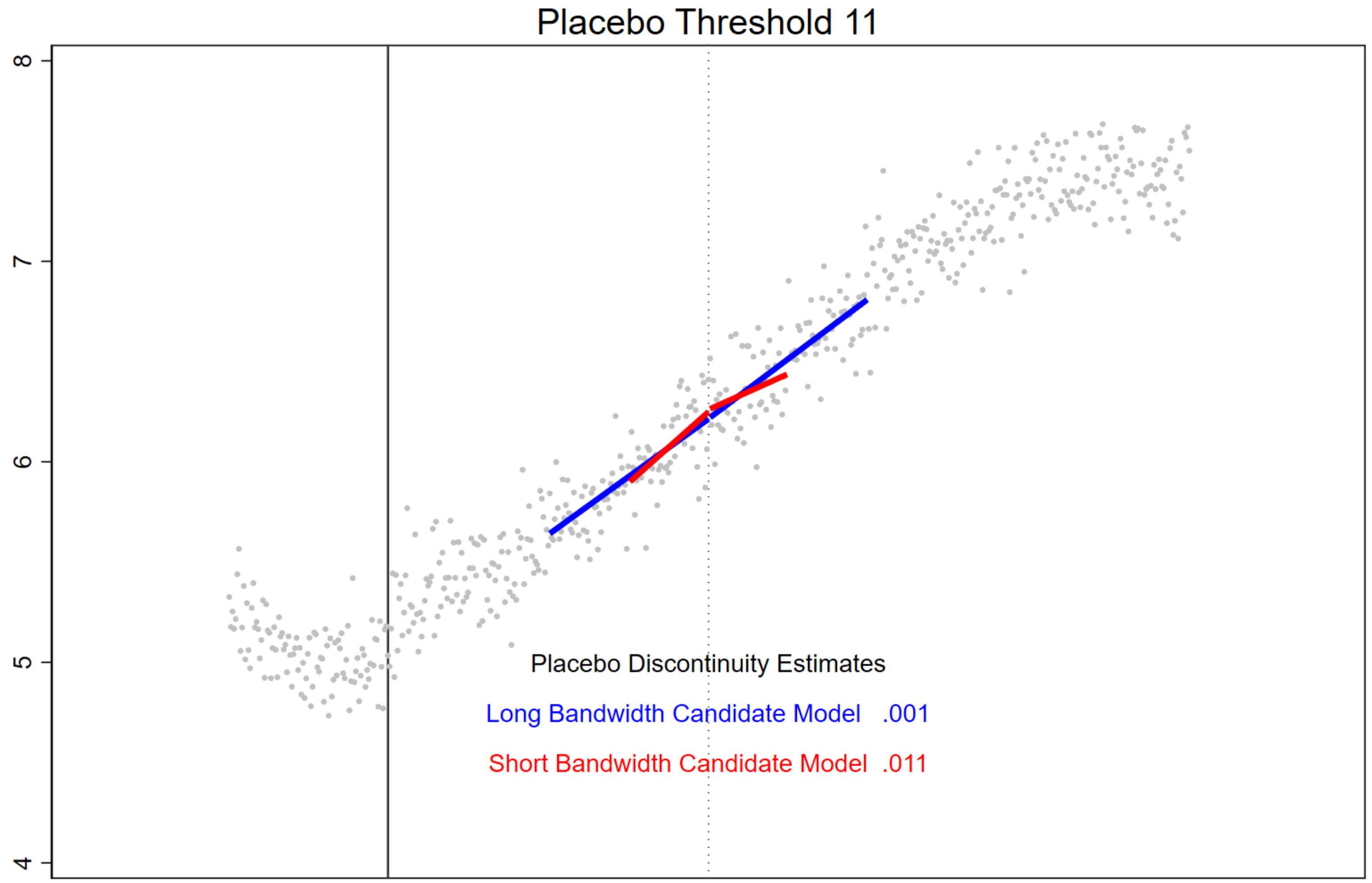


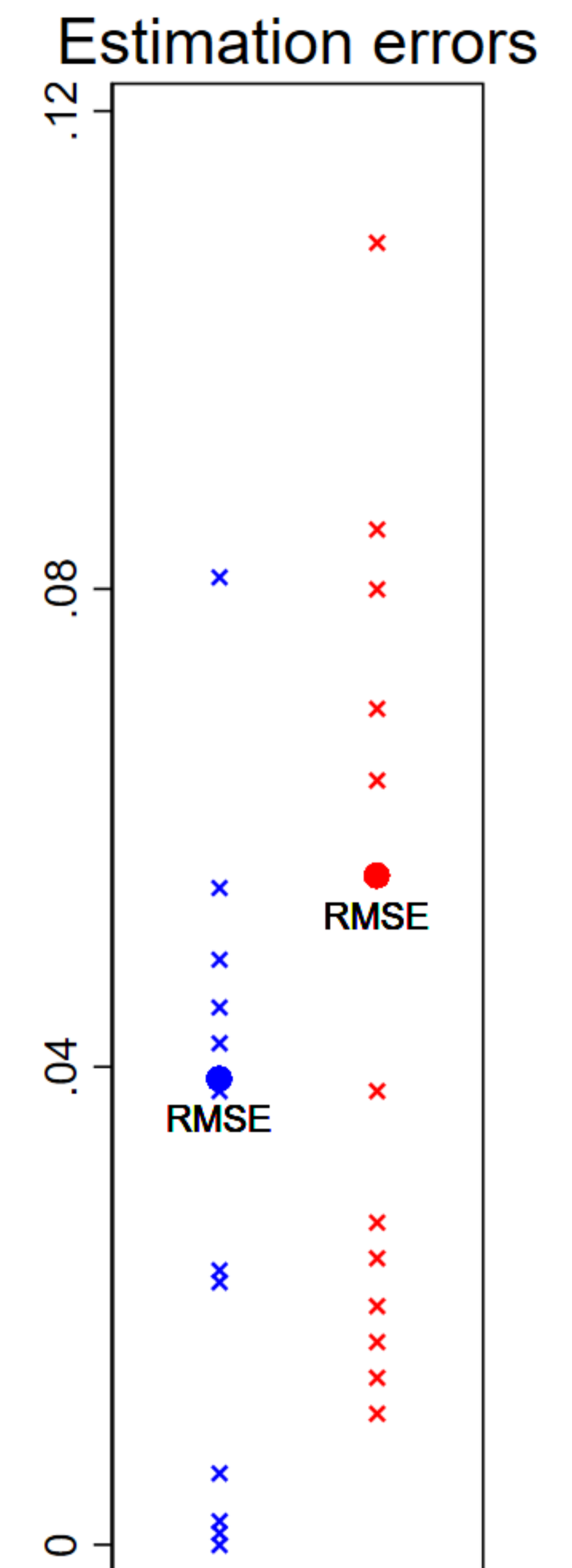
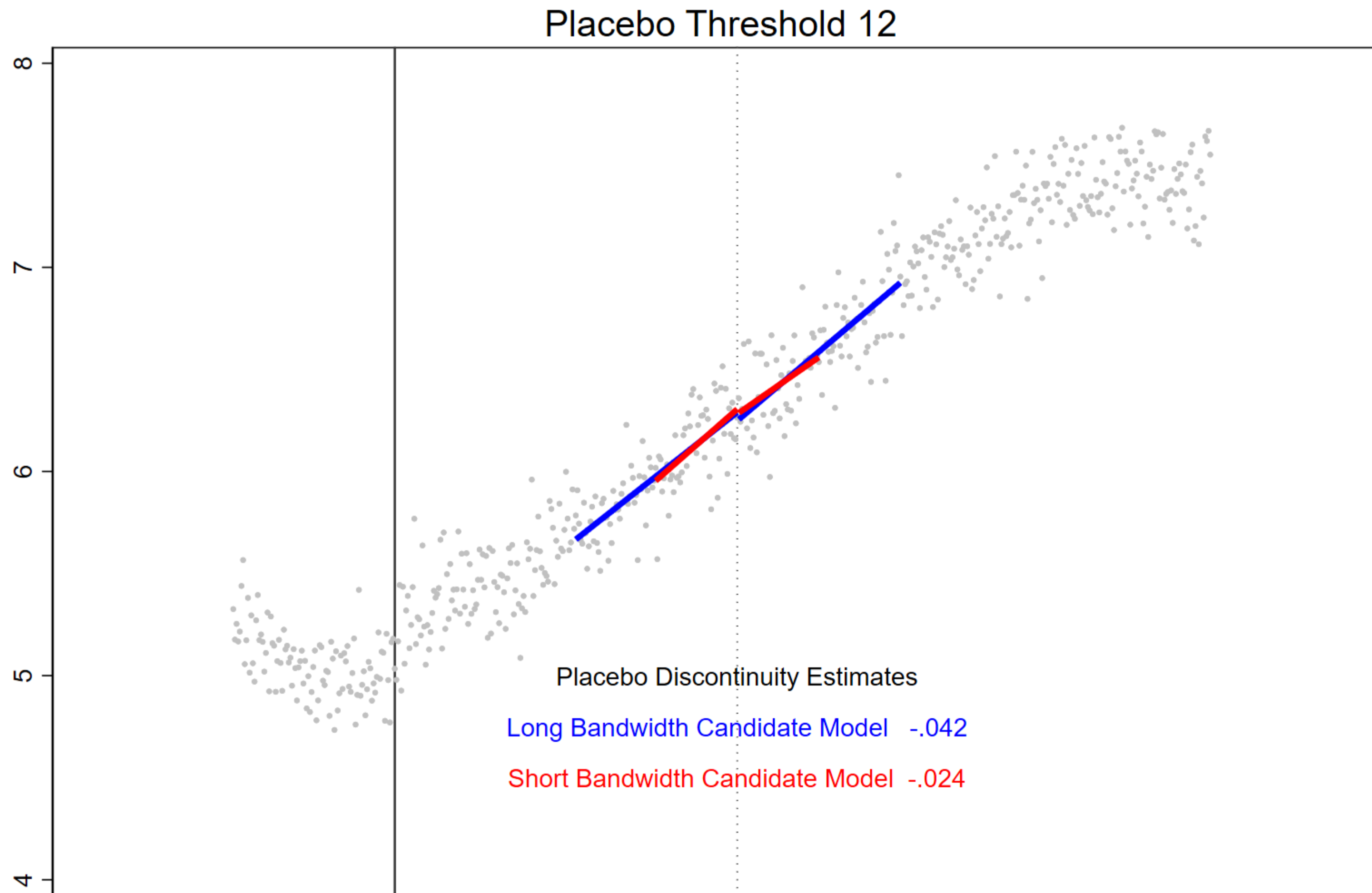


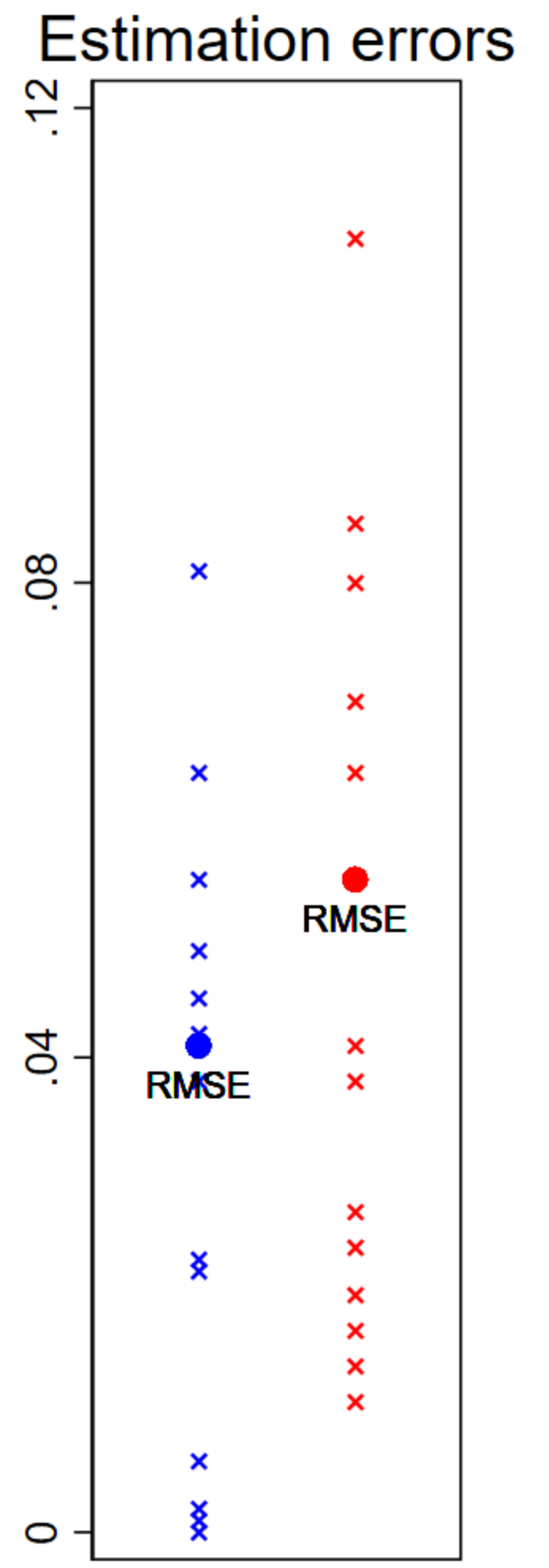
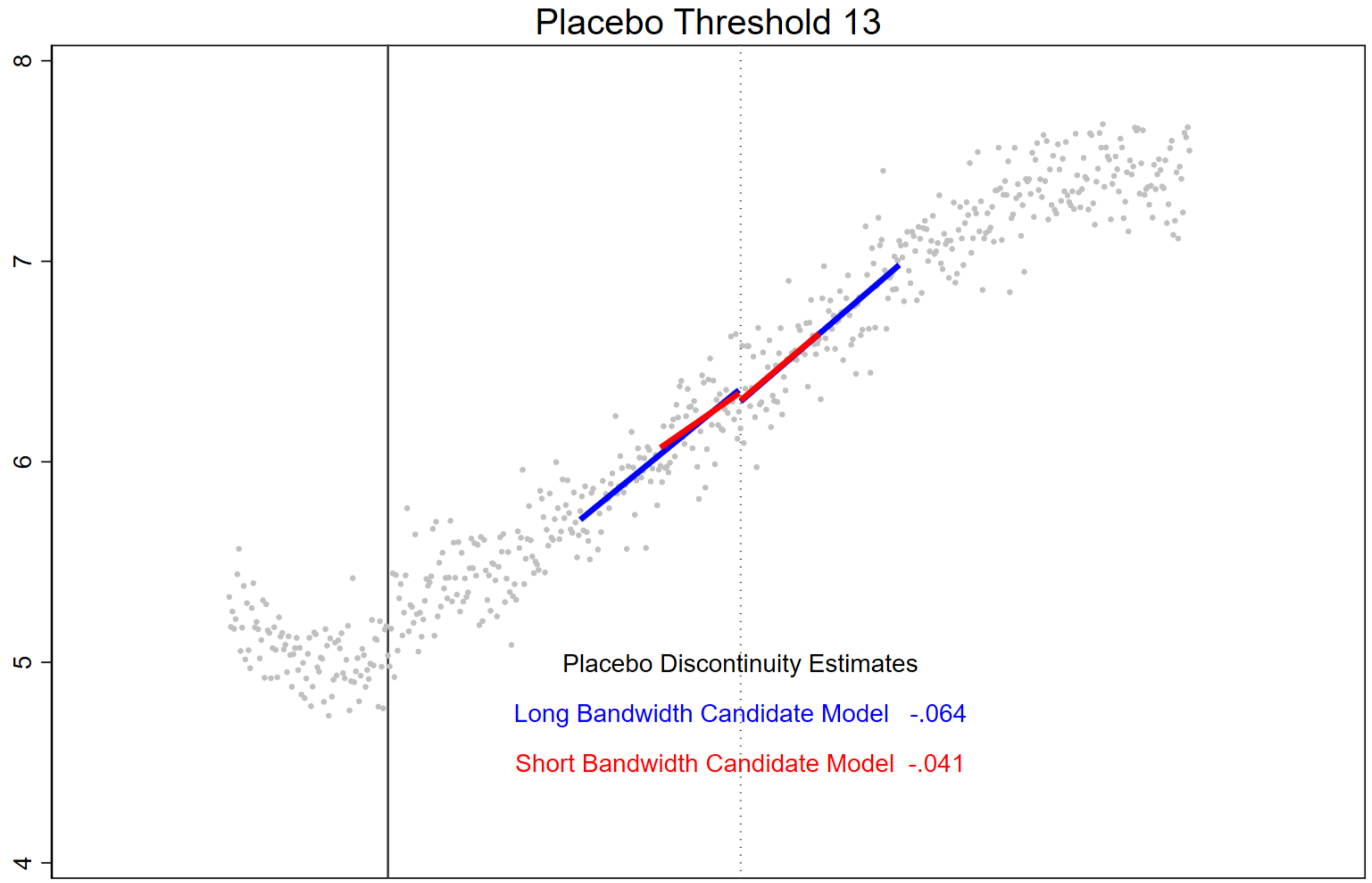


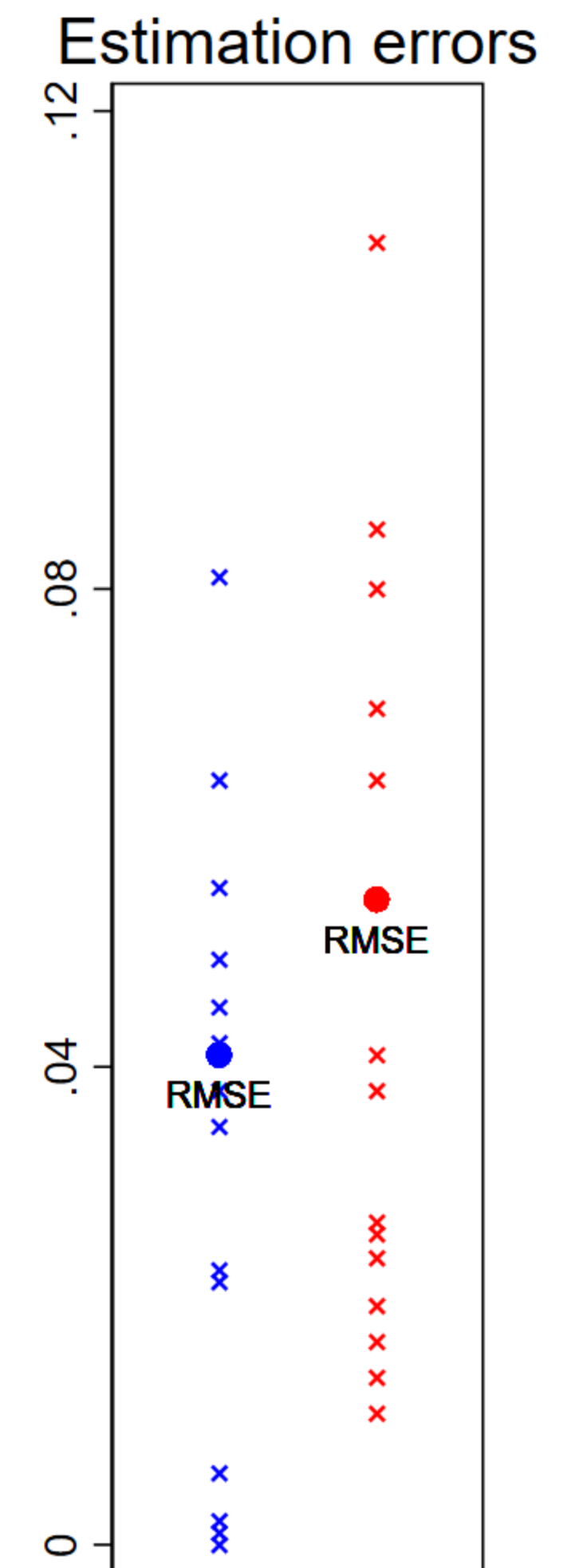
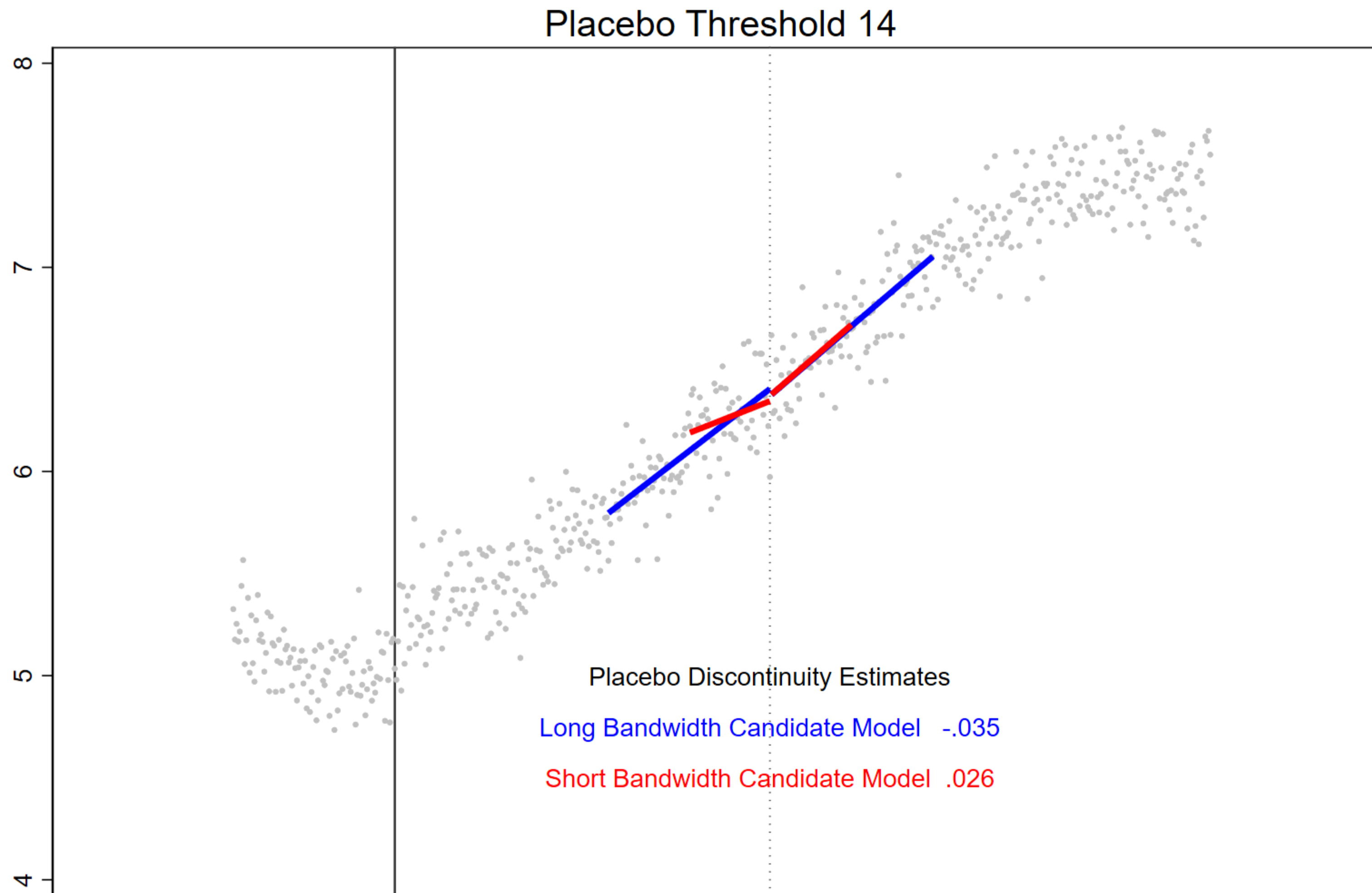


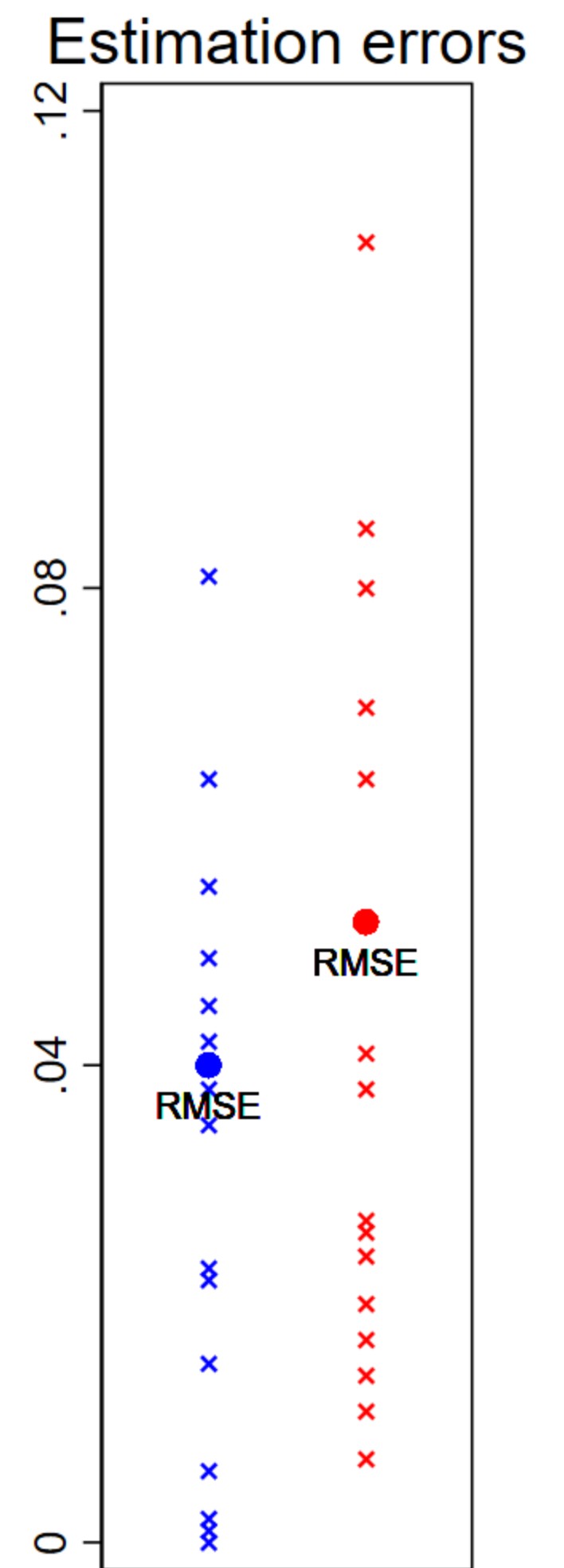
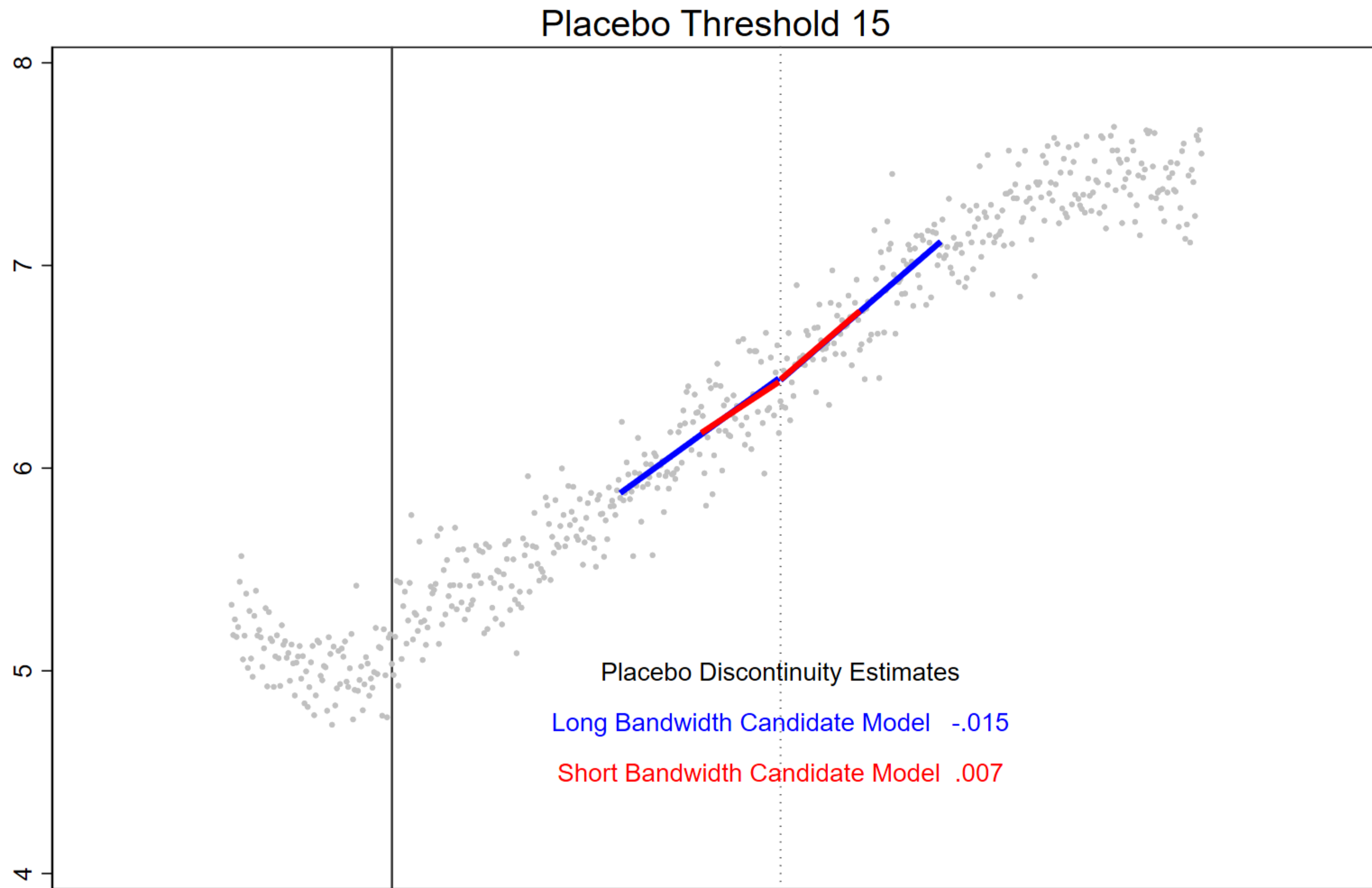


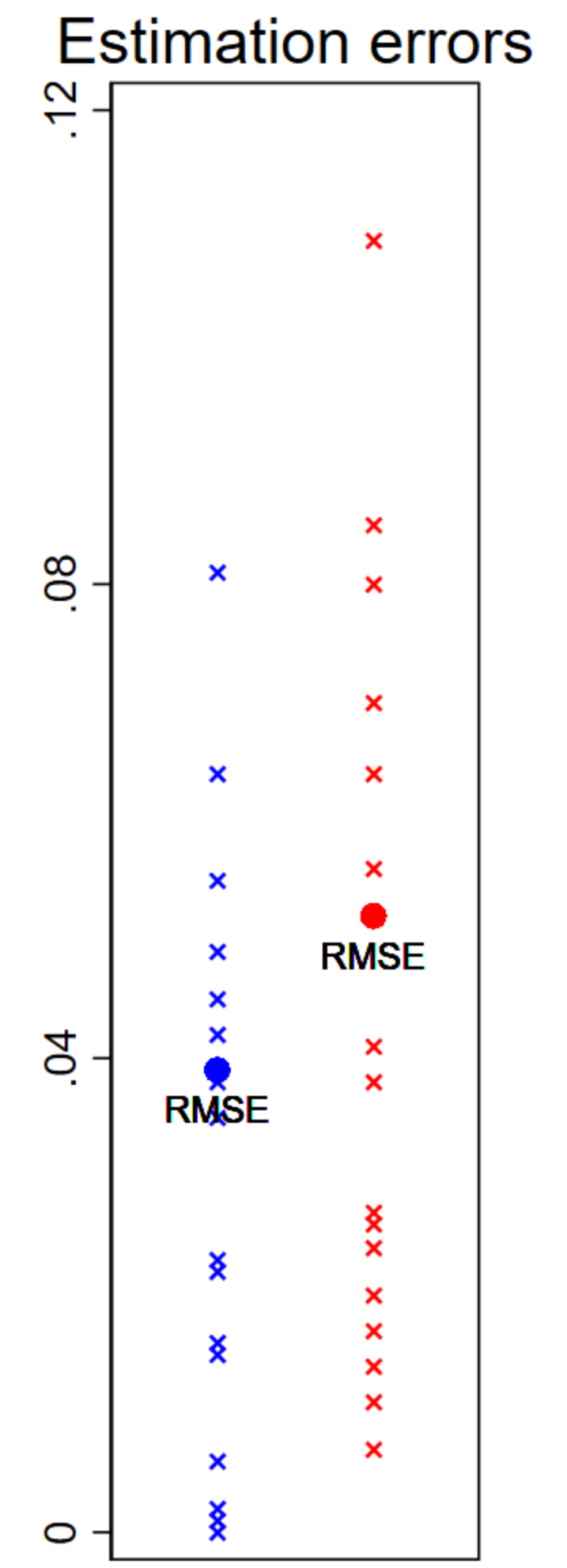
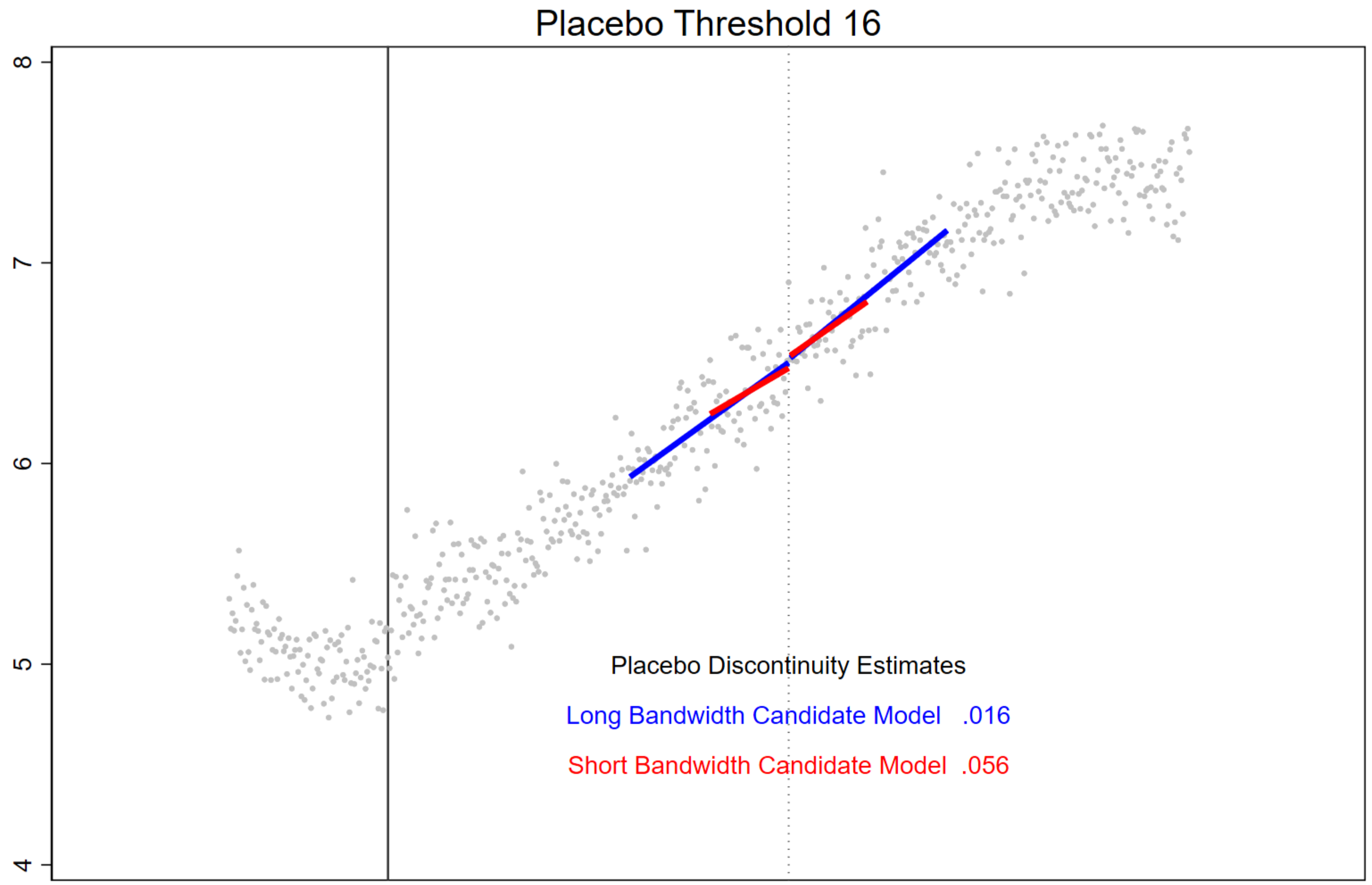


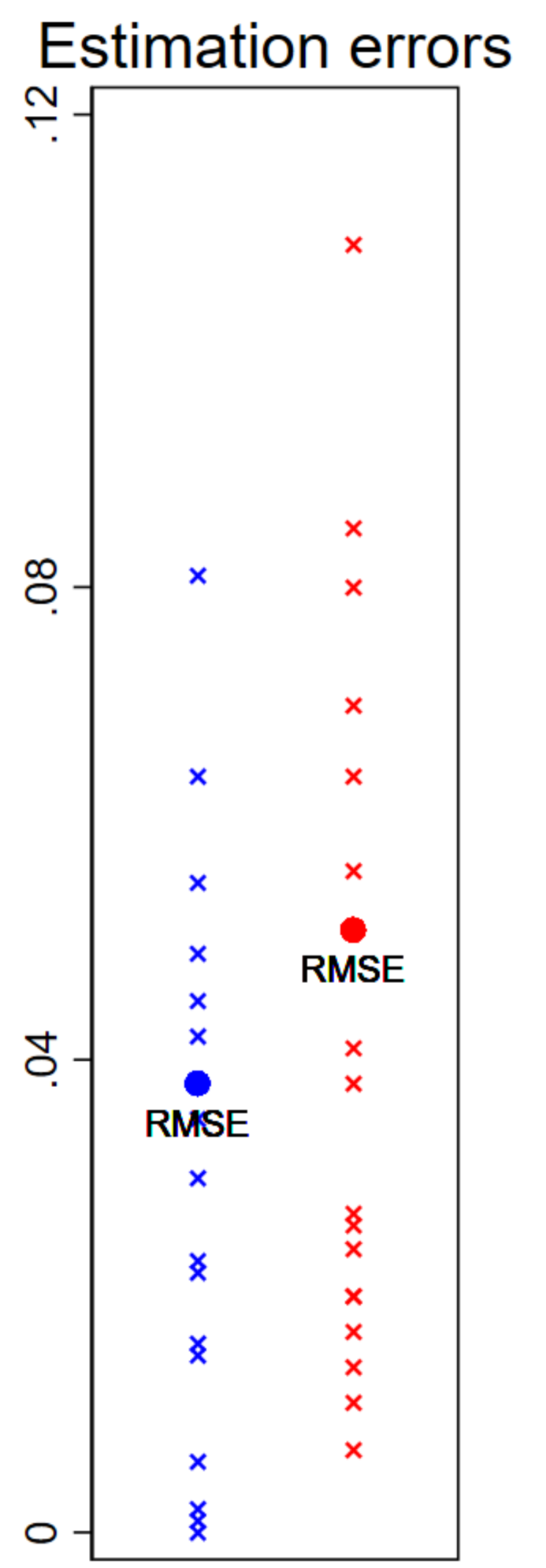
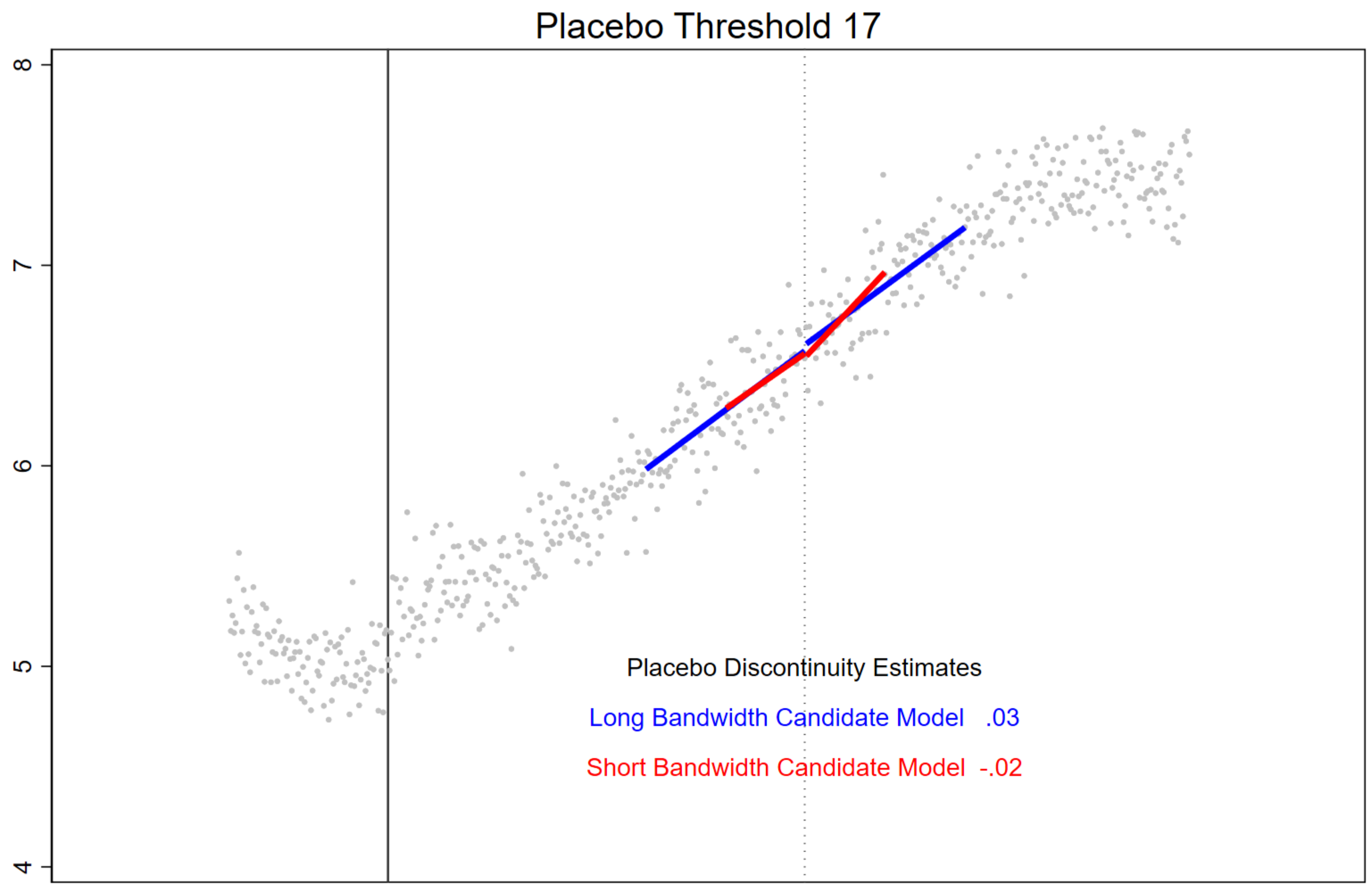


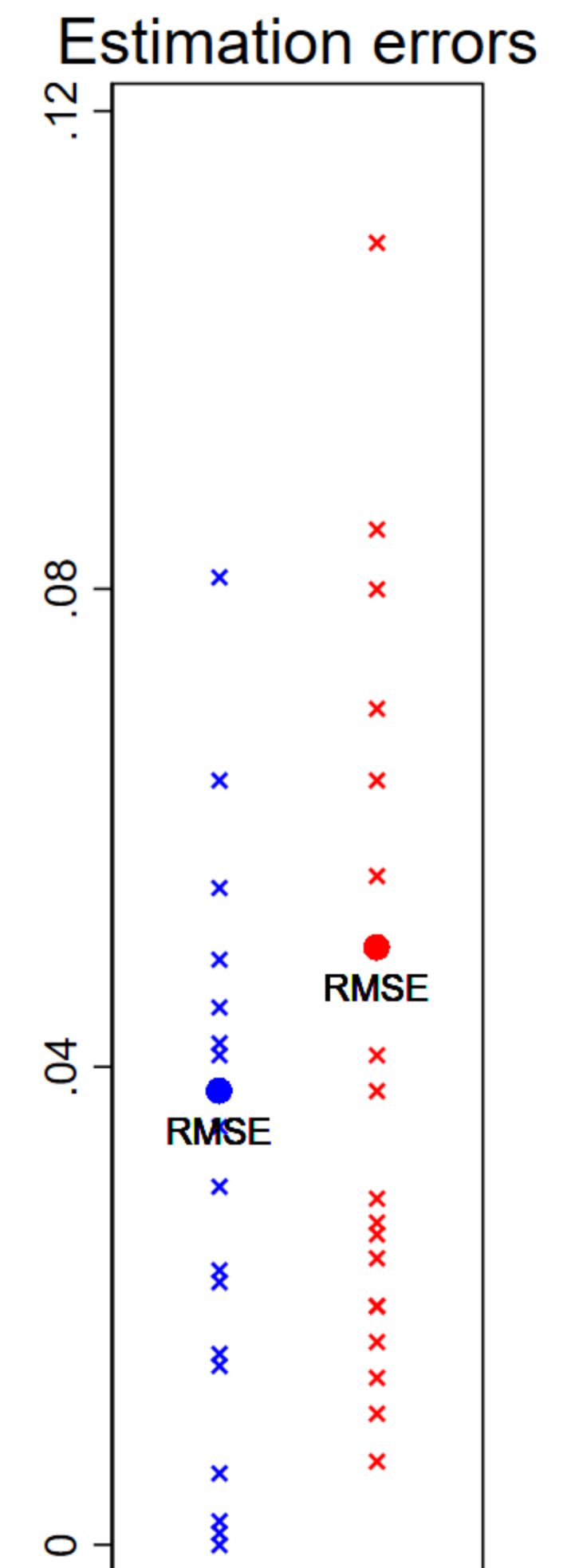
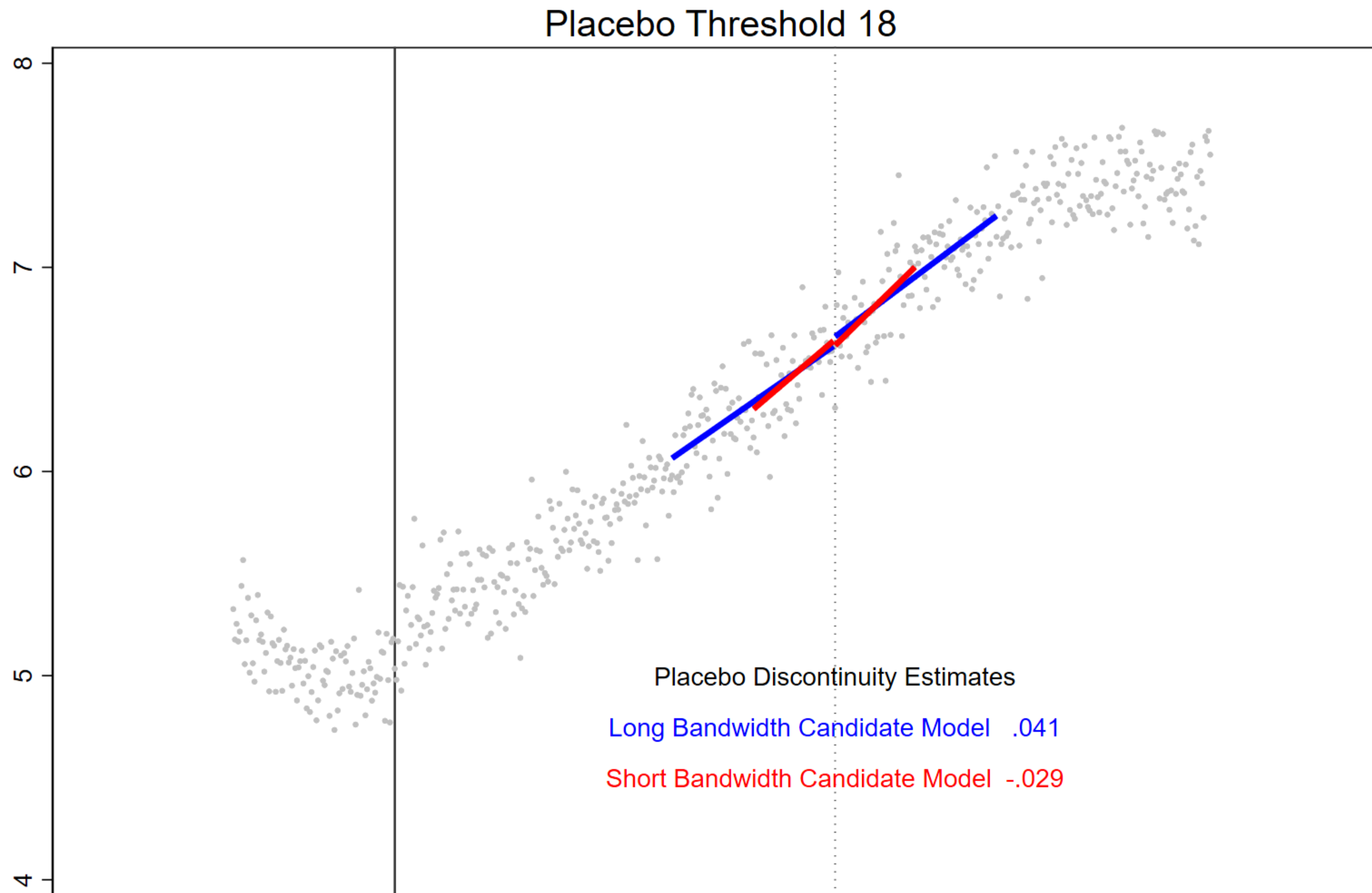


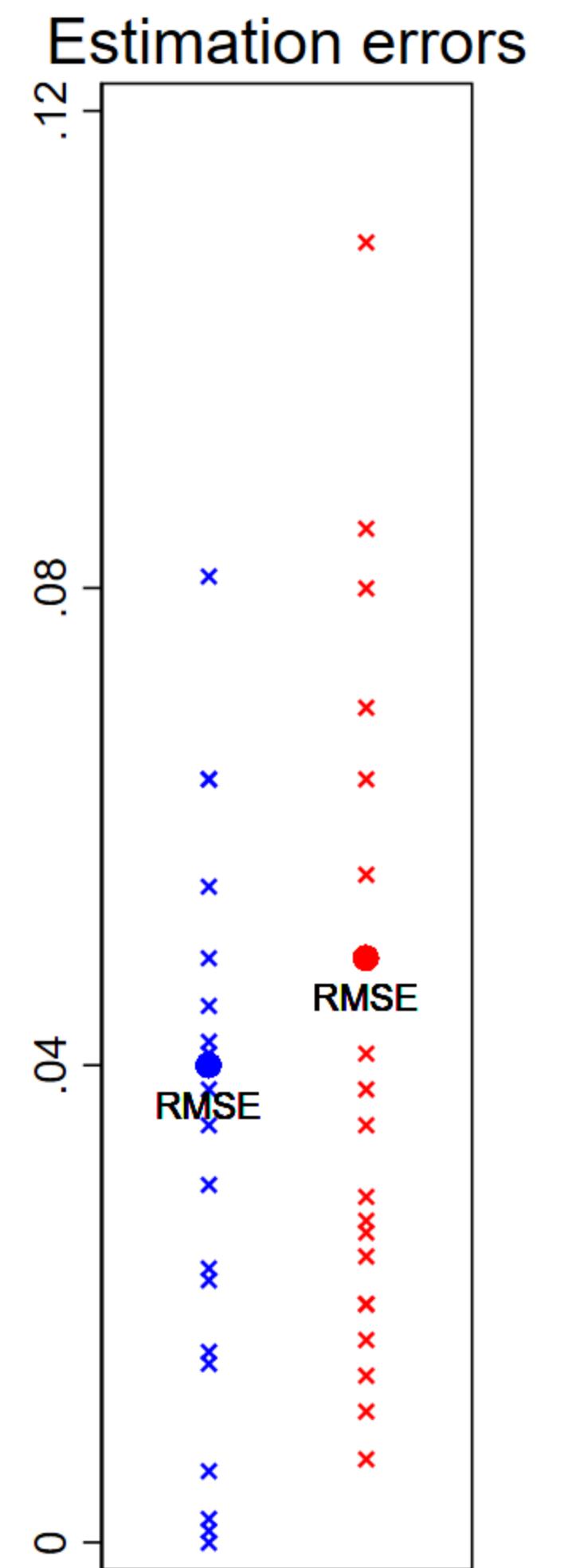
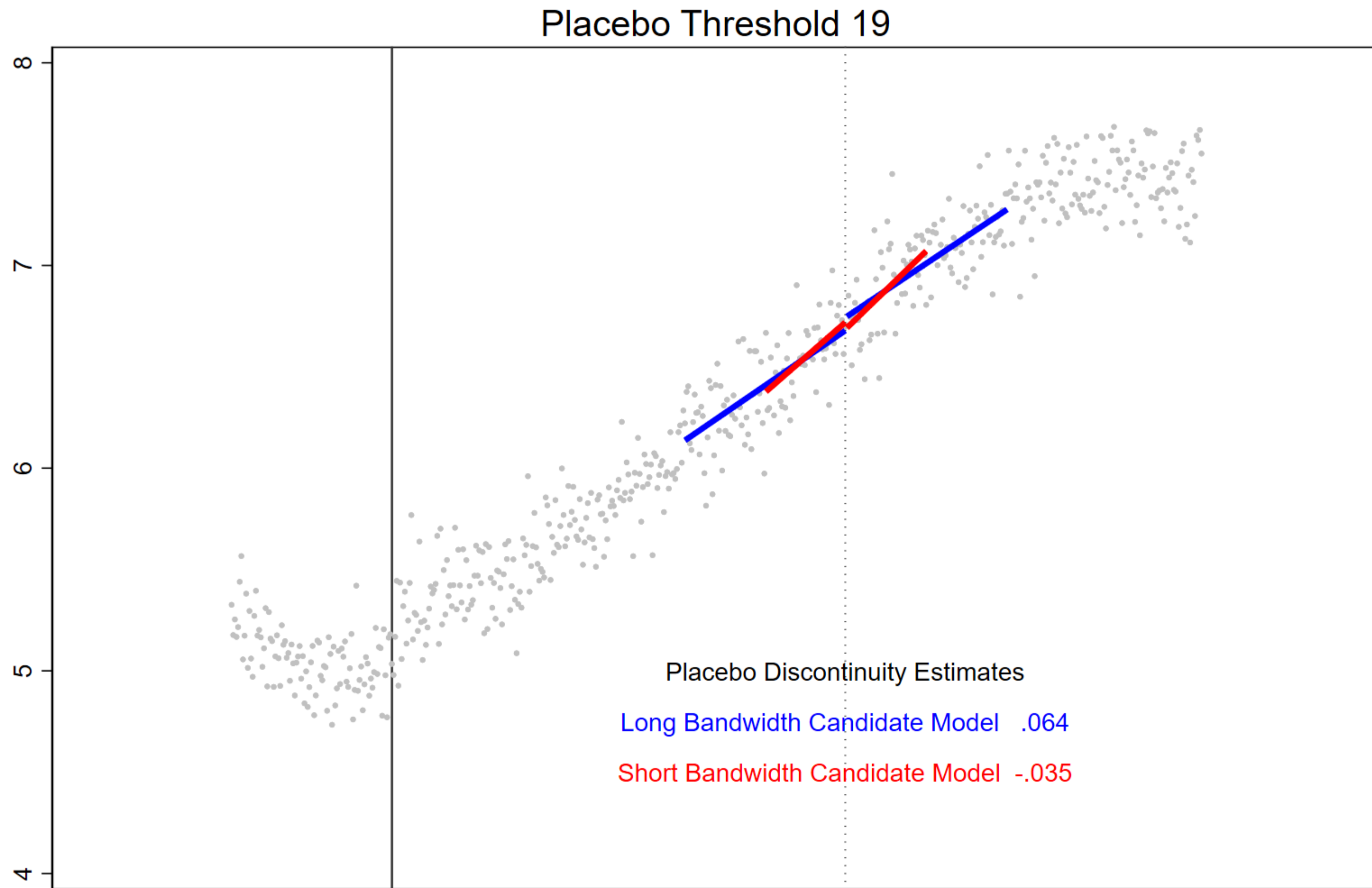


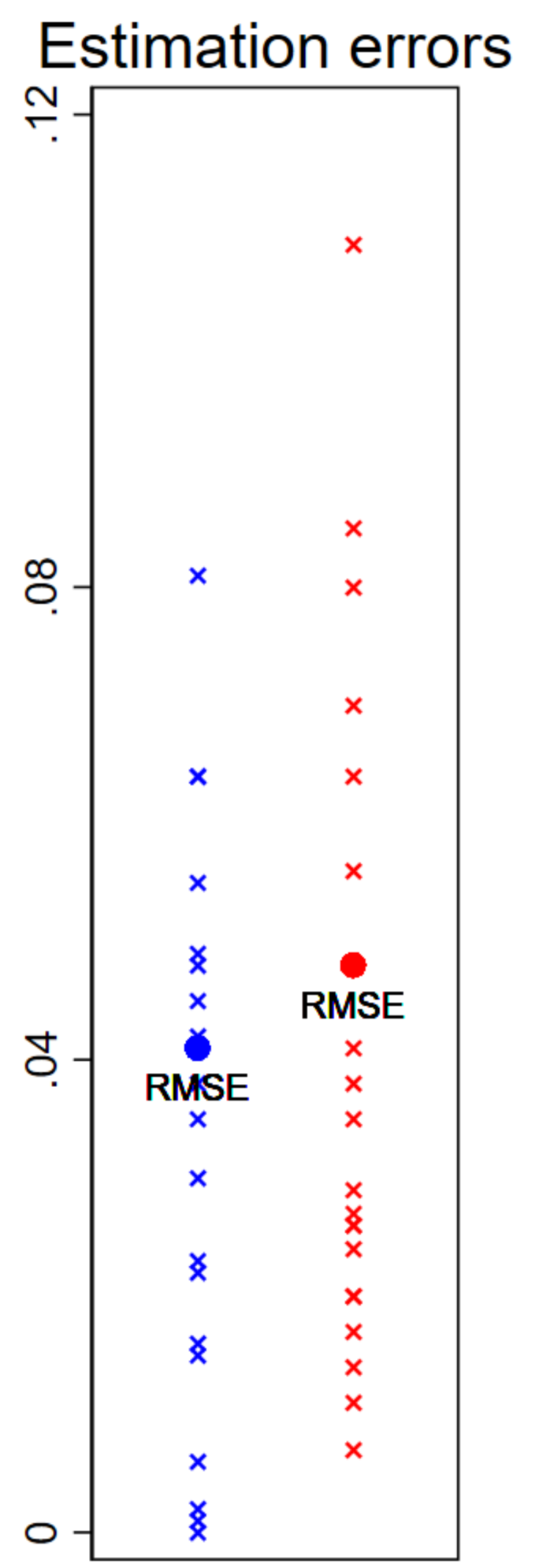
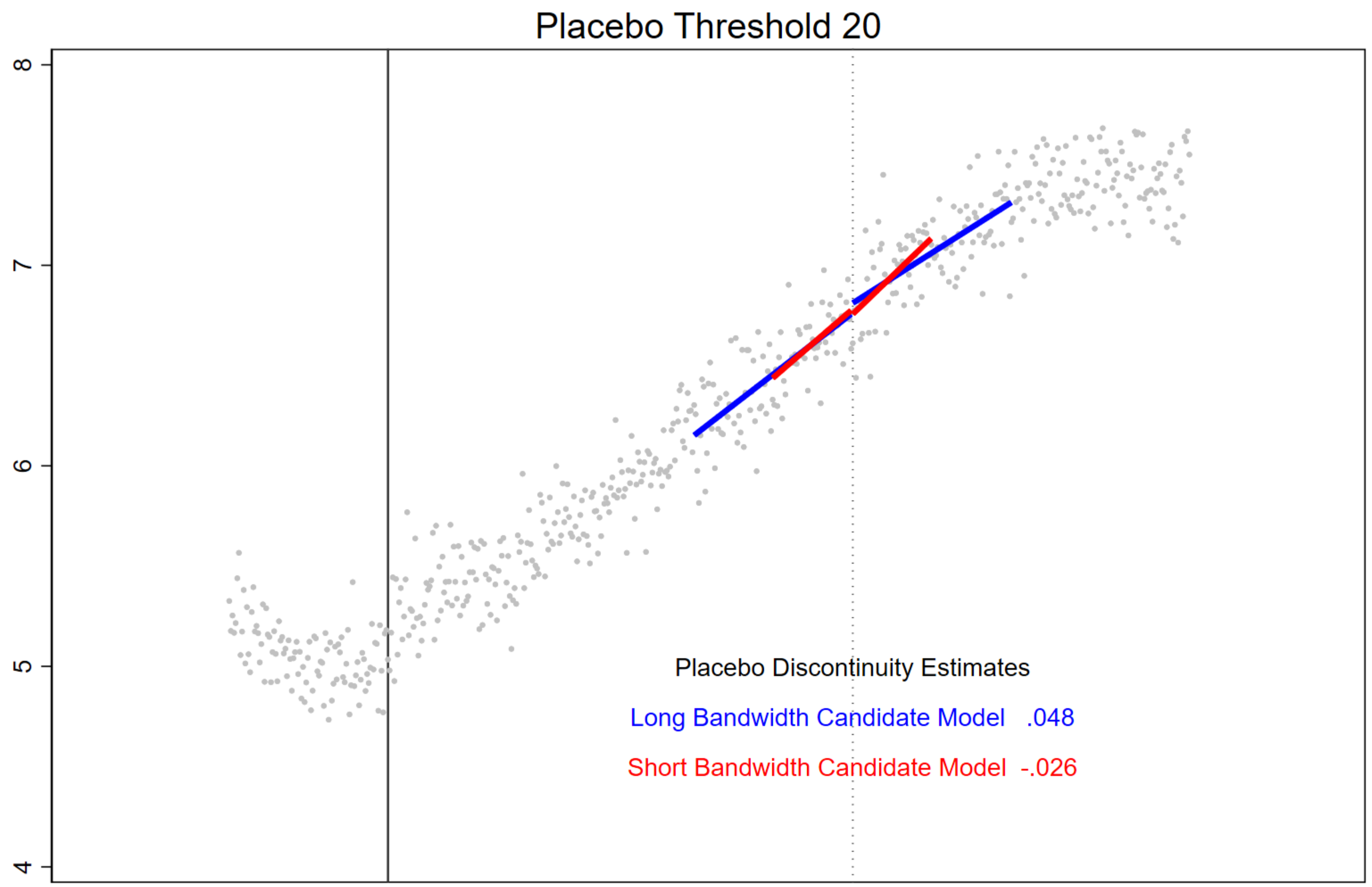


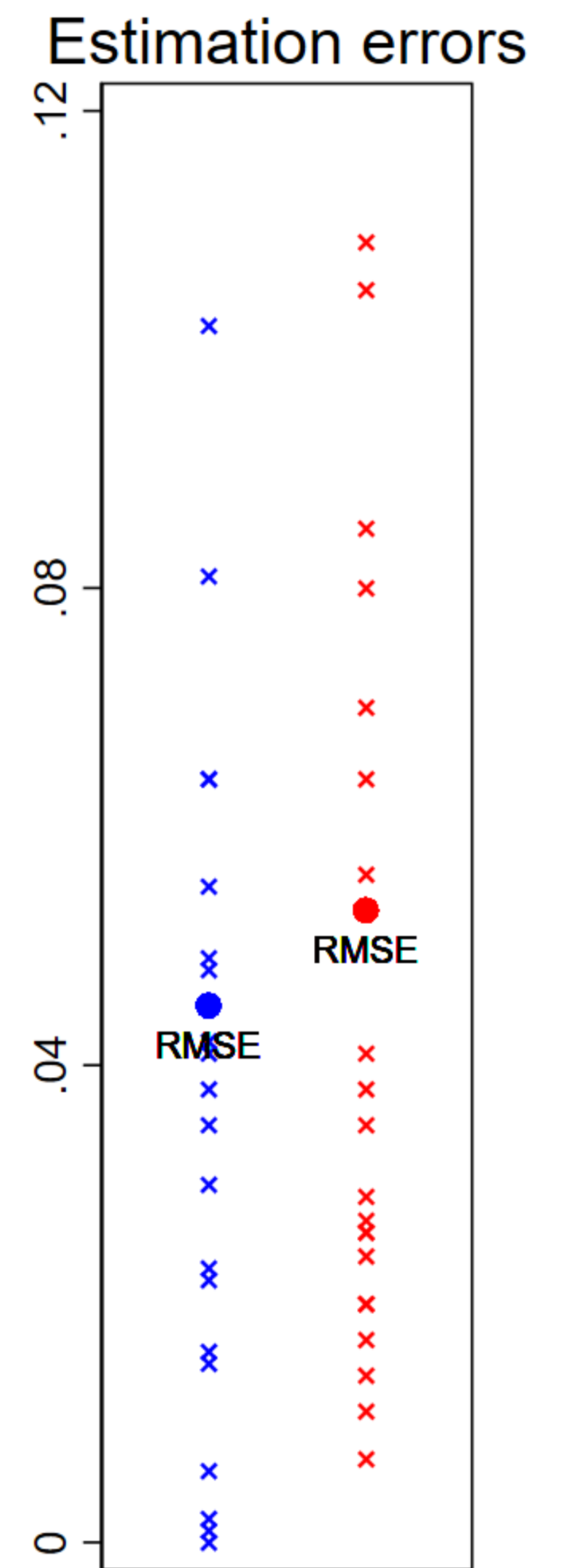
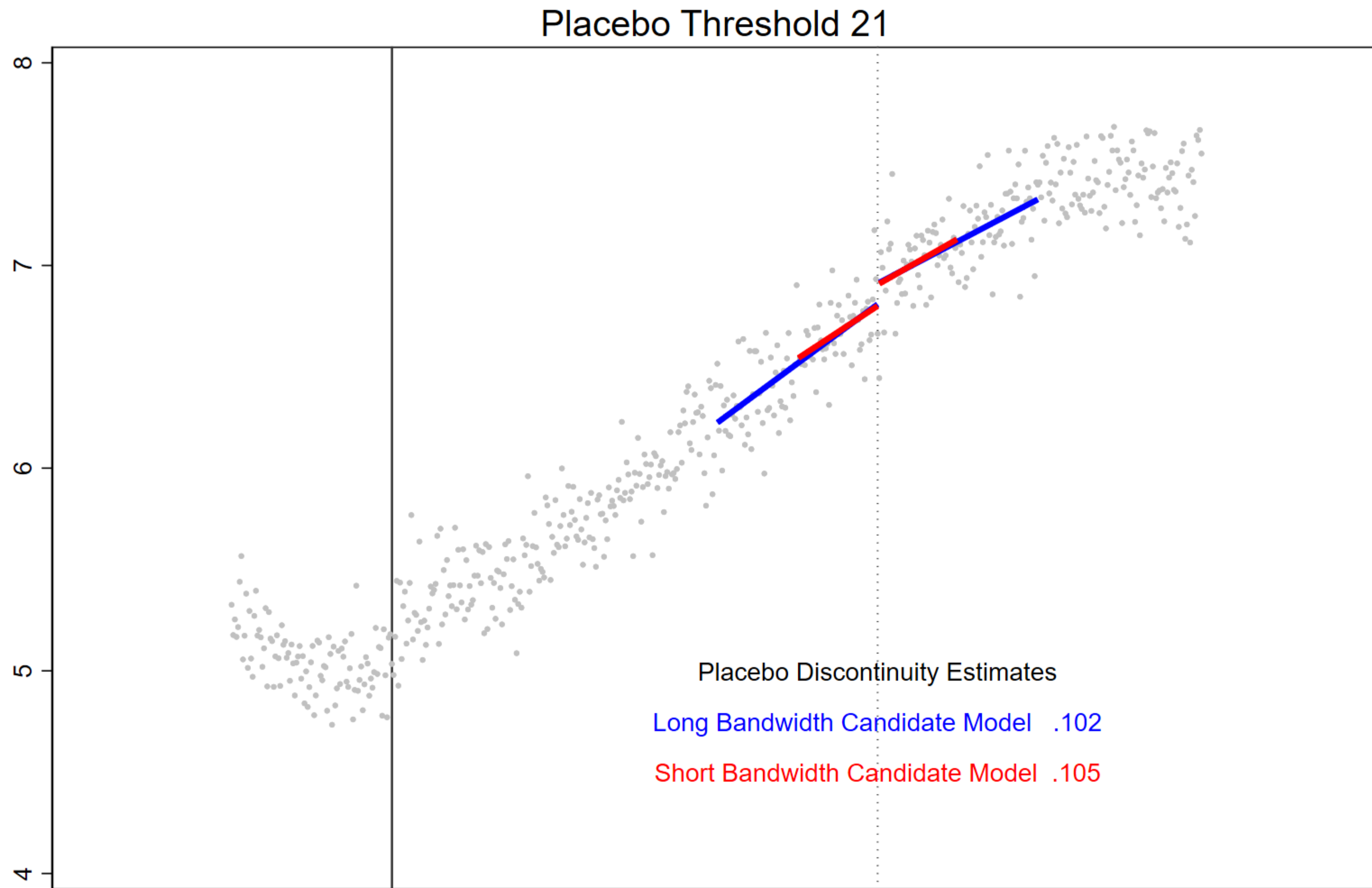


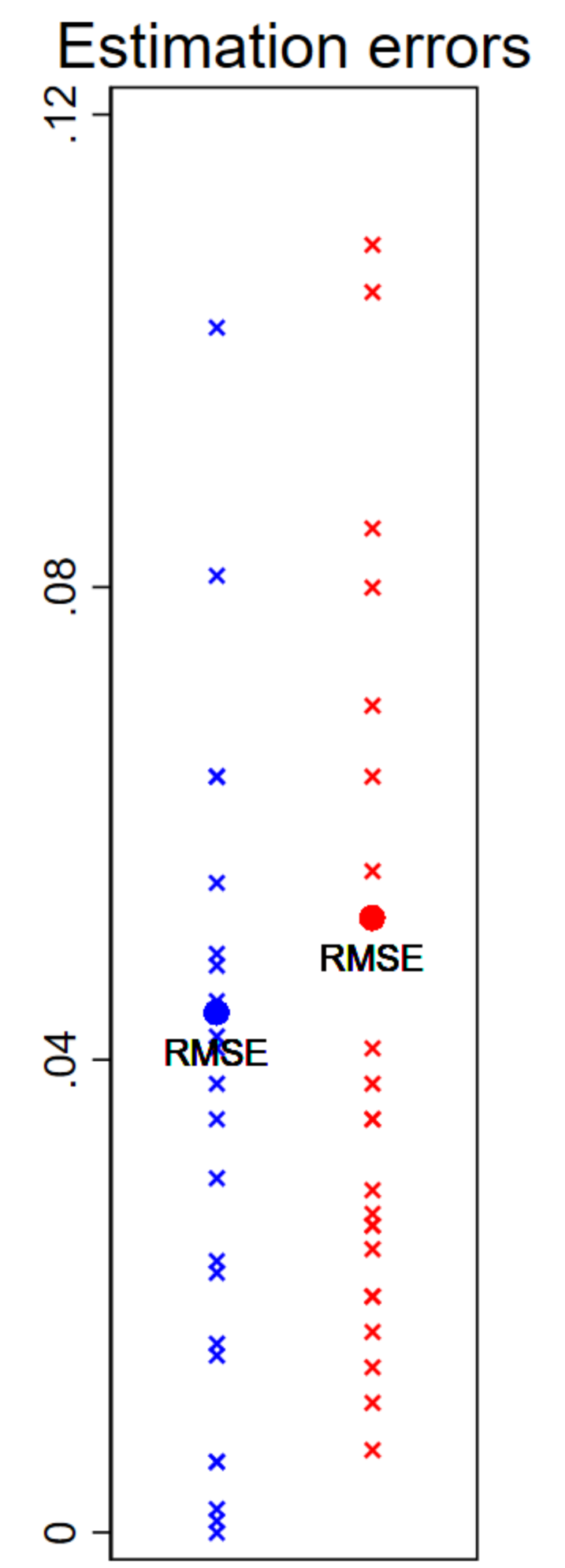
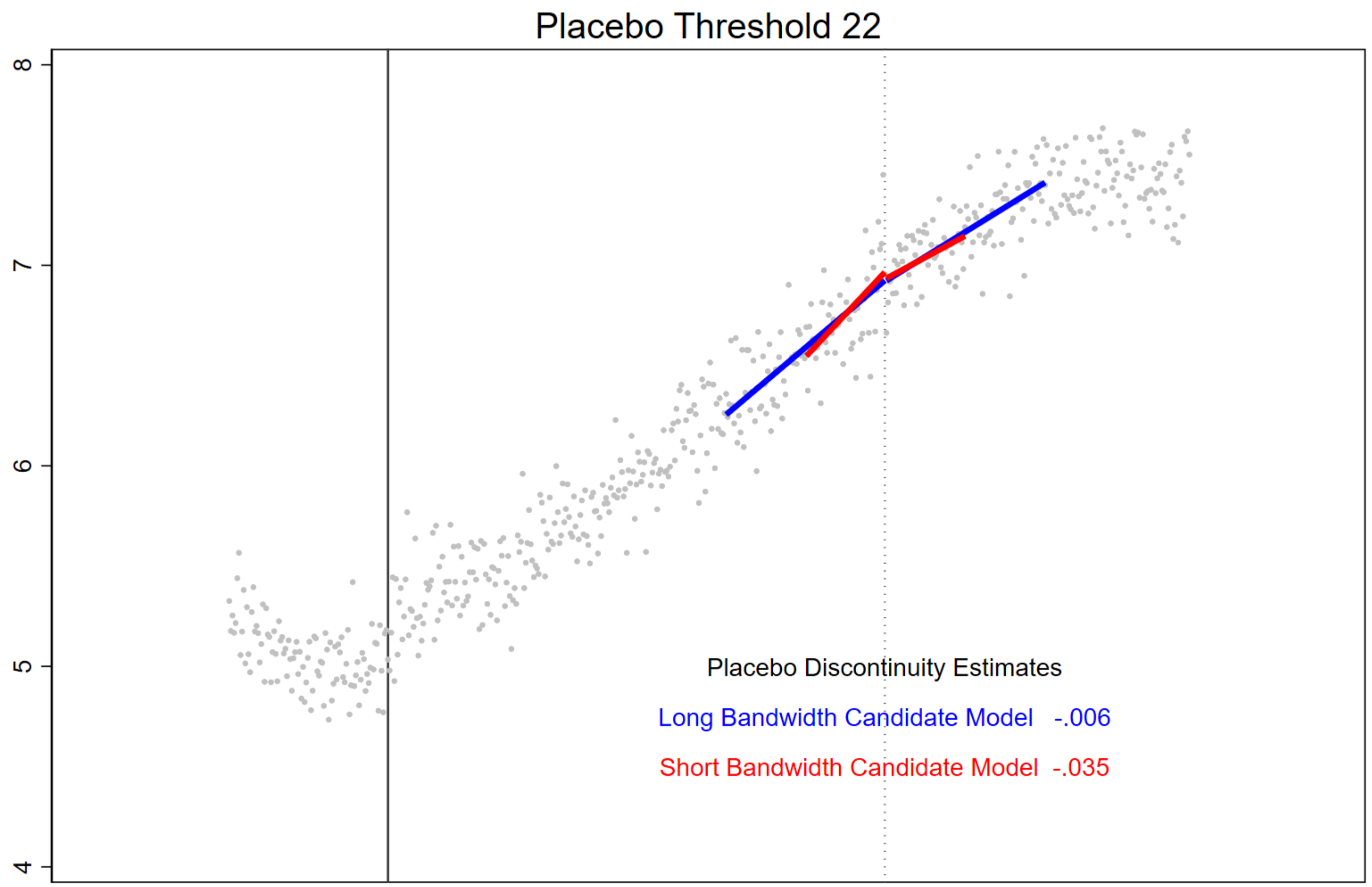


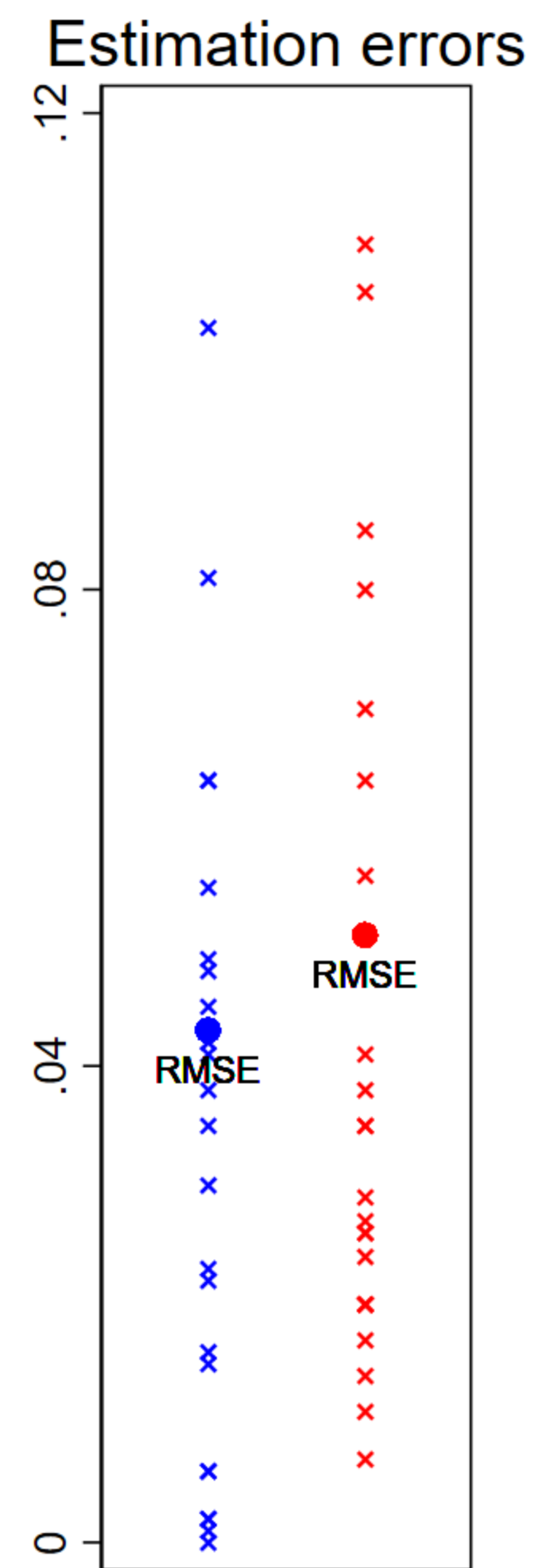
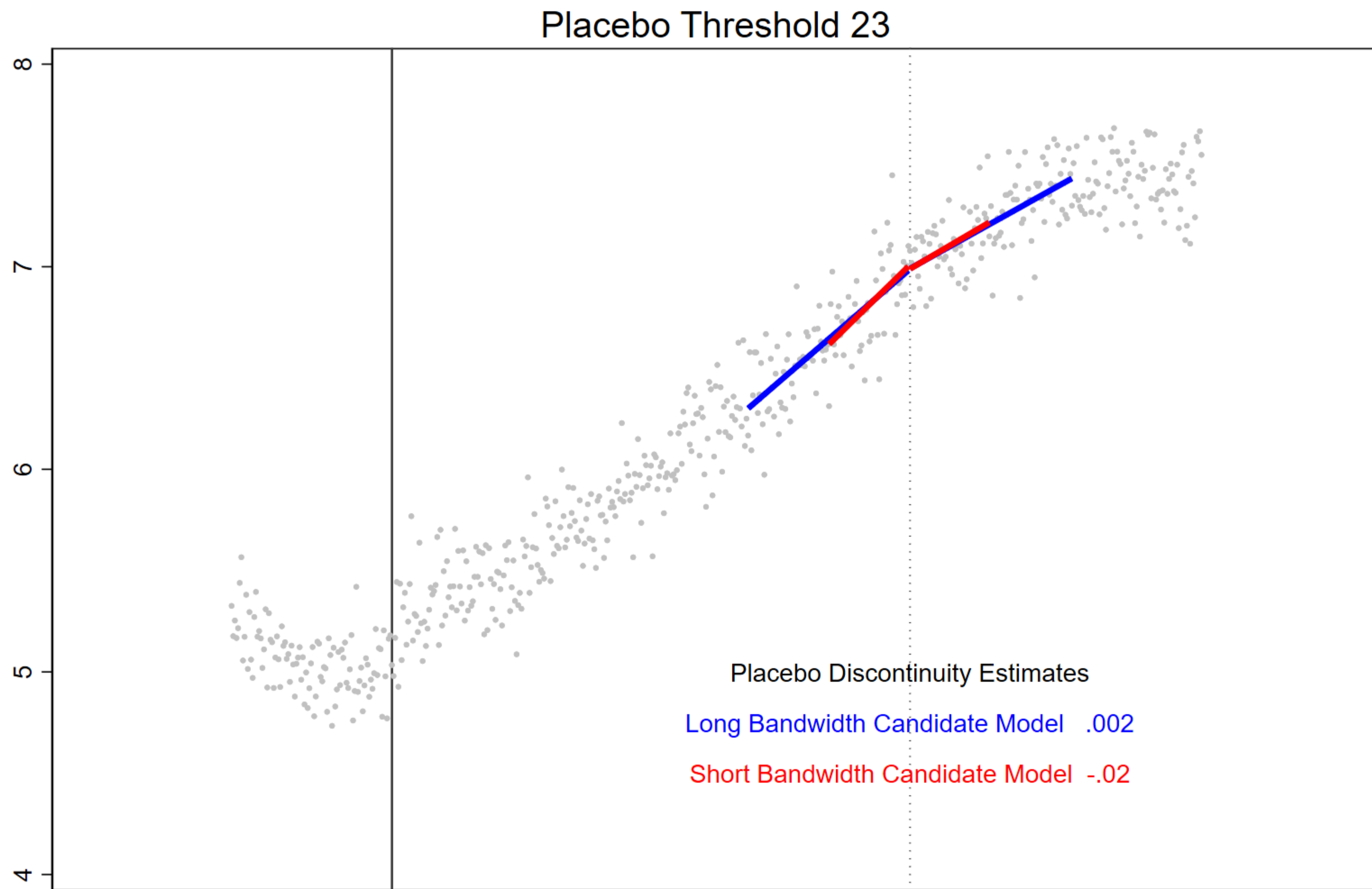


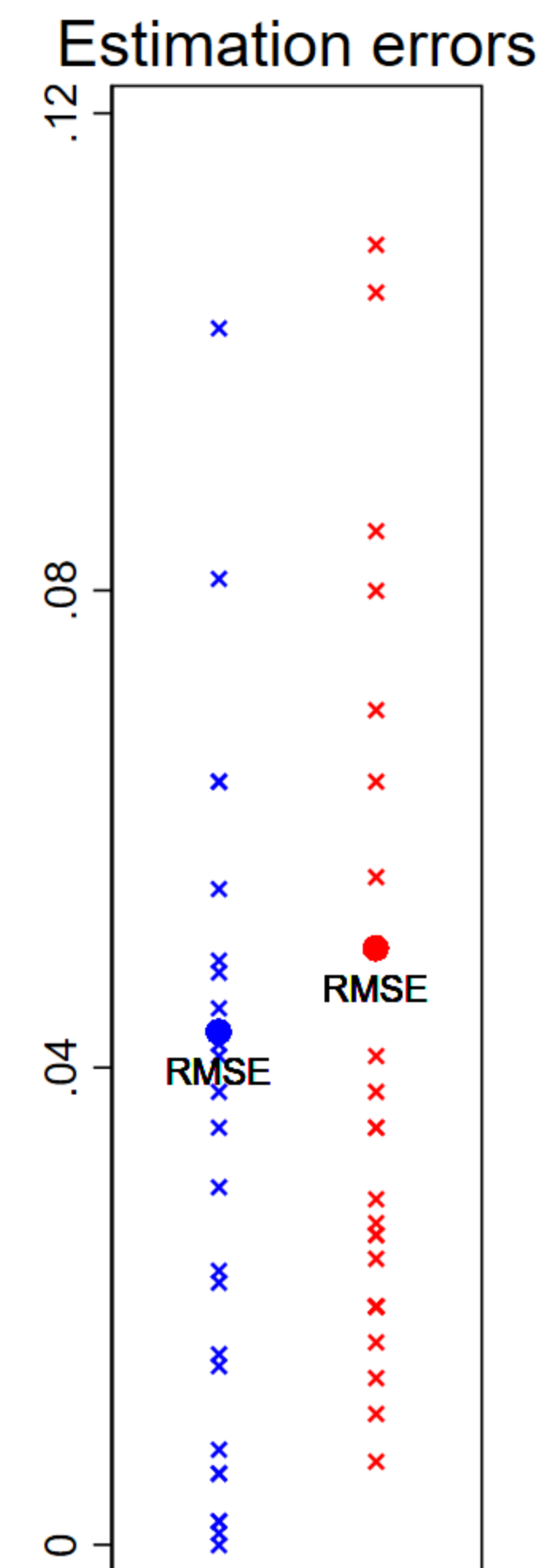
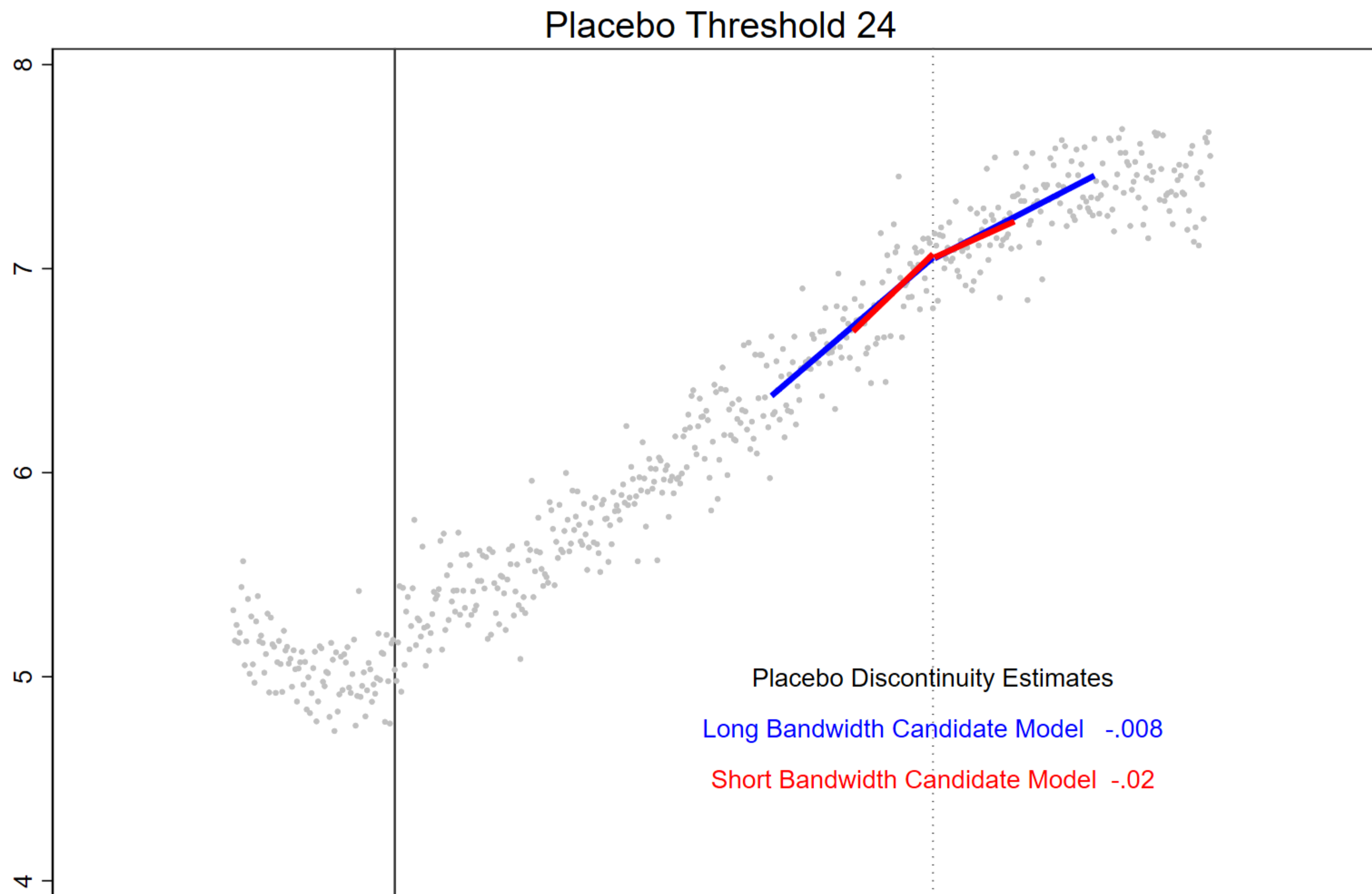


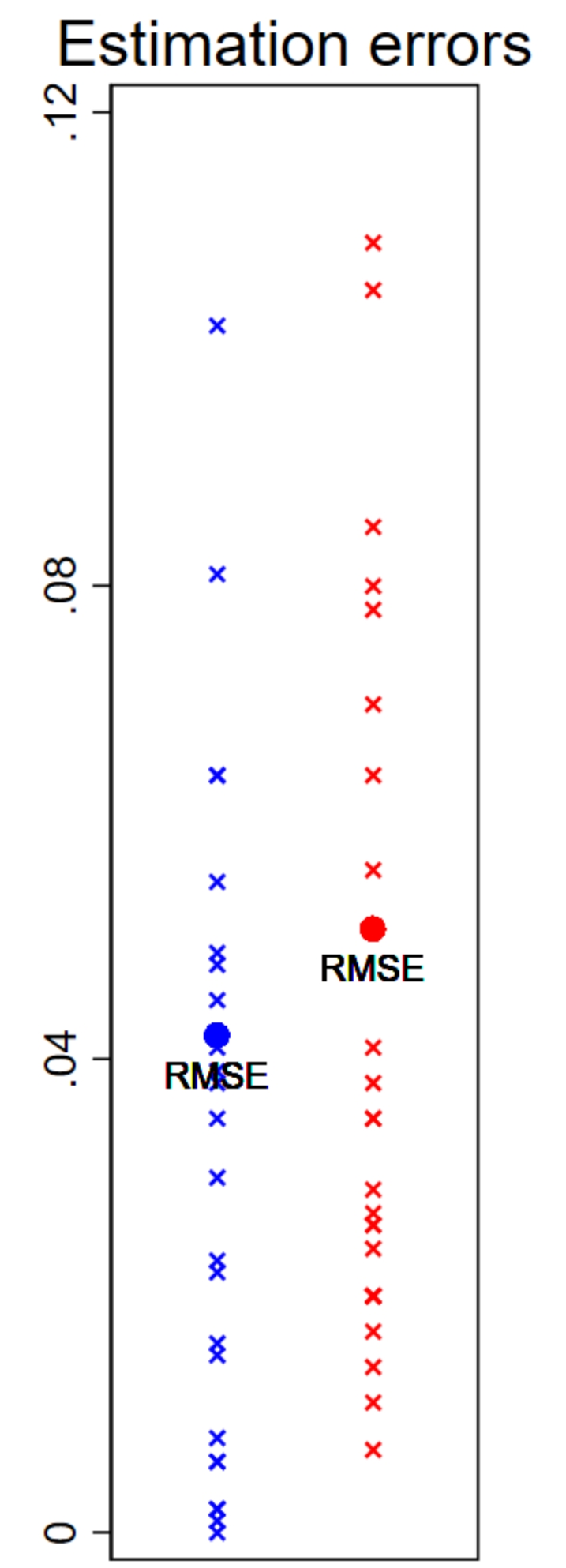
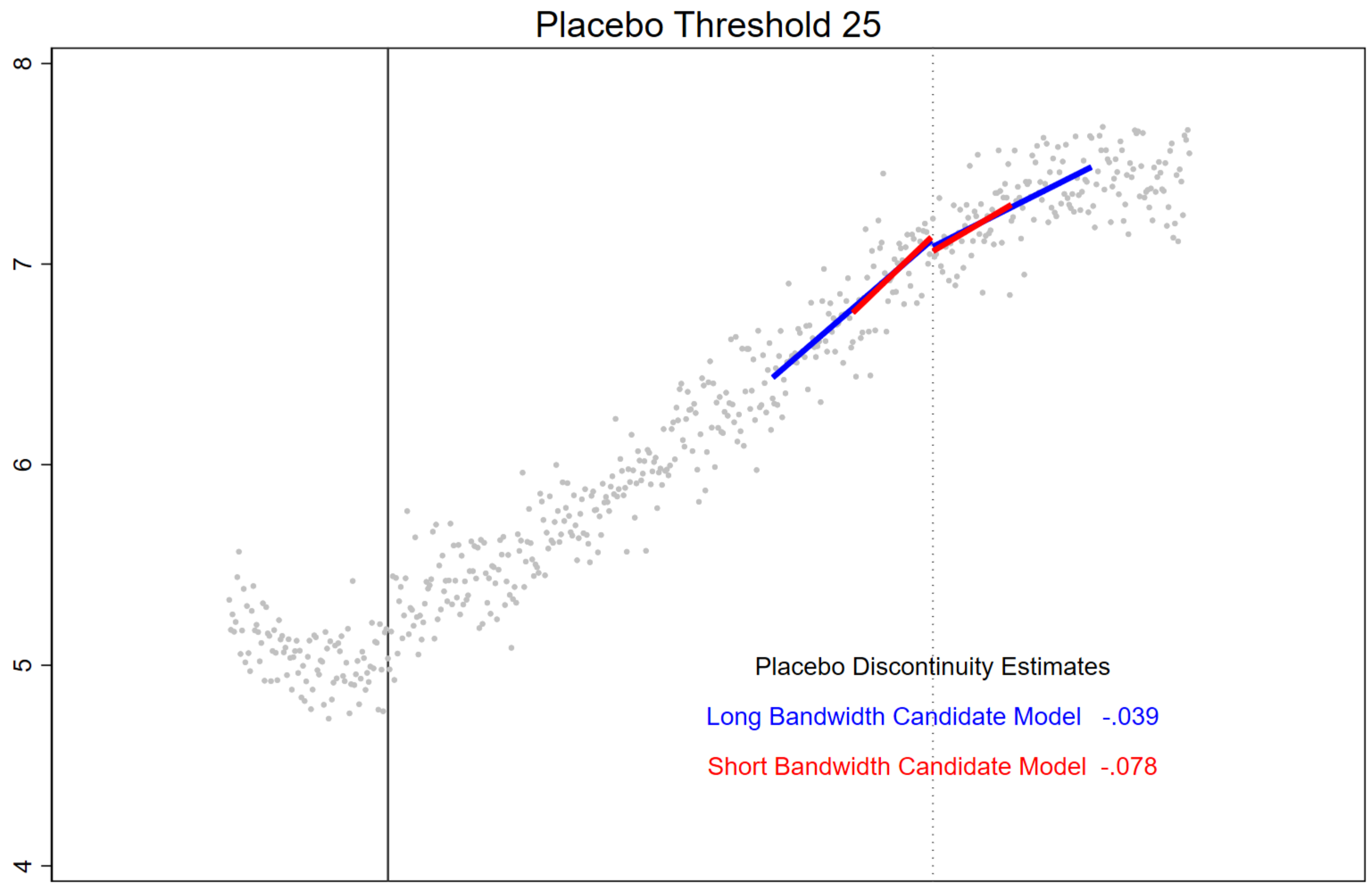


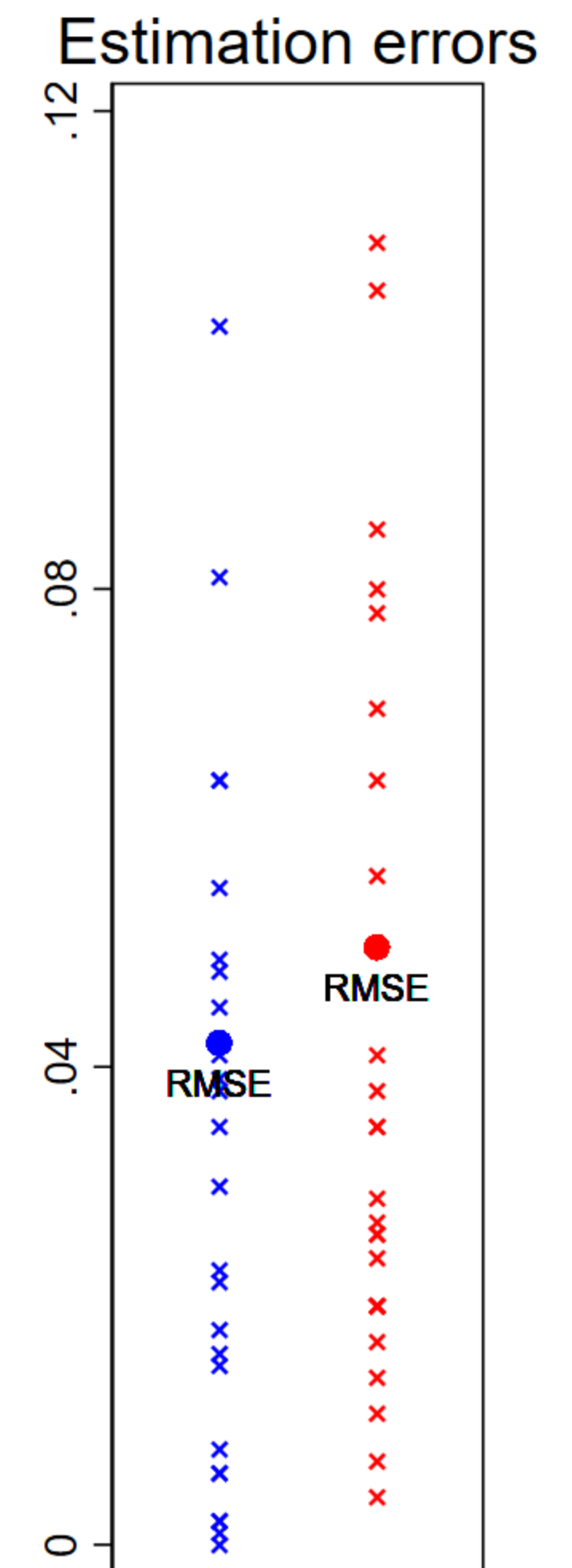
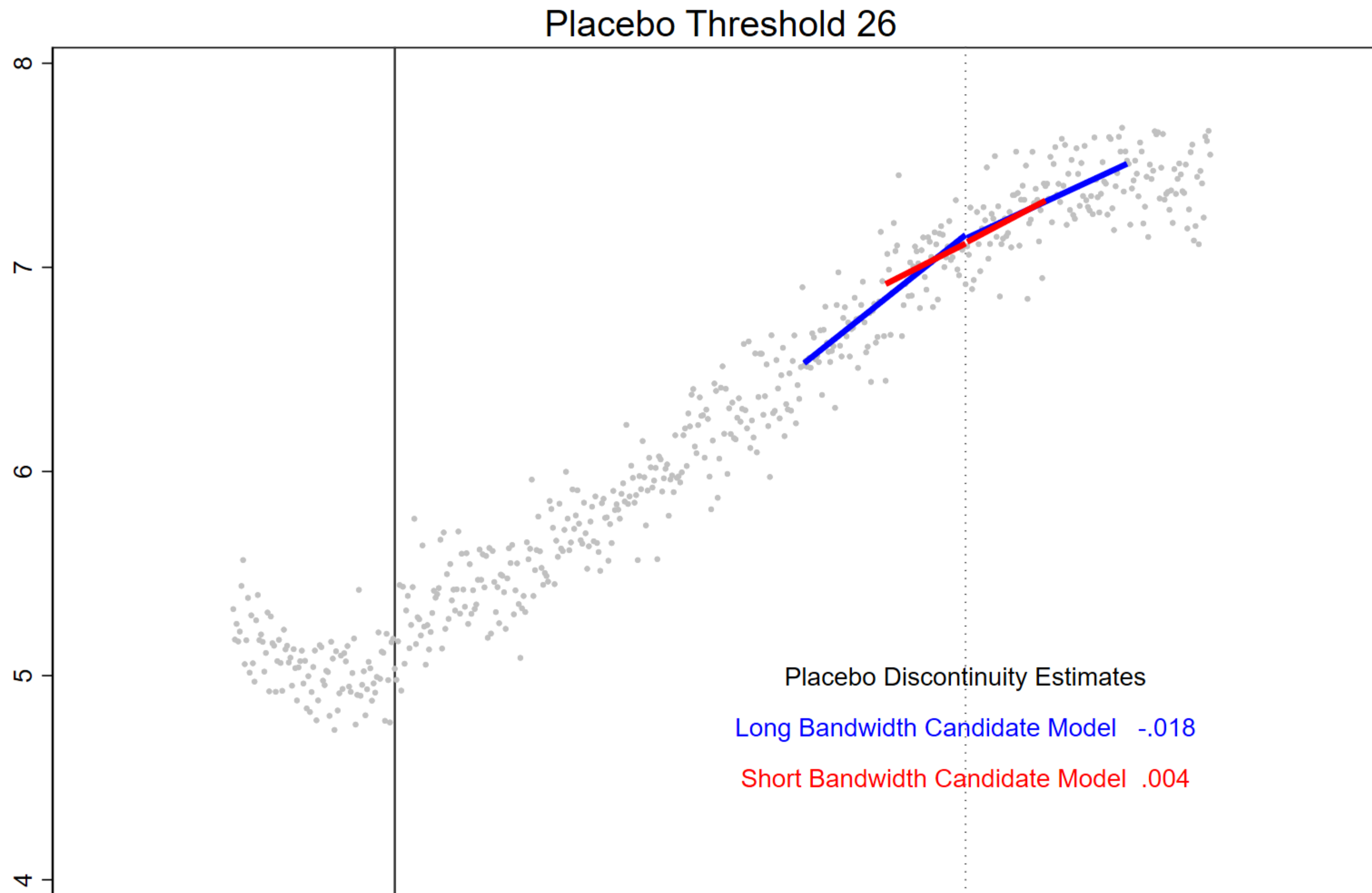


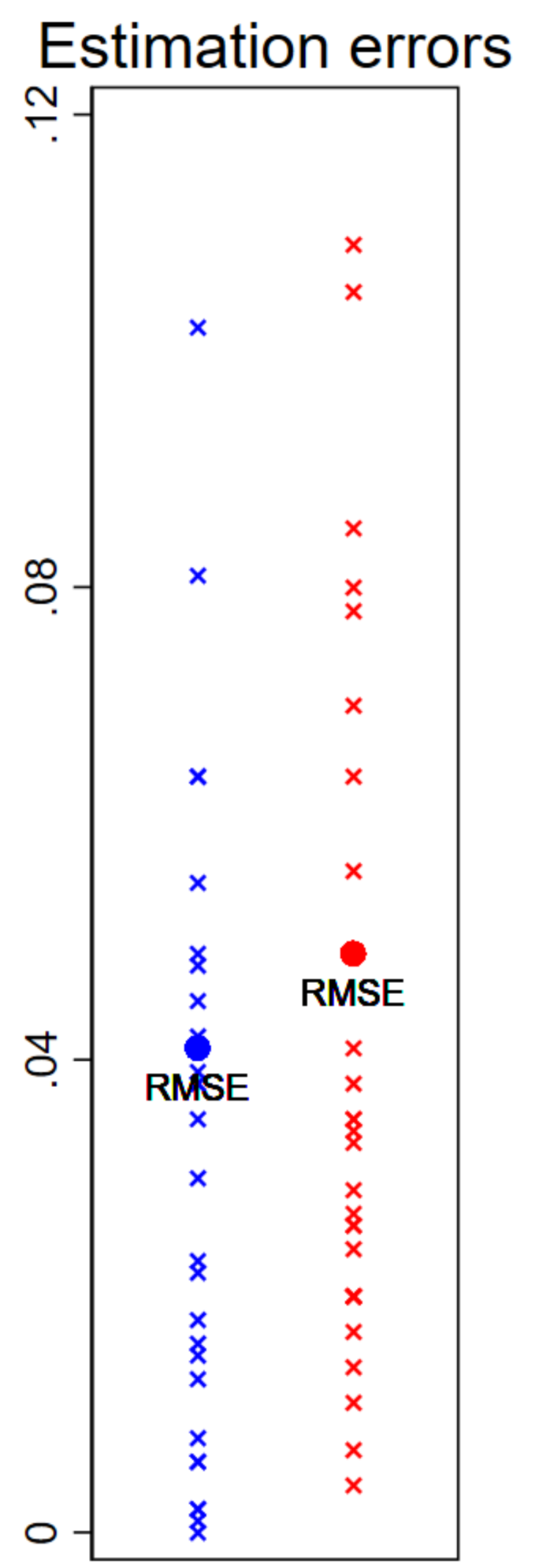
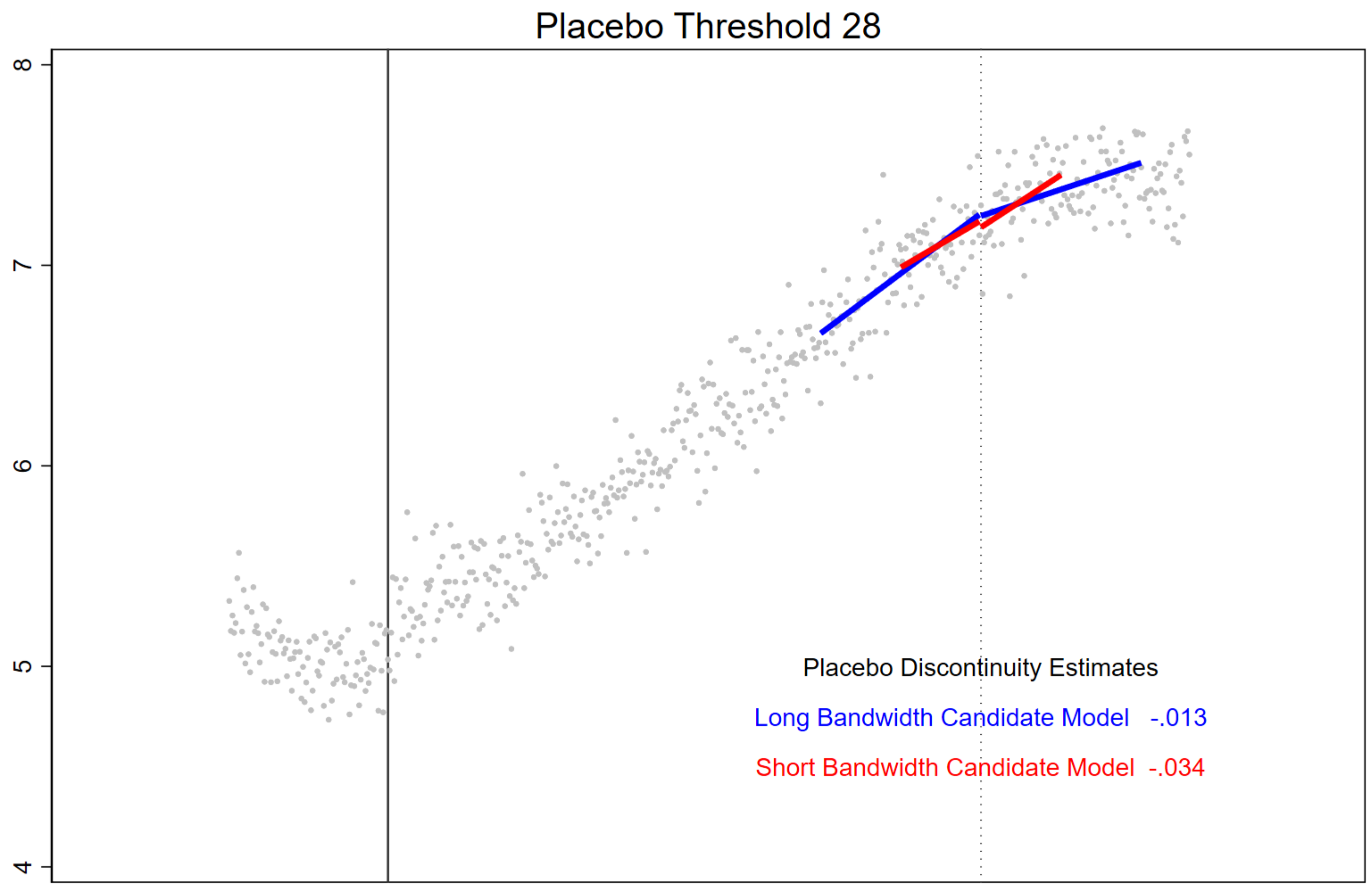


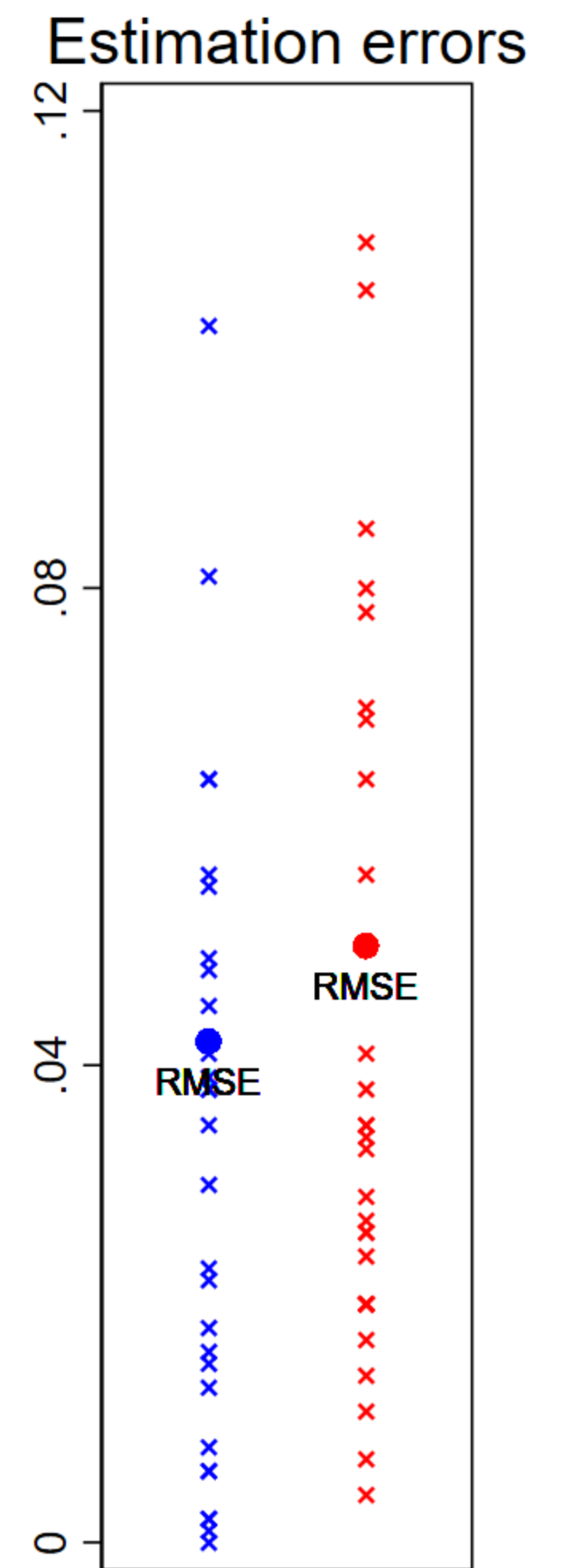
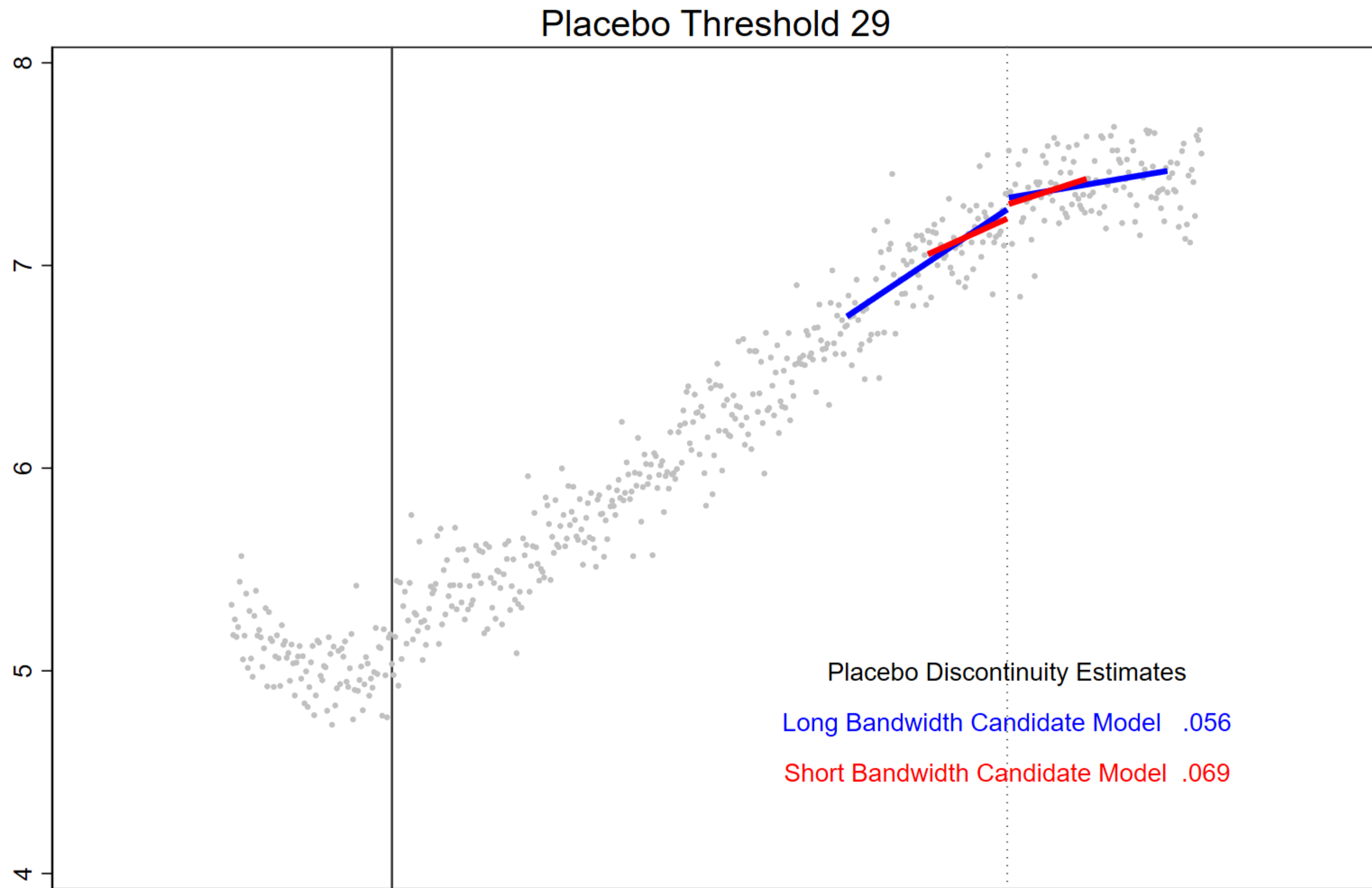


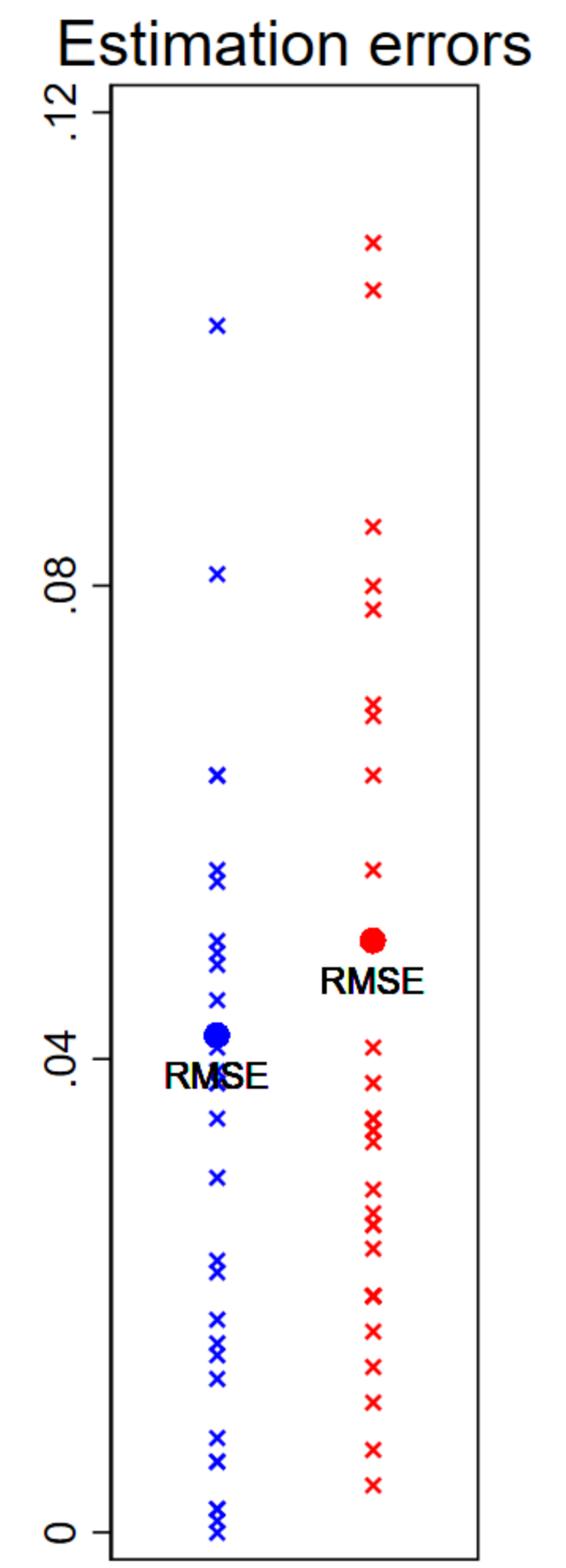
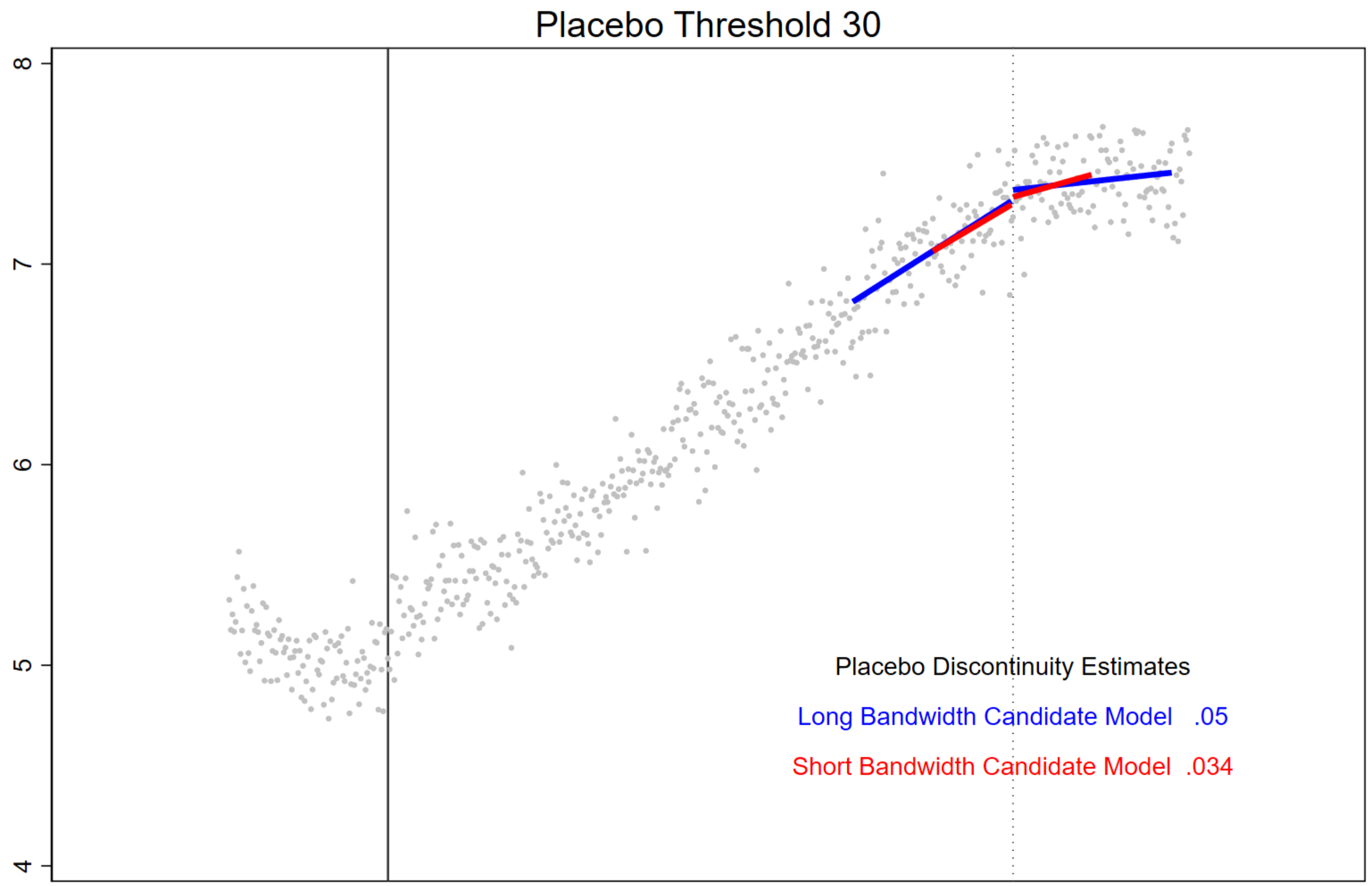


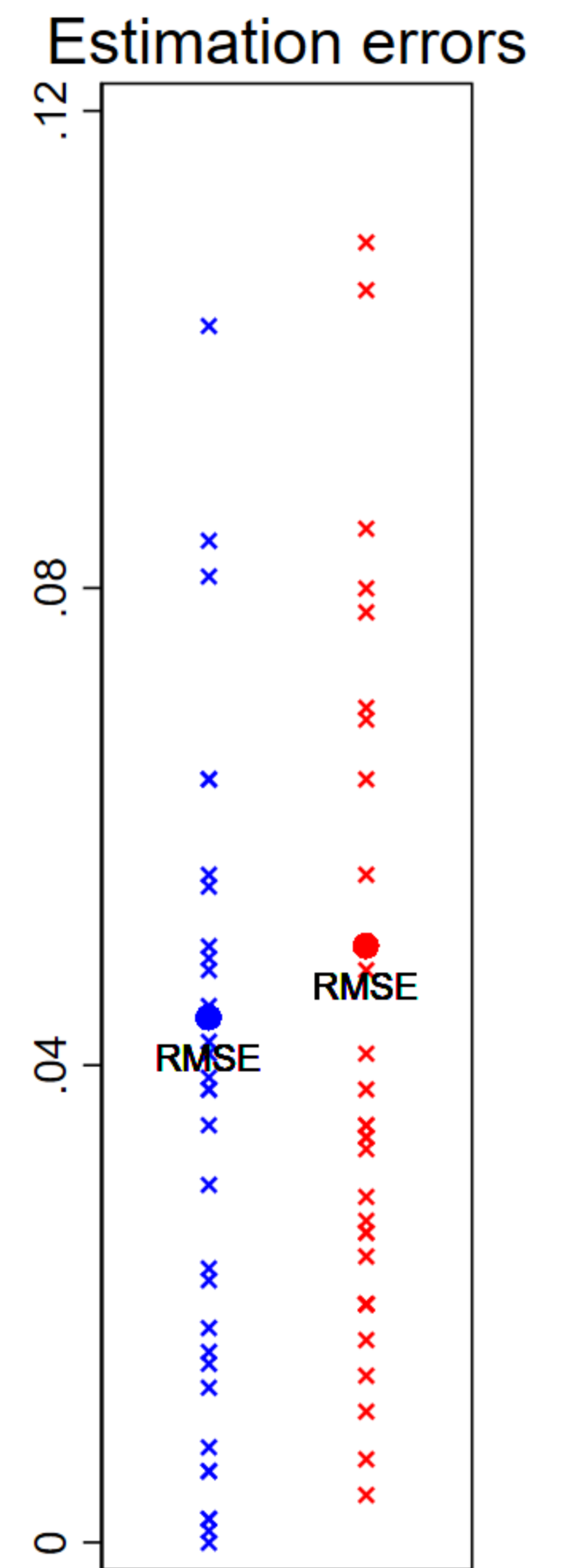
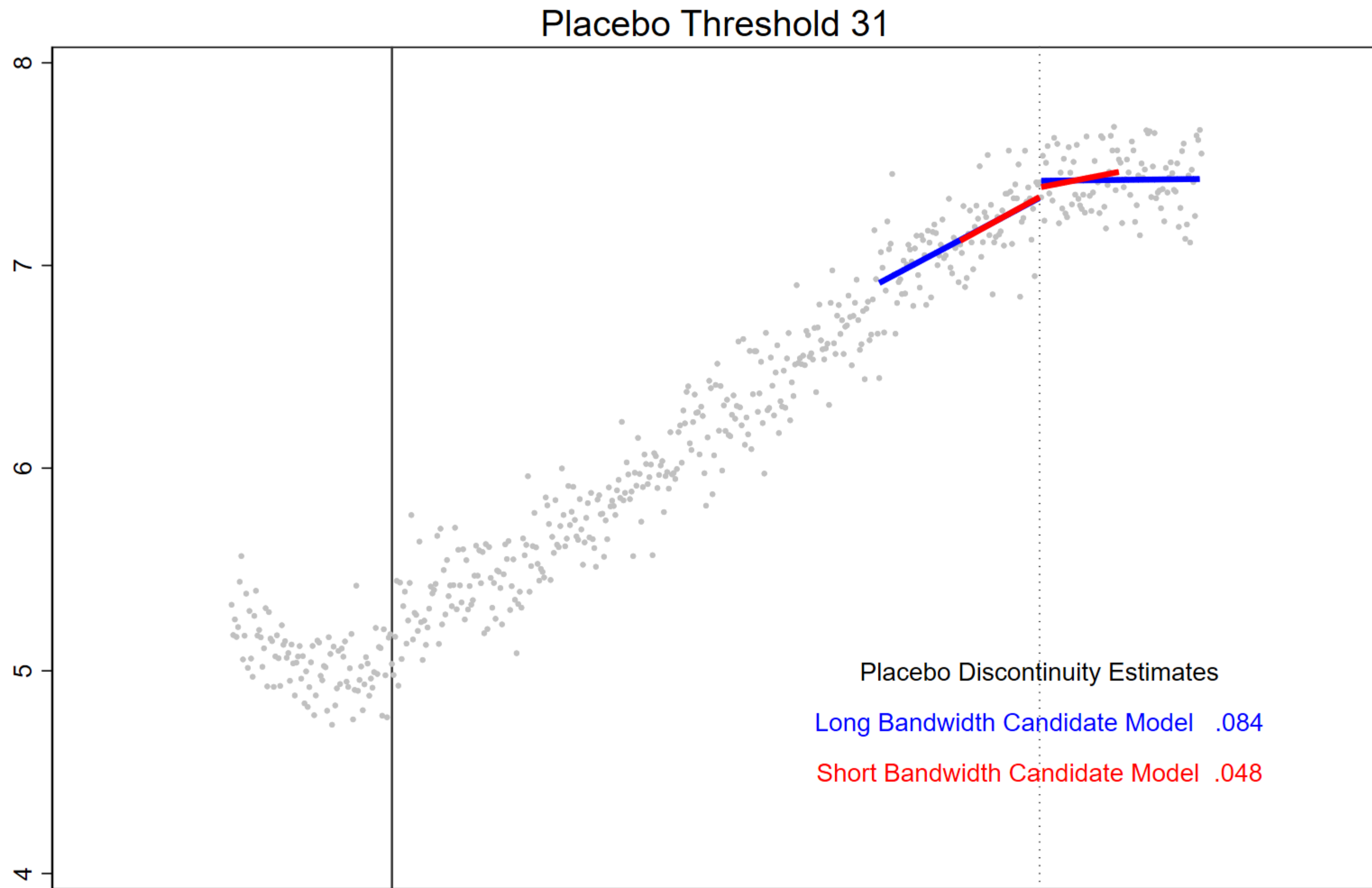




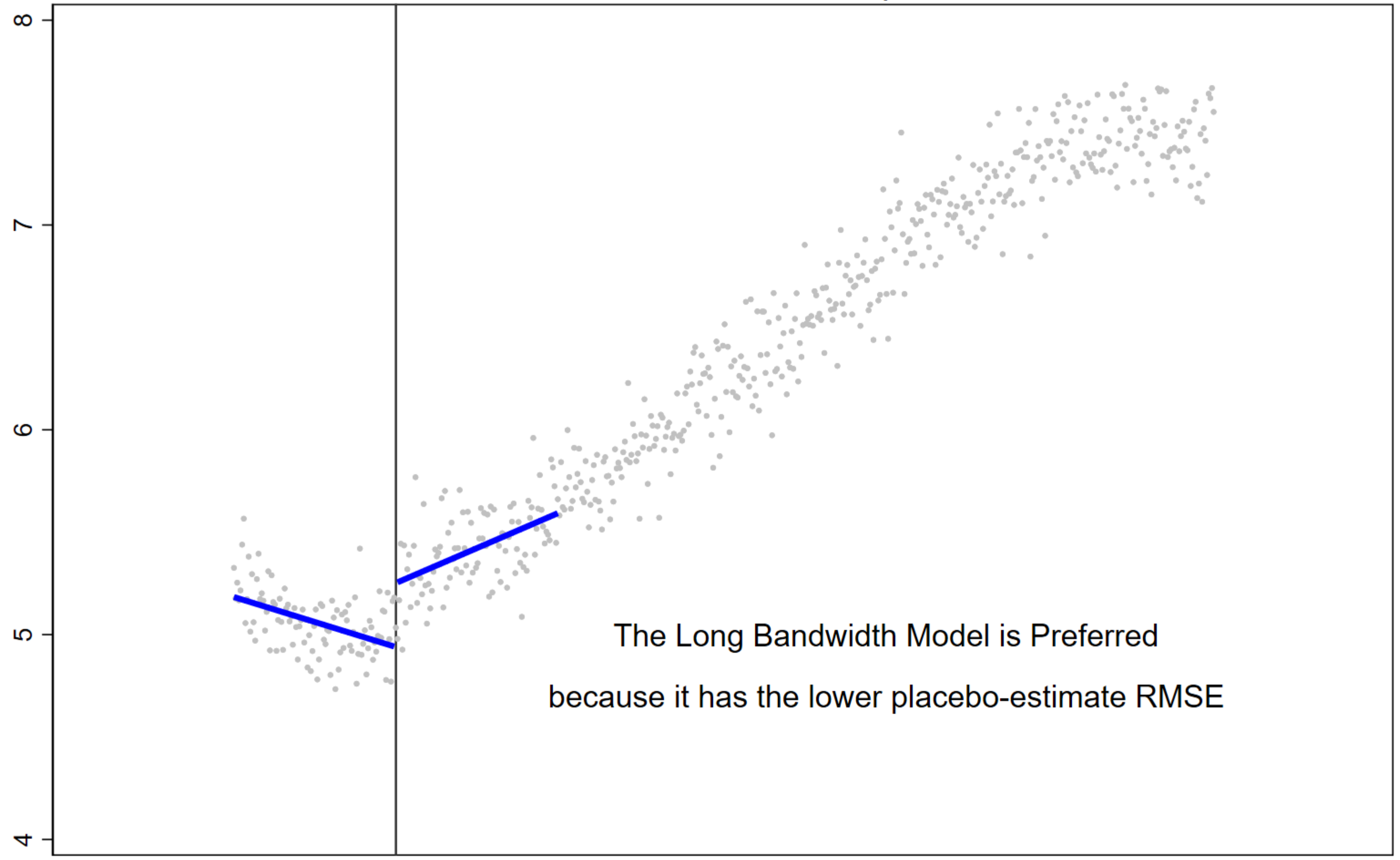




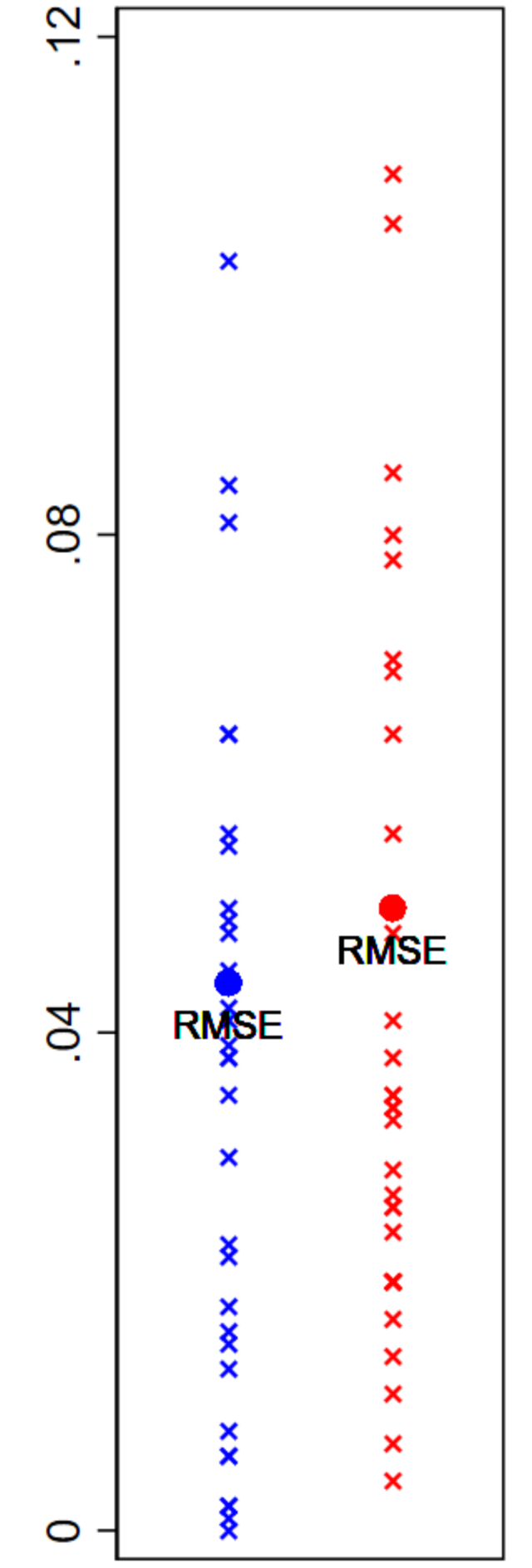




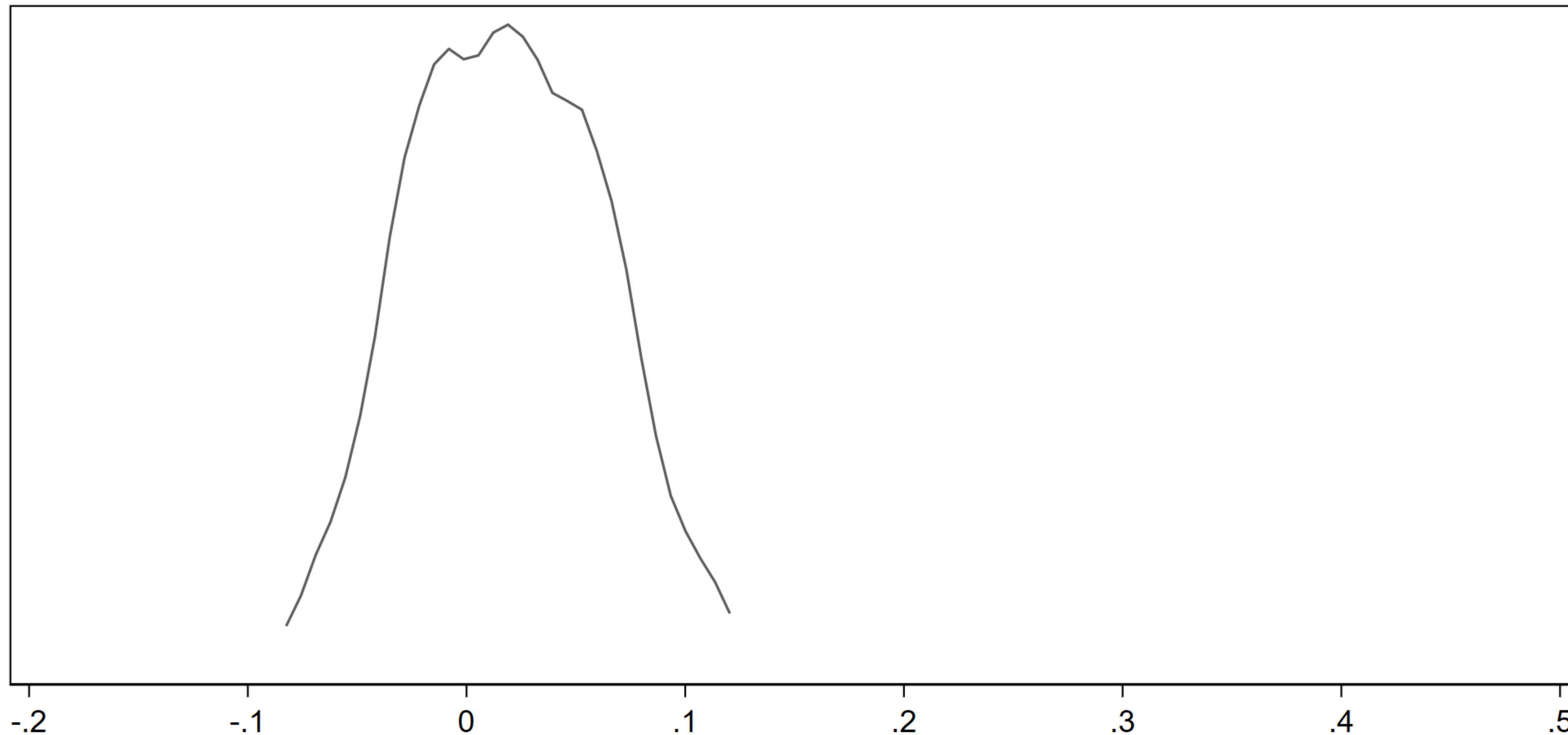
Results after 31 Placebo Repetitions



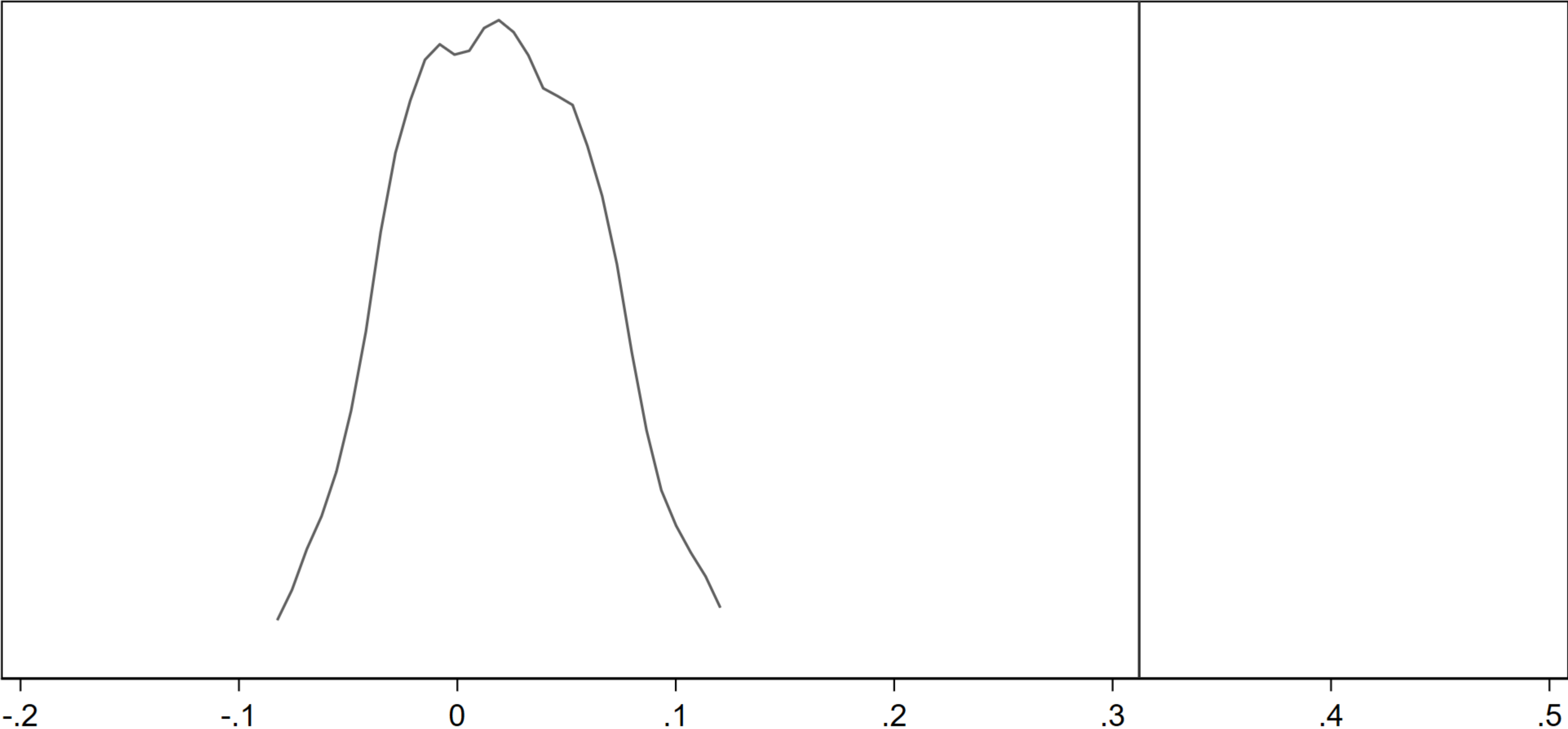
Estimation errors



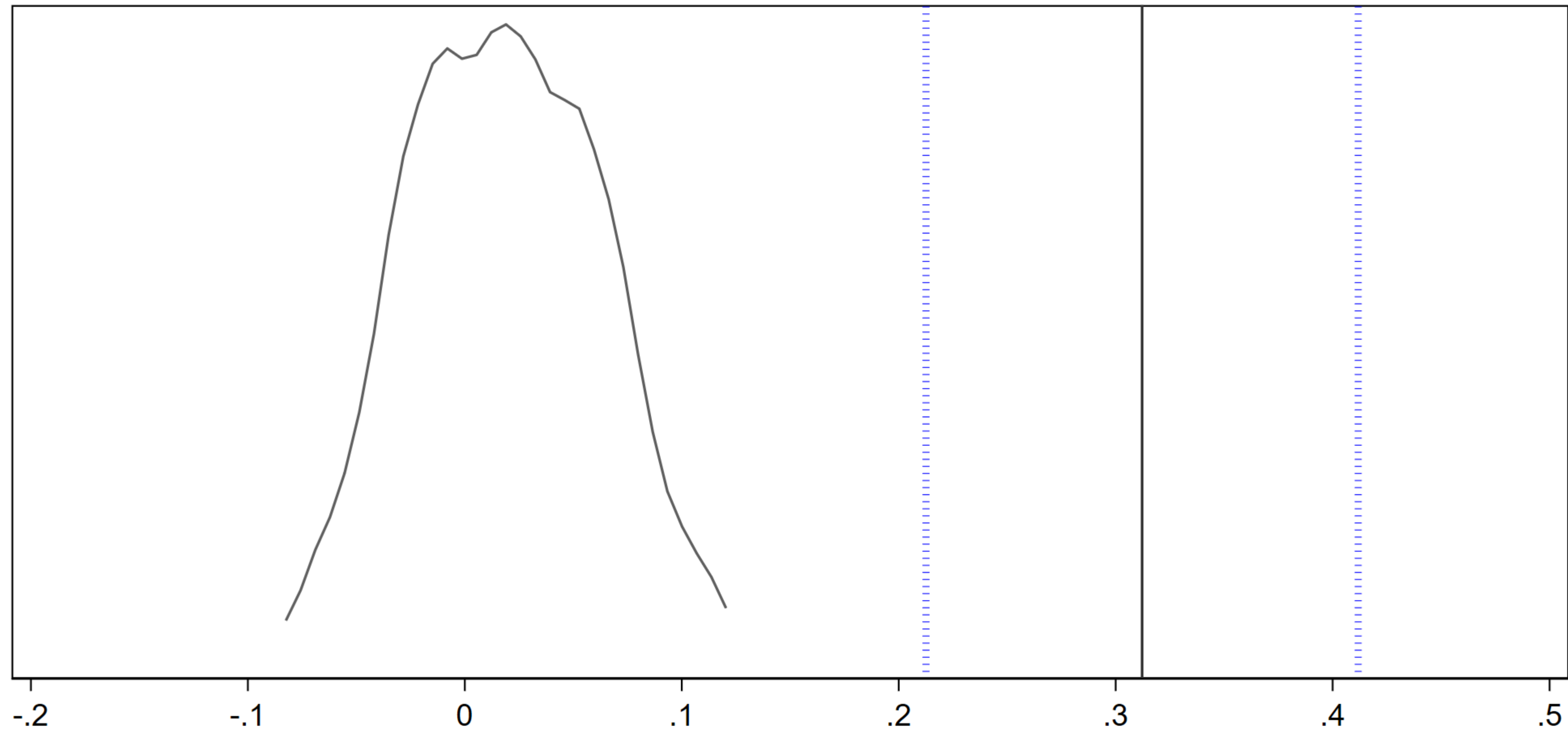
Distribution of Placebo estimates from Preferred Model



The discontinuity point-estimate alongside the placebos



A Randomisation Inference 95% Confidence Interval



We assume the placebos are drawn from a normal distribution and adjust for serial correlation

Outline

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- **Motivating application – 2 policy changes affecting Learner drivers in NSW**
- Theory – show our approach is asymptotically optimal, under restrictive conditions
- Simulations – our approach performs favourably compared to other procedures using stylised and realistic DGPs

Motivating Example – 2 Policy Changes affecting **Learner Drivers** in NSW



These policy changes were intended to reduce (subsequent) crashes



Motivating Example (cont.) These policy changes increased the number of **Mandatory Supervised Driving Hours**, from zero to 50, and then to 120 hours

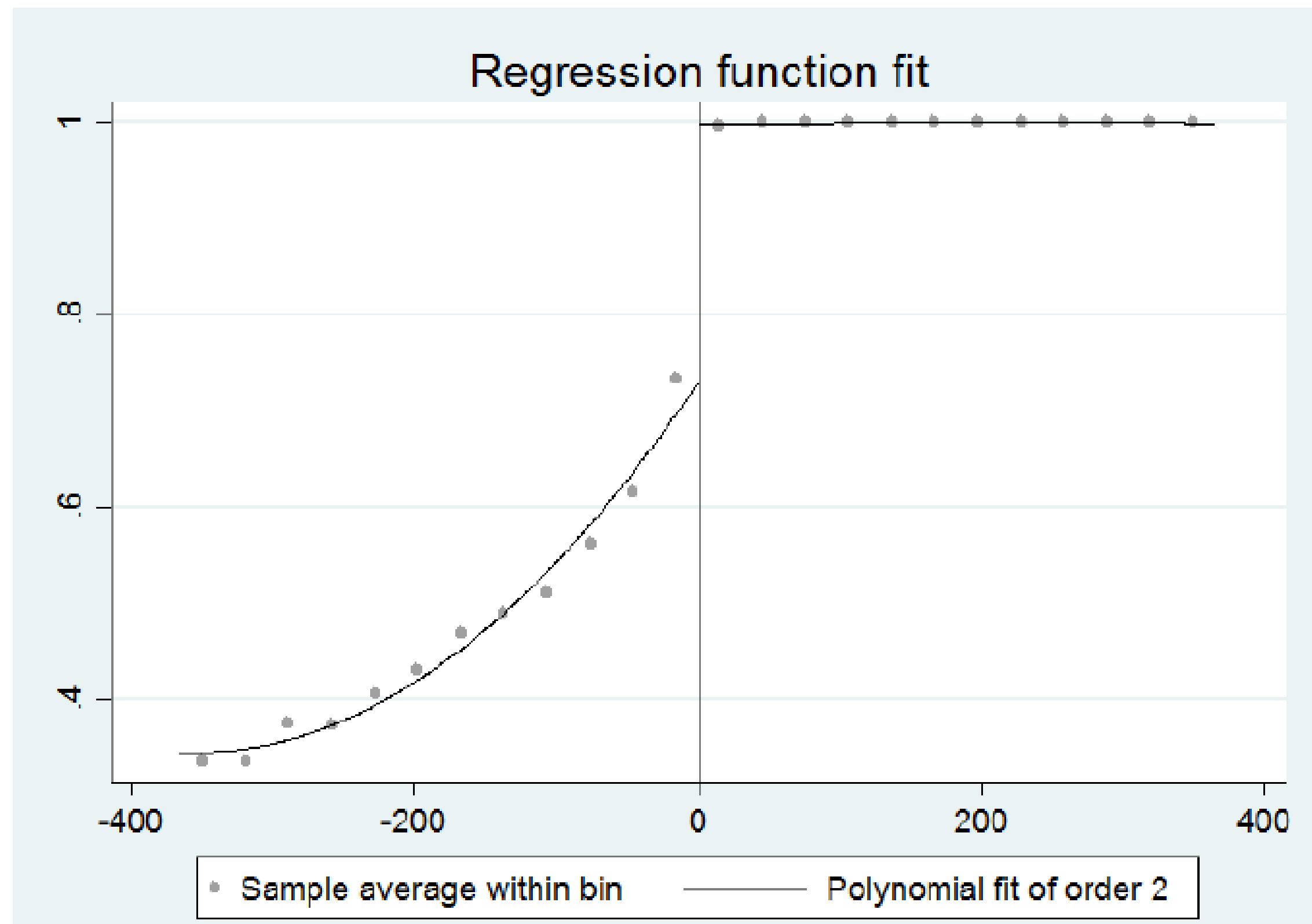


Motivating Example (cont.)

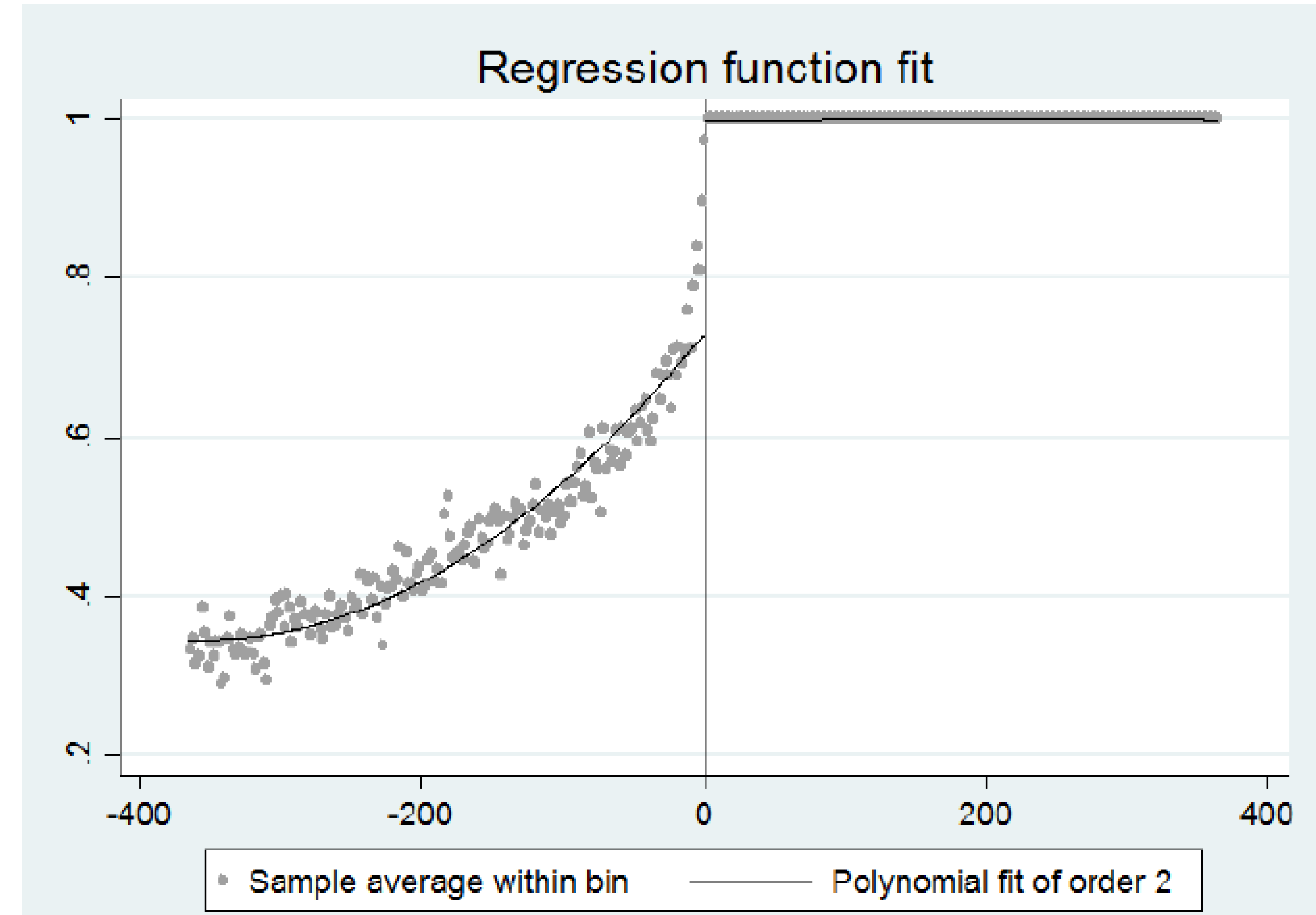
- **Mandatory Supervised Driving Hours (MSDH) for Learner Drivers in NSW**
- On 1 July 2000, MSDH increased from 0 to 50 hours.
- Because of the age requirement for a learner's permit (16 years), people born just before 1 July 1984 could avoid the policy. Those born after 1 July 1984 could not.
- A similar increase occurred on 1 July 2007 from 50 to 120 hours.
- This created discontinuities(?) in the probability of treatment by date of birth (DOB).
- Did these policy changes affect crash rates??

First-stage relationship between DOB (centred around 1 July 1984) and 'Treatment' (i.e. 50+ hours versus 0+)

P(Mandated 50 hours | Ps prior to age 30) - 1 month bins

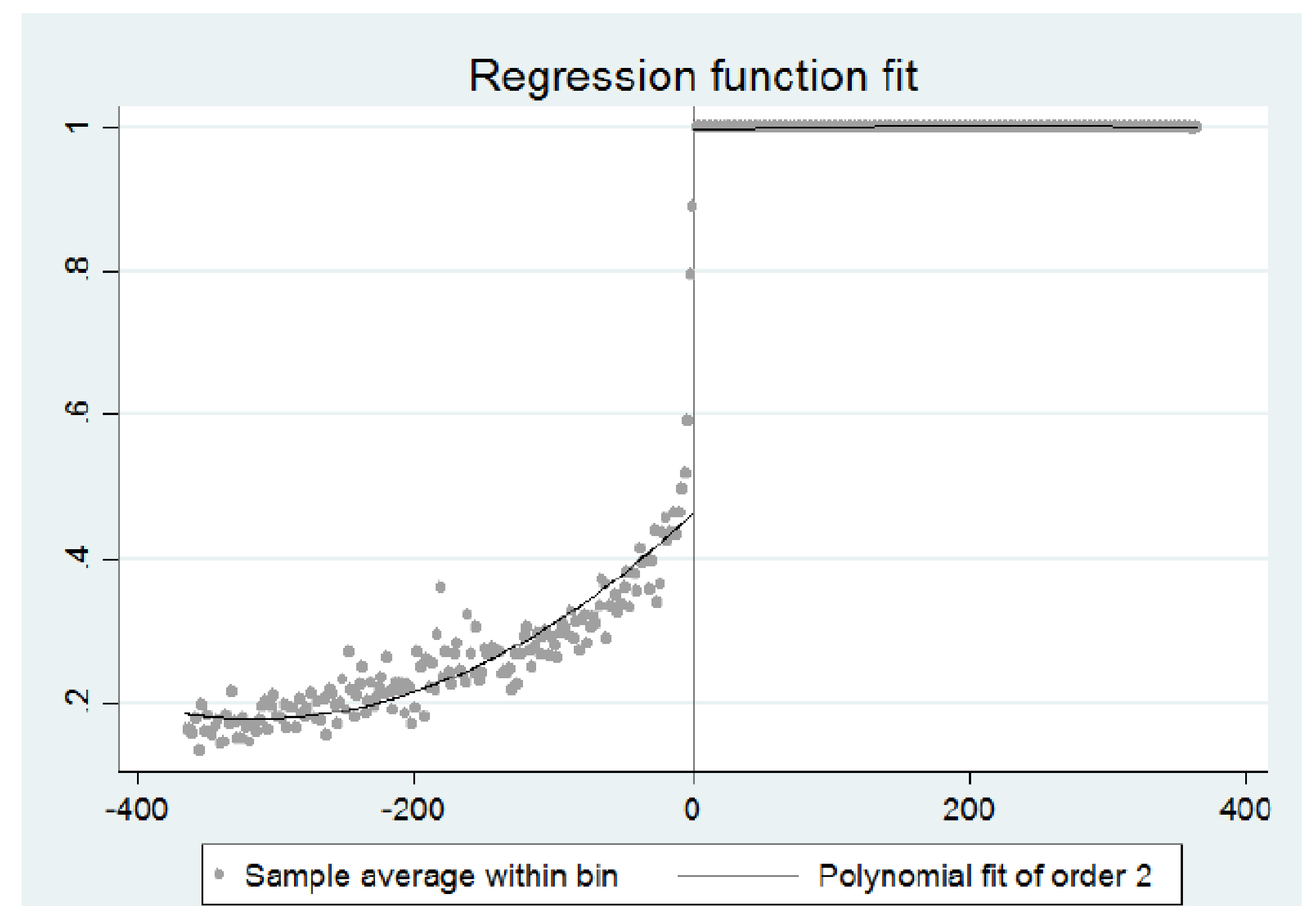
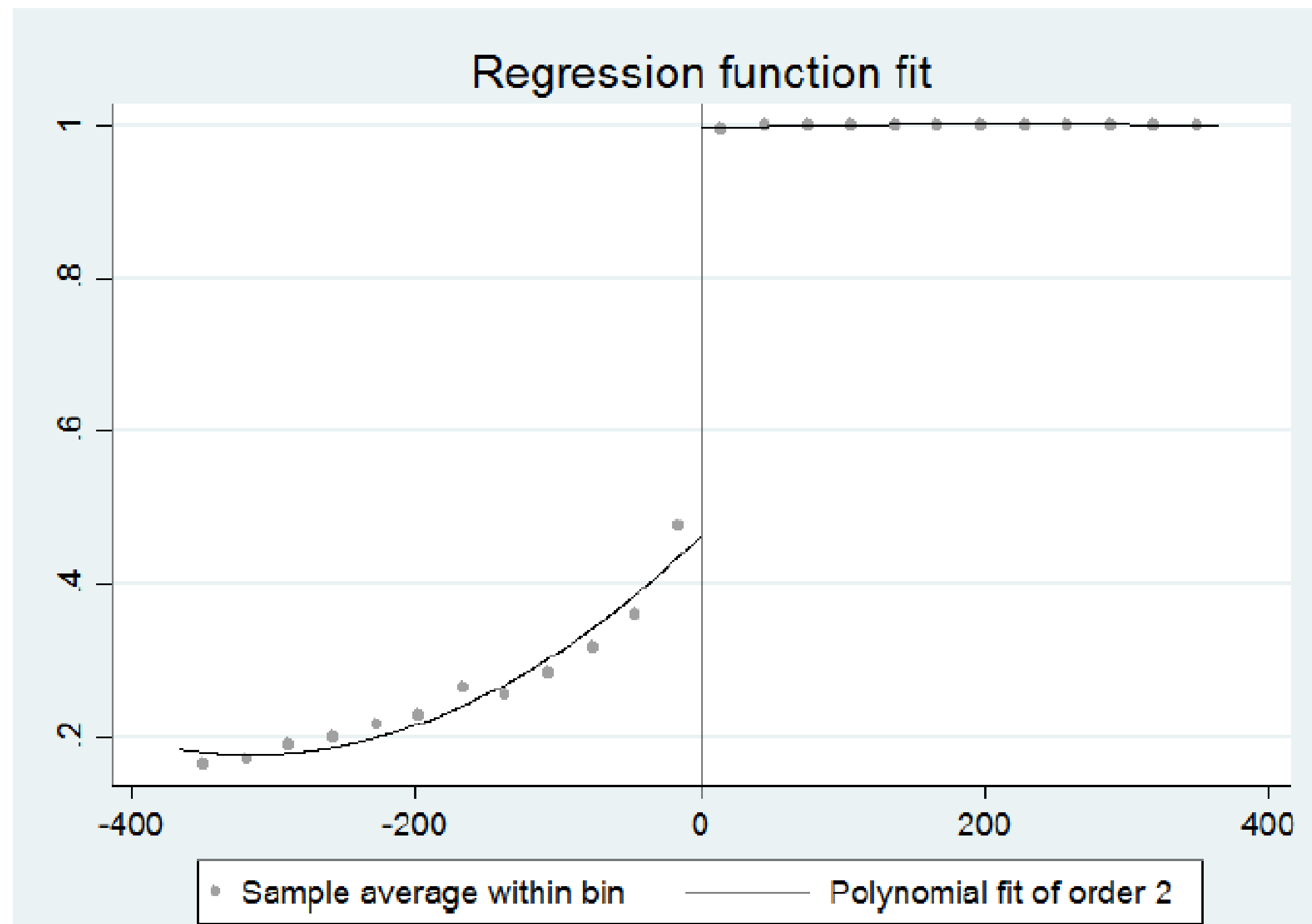


P(Mandated 50 hours | Ps prior to age 30) - 2 day bins



First-stage relationship between DOB (centred around 1 July 1991) and 'Treatment' (i.e. 120+ hours versus 50+)

P(Mandated 120 hours | Ps prior to age 23) - monthly bins



Outline

- Key Issues and our Contribution
- Motivating Application – estimating effects of 2 policy changes affecting Learner Drivers in NSW
- **Theory – show our approach is asymptotically optimal, under restrictive conditions**
- Simulations – our approach performs favourably compared to other procedures using stylised and realistic DGPs

Proof of Asymptotic Optimality (Overview)

- Our approach is ‘asymptotically optimal’ if the best treatment effect estimator also has the lowest mean squared placebo estimates, when the placebo zone is large.
- This is the case if the MSE at each placebo threshold equals the MSE at the treatment threshold.
- We show this, assuming the global DGP’s CEF has a zero fourth-derivative, under homoskedasticity, and uniformly distributed x .
- We focus on sharp RDD, with local linear estimators. The results also translate to higher-order polynomials.

Proof of Asymptotic Optimality (Outline)

1. We show that bias of linear RDD estimators is proportional to the 3rd derivative of the DGP's CEF.
2. This bias is constant across the support of the running variable, assuming the 4th derivative is zero. Therefore the bias of placebo estimators equals the bias of the treatment effect estimator
3. The variance of placebo estimators also equals the variance of the treatment effect estimator, assuming homoskedasticity and a uniform distribution
4. For each estimator, the observed mean of squared estimates across the placebo zone approaches the (unobservable) MSE of the treatment effect estimator as the placebo zone becomes large
5. Therefore the approach is 'asymptotically optimal'

Outline

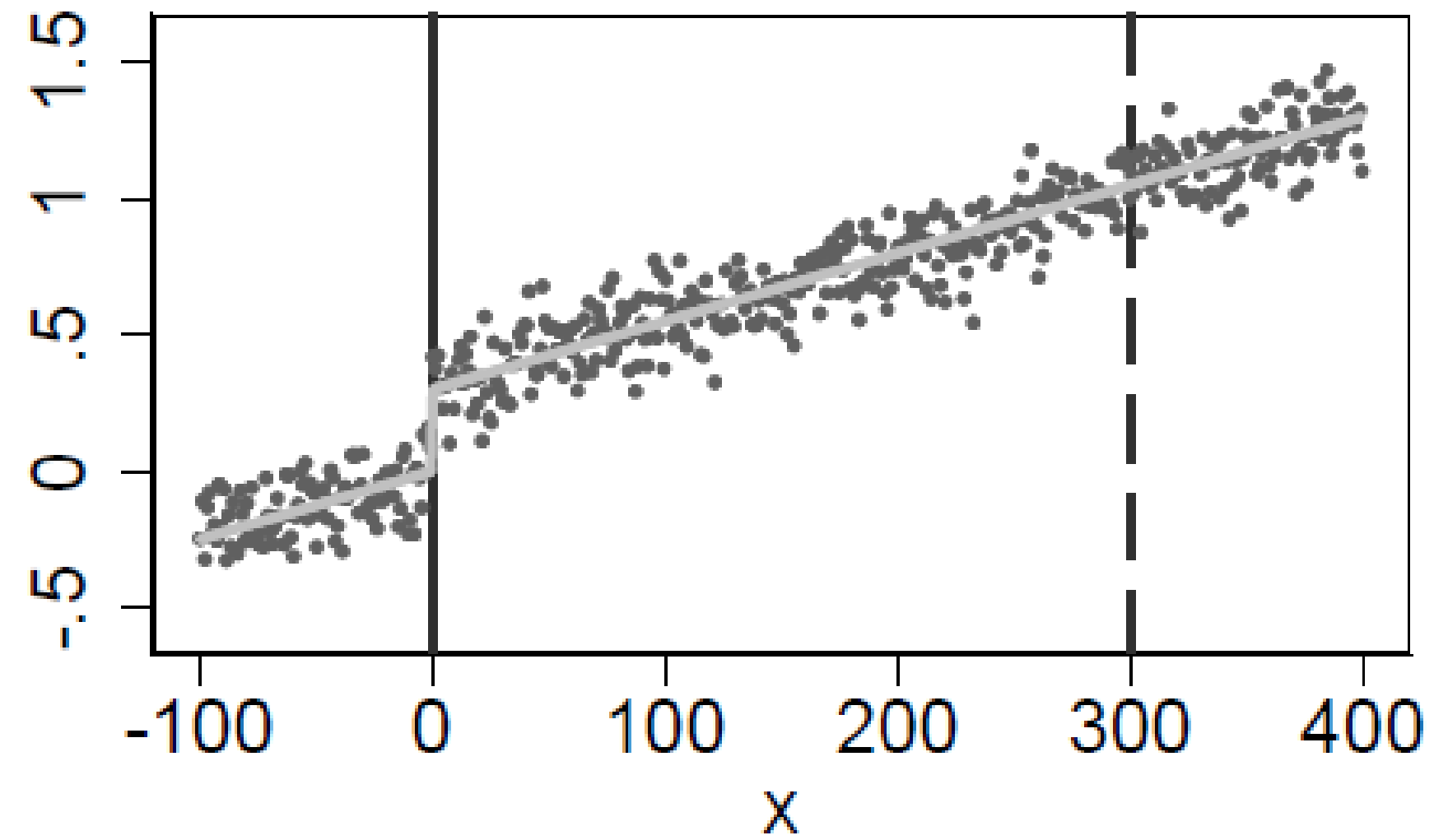
- Key Issues and our Contribution
- Motivating Application – 2 policy changes affecting Learner Drivers in NSW
- Theory – show our approach is asymptotically optimal, under restrictive conditions
- **Simulations – our approach performs favourably compared to other procedures using stylised and realistic DGPs**

Monte Carlo Simulations

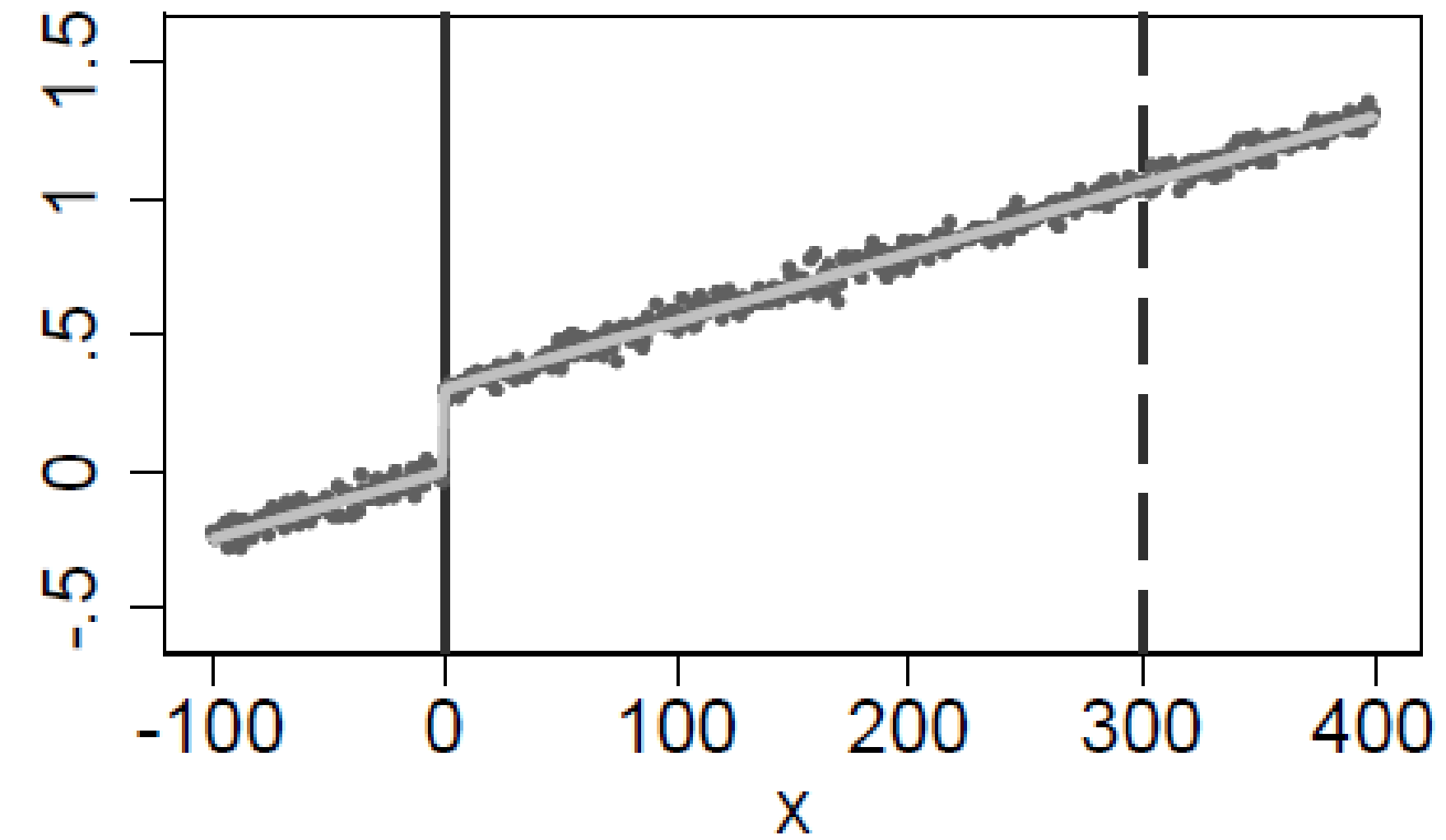
- Great in theory, but does our approach work well with:
 - Finite placebo zones?
 - DGPs with non-zero fourth derivatives?
 - DGPs with non-constant density
 - Realistic DGPs, such as those in prominent well-known RDD studies?
- Our simulation approach closely follows related simulation work
- Use 1000 reps. In each rep, we
 - i) trial many candidate estimators (linear and quadratic models with a wide range of BWs) through the placebo zone and pick the best performer on RMSE
 - ii) Apply that model to estimate the actual treatment effect
- Compare RMSE (across reps) of our approach to those chosen by the CCT and IK algorithms

Stylised DGPs (first draw of 1000 reps)

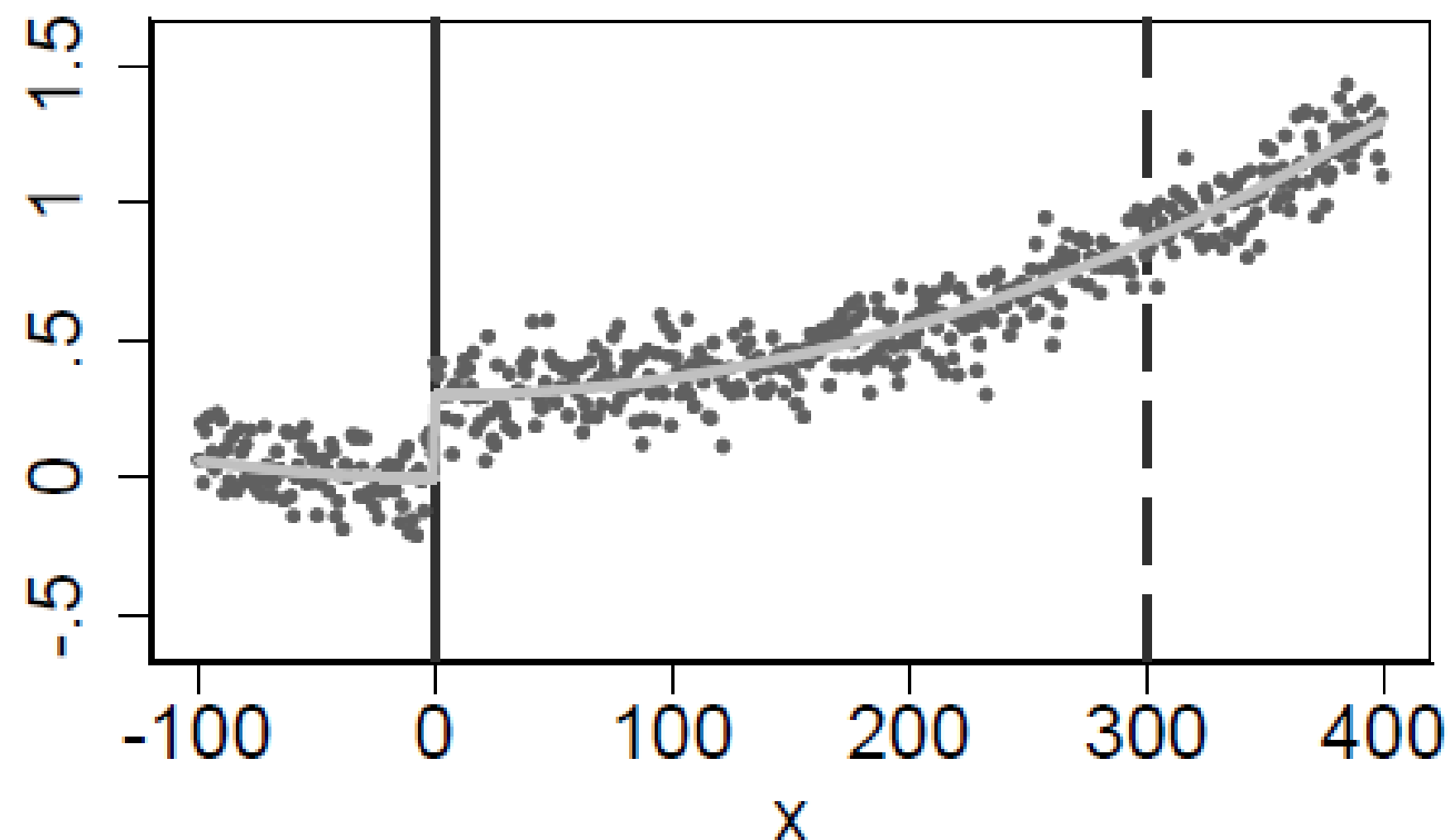
A: Linear, Large Variance



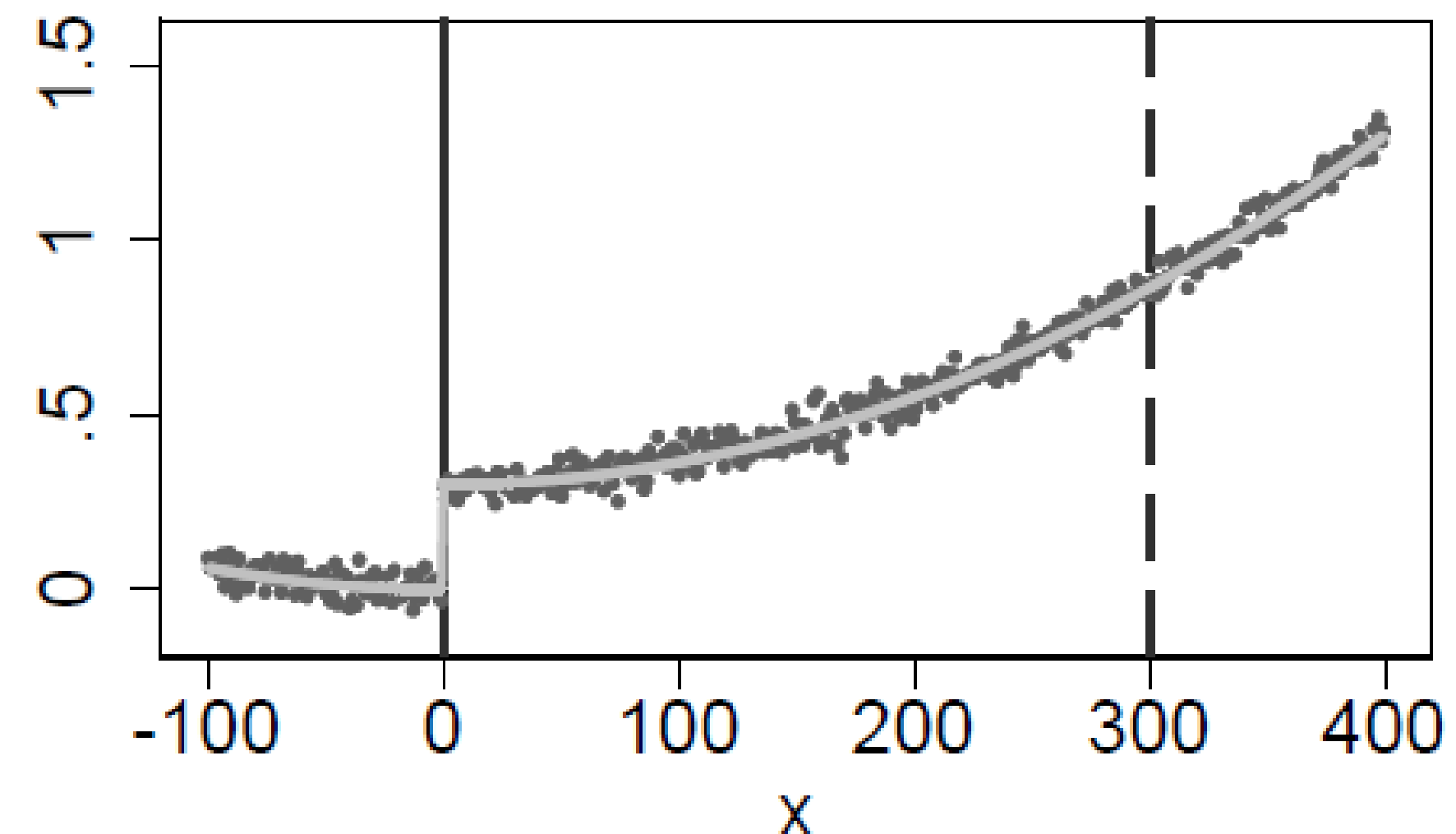
B: Linear, Small Variance



C: Quadratic, Large Variance

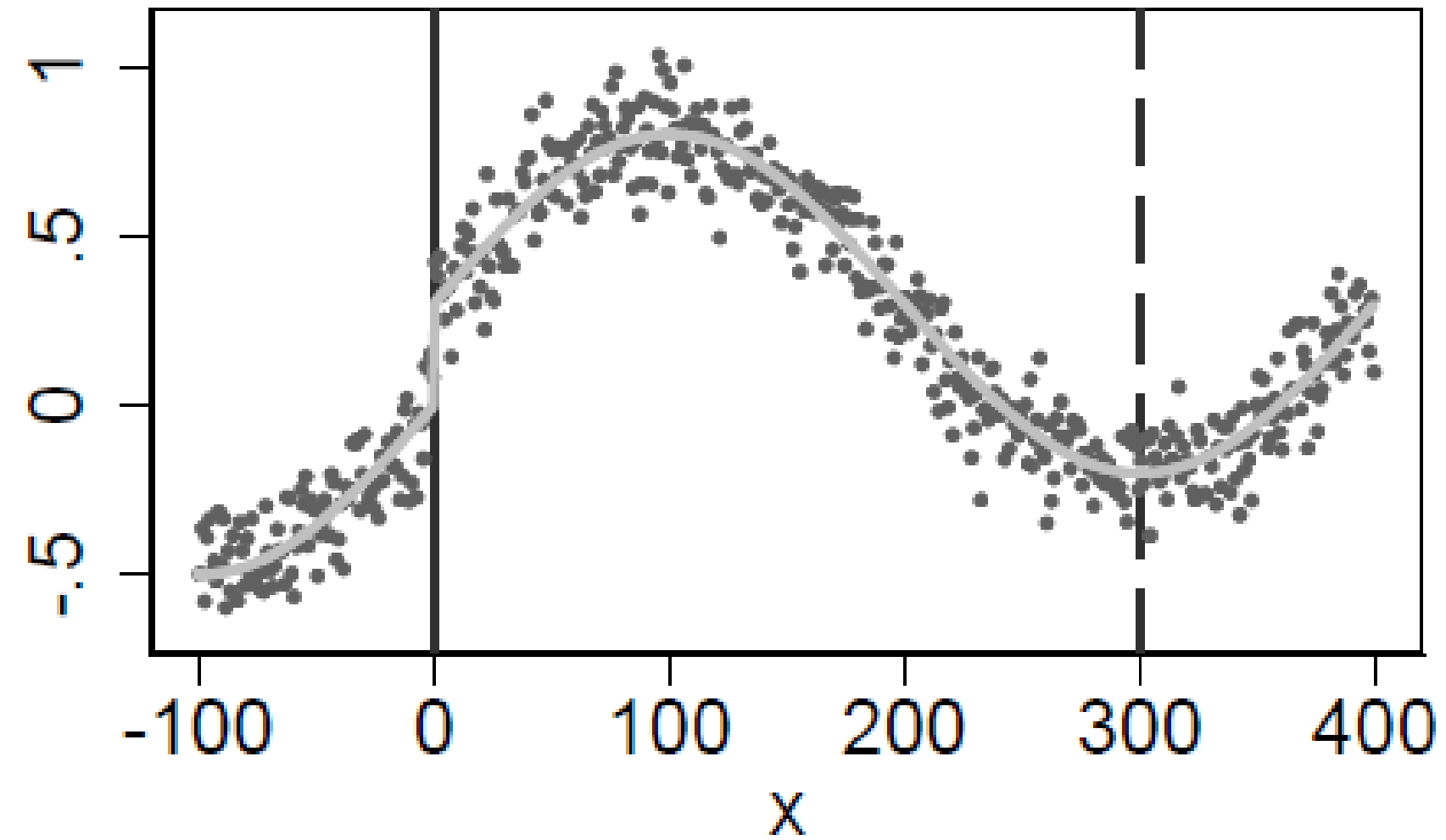


D: Quadratic, Small Variance

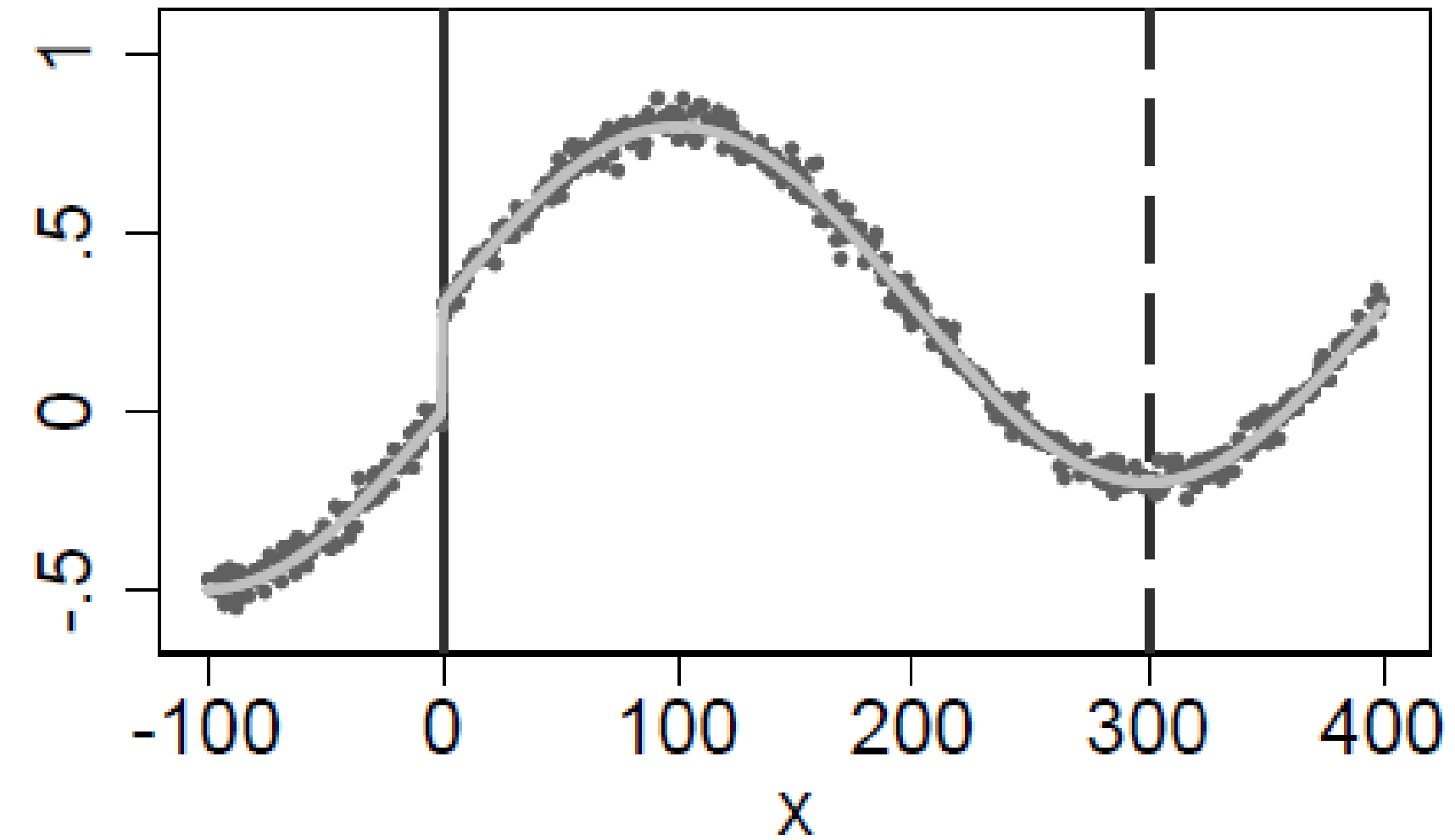


Stylised DGPs (first draw of 1000 reps)

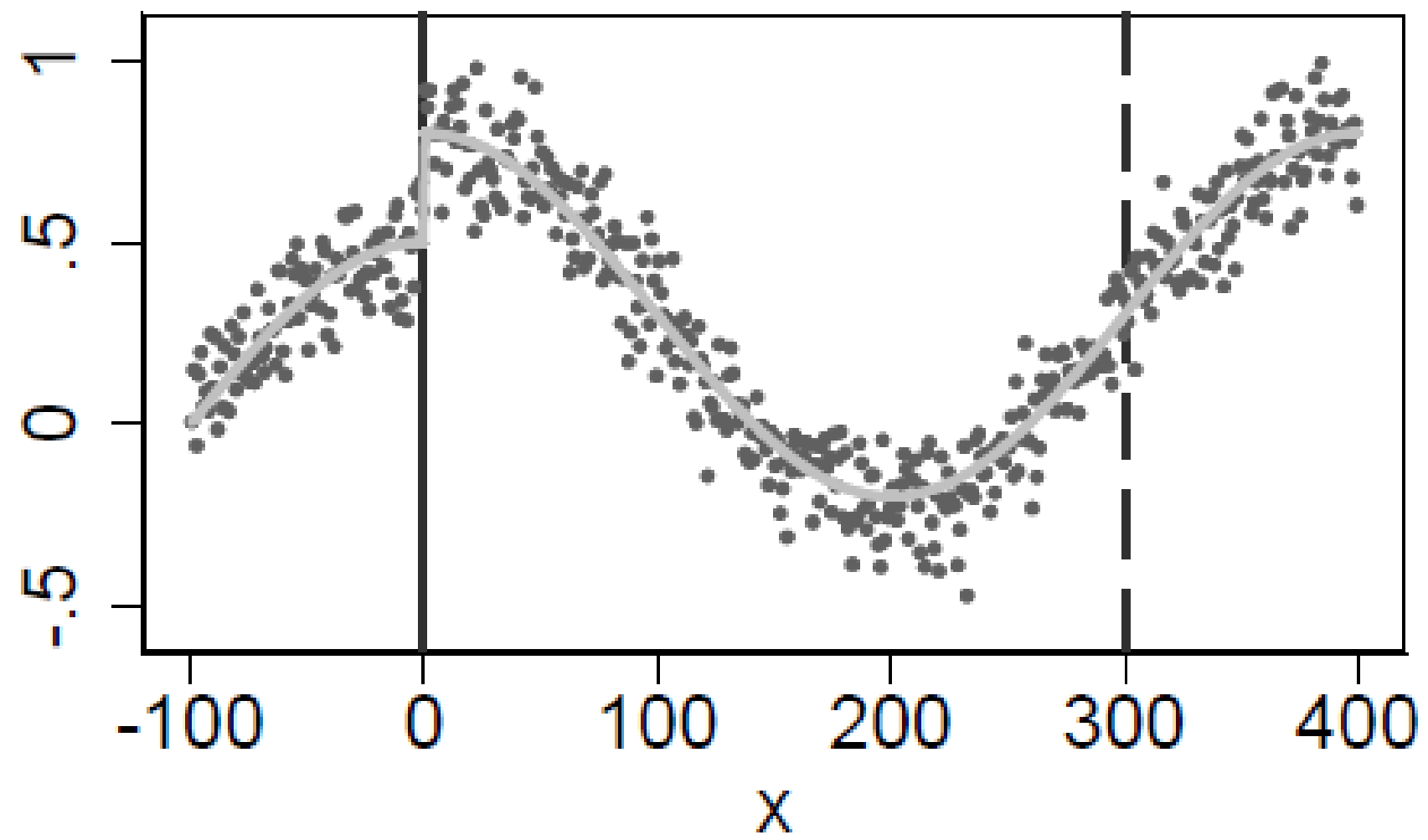
E: Sine, Large Variance



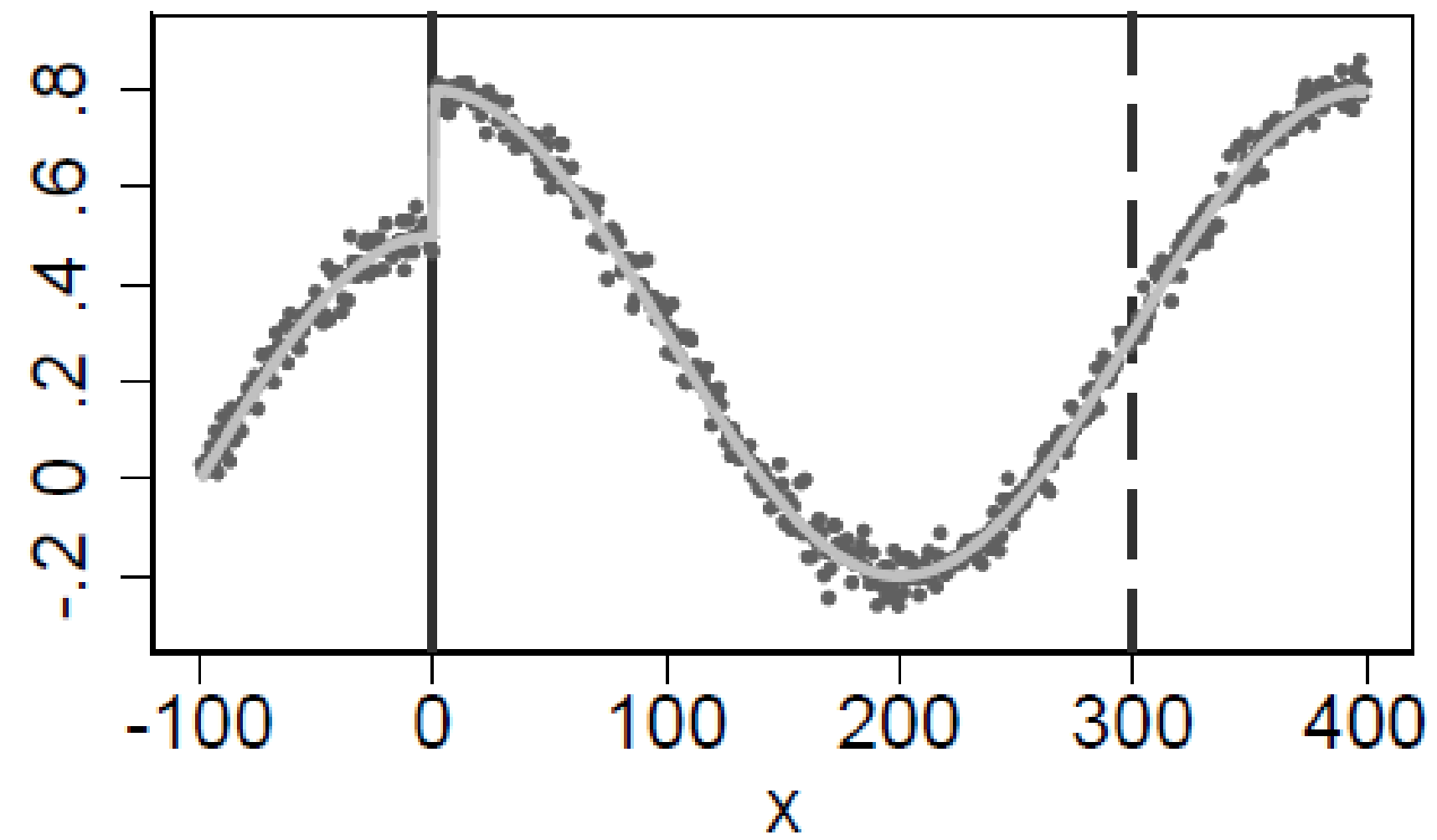
F: Sine, Small Variance



G: Cosine, Large Variance



H: Cosine, Small Variance



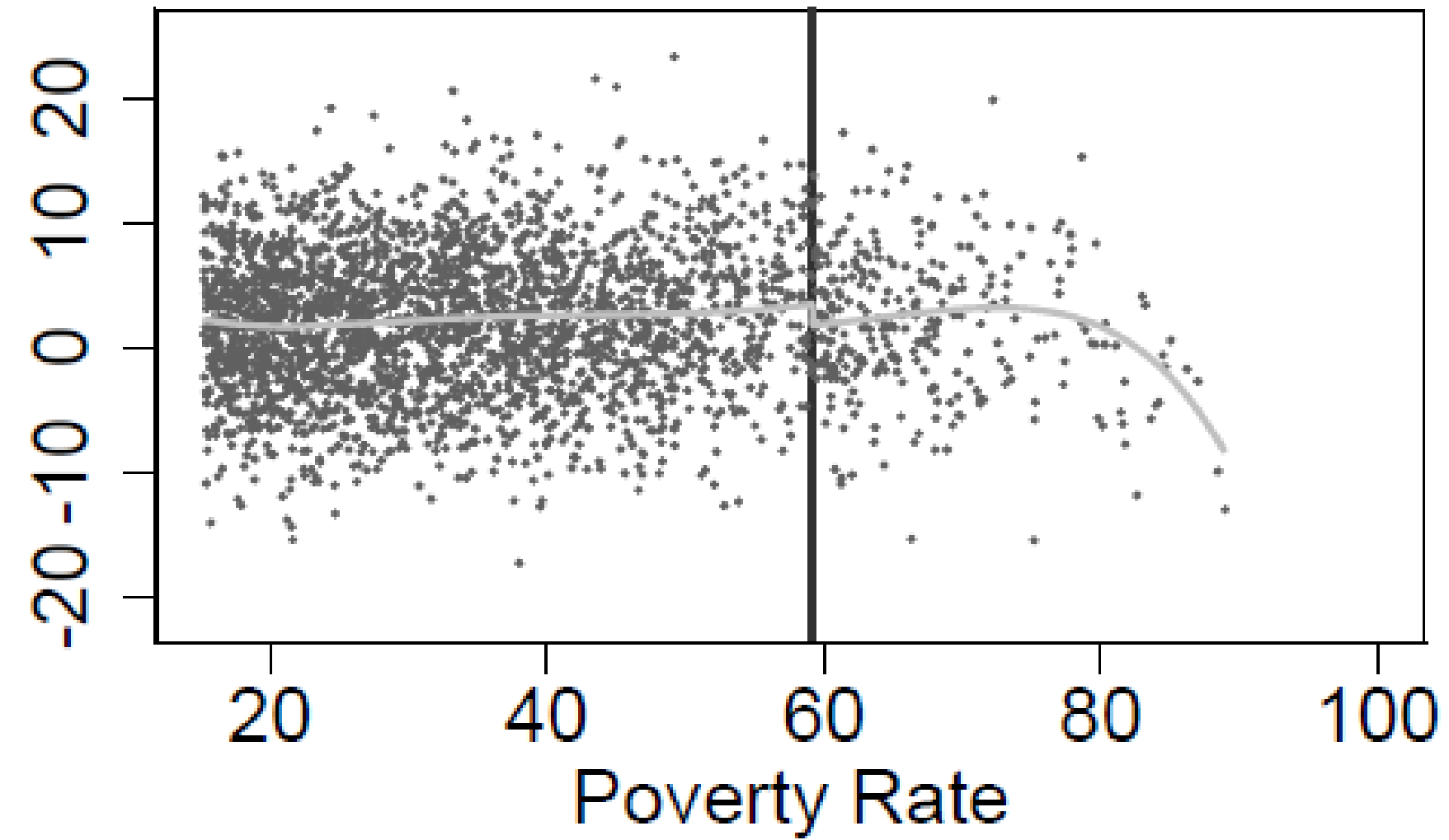
Monte Carlo Simulations (cont.)

- **Realistic DGPs**

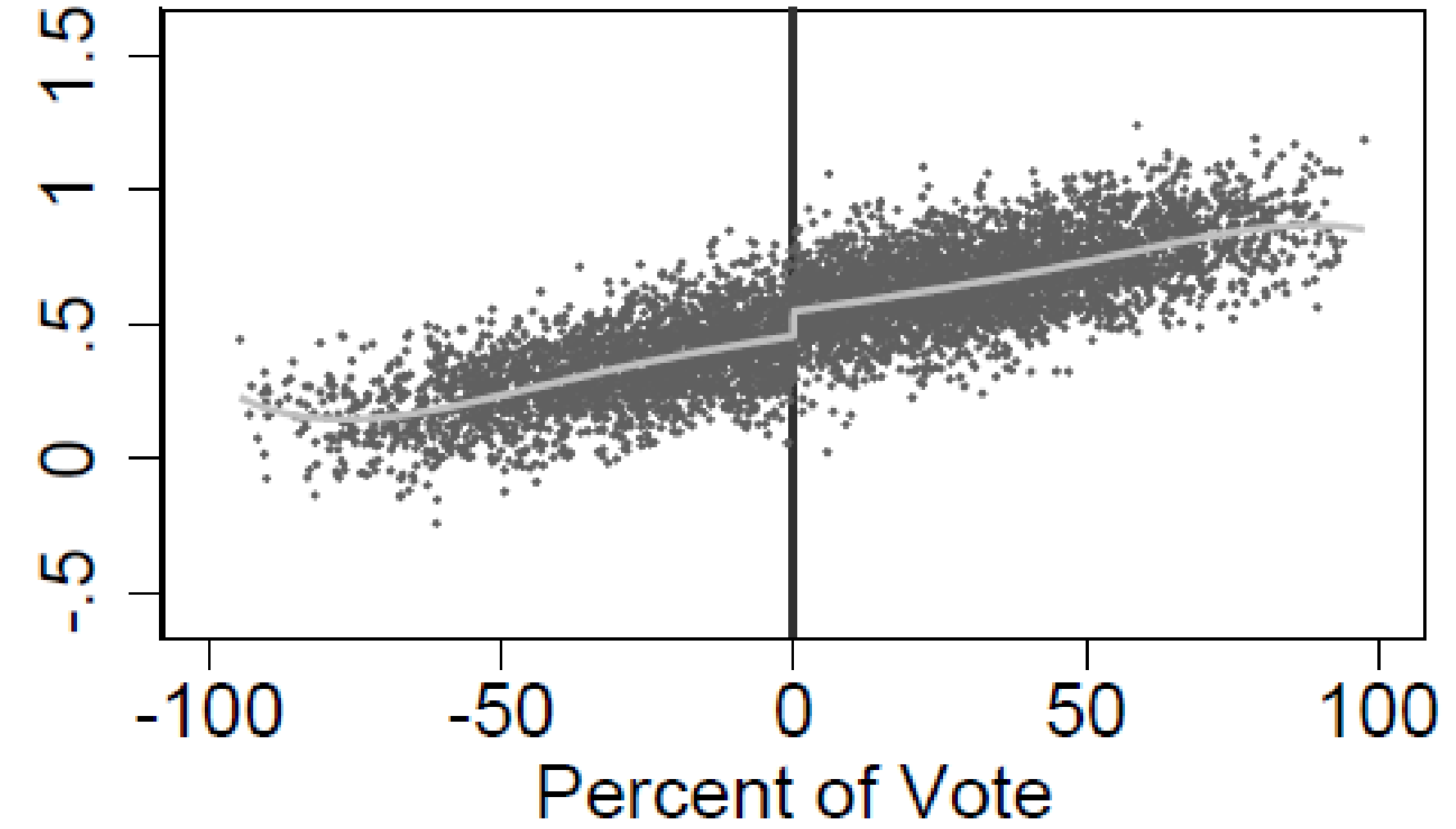
- Mimic well-known applications: Head Start (Ludwig & Miller, 2007), political incumbency (Lee, 2008), and Minimum Legal Drinking Age (MLDA)
- Fit $f(x)$, 5th order polynomial to original data, allowing a discontinuity and kink
- Fit Beta-distribution to summarise distribution of running variable.
- In each iteration, sample size is set equal to the original sample.
- Randomly draw values of the running variable from the beta distribution.
- Set $y = f(x) + e$, where e is normally distributed with zero mean and variance equal to the variance of the residuals from the regression in 1st step.

Realistic DGPs (first draw of 1000 reps)

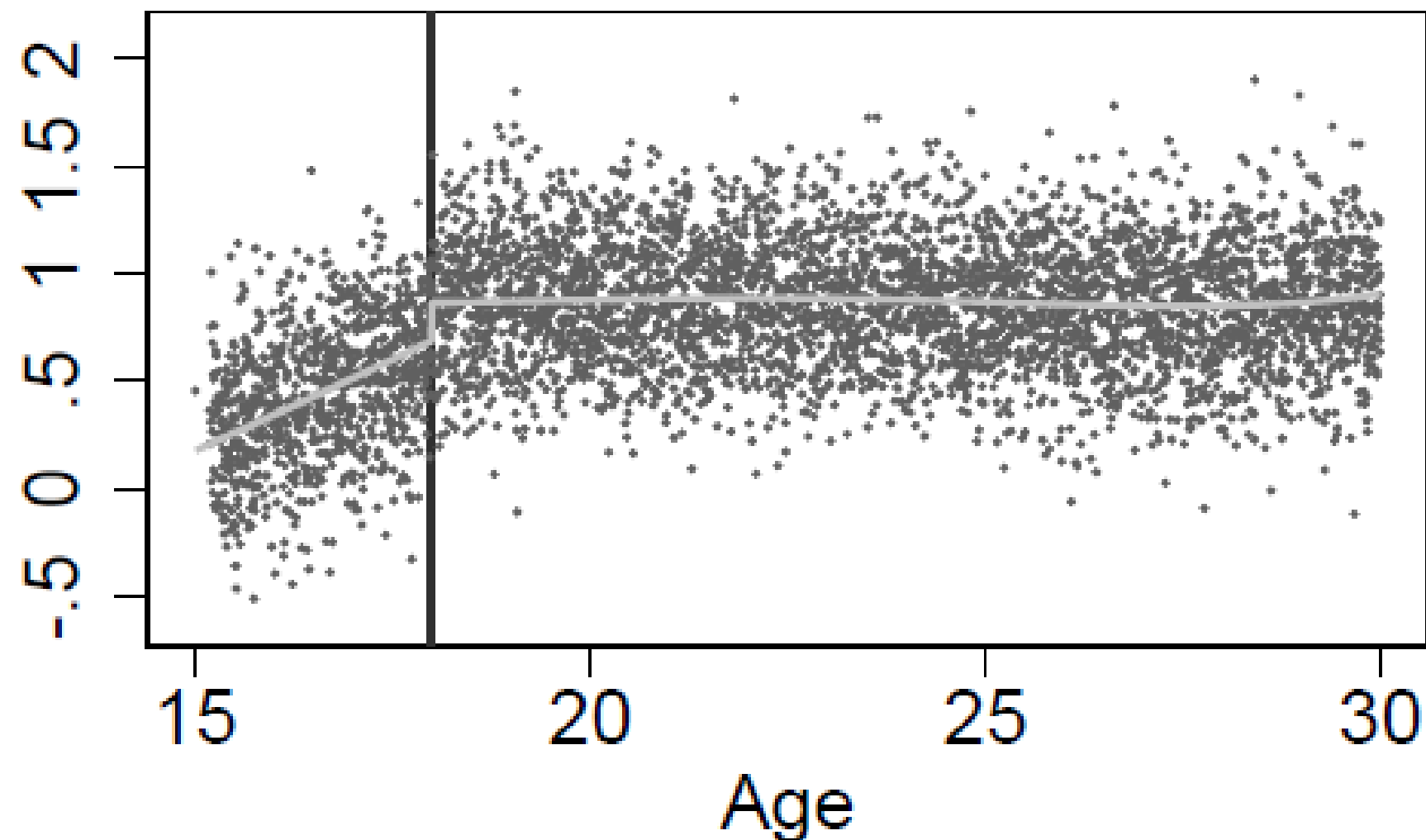
A: Head Start Simulation



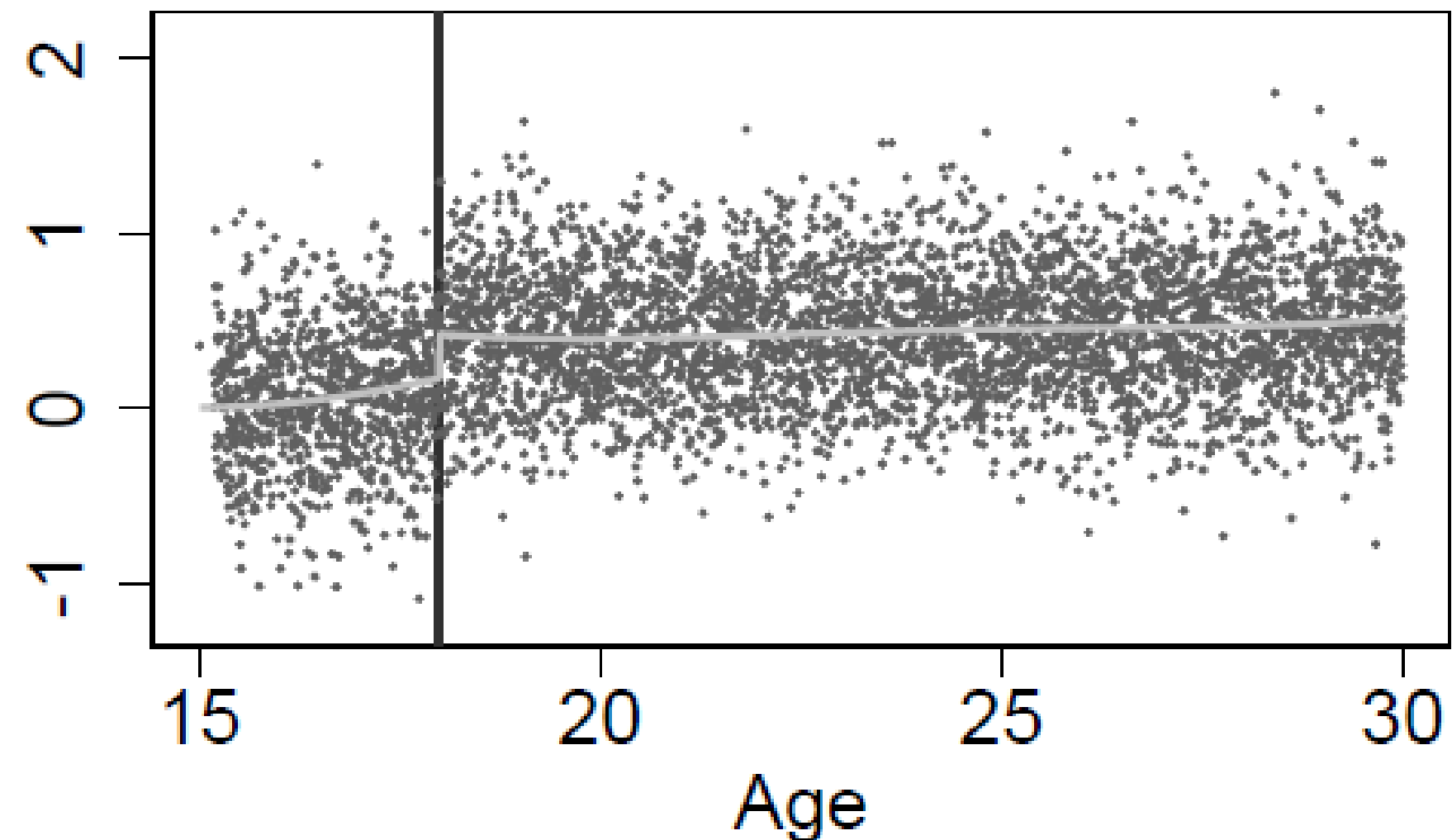
B: Incumbency Simulation



C: MLDA Sim - Ever Drinks



D: MLDA Sim - Drinks Regularly



Monte Carlo Results (4)

	KS			CCT		IK	
	RMSE	mean BW	order (%)	RMSE	mean BW	RMSE	mean BW
			F: Head Start		DGP		
Mortality	0.6794	13.86	0.999	1.4265	7.93	0.9853	14.87
			G: Political incumbency		DGP		
Wins	0.0111	21.59	0.988	0.0125	22.76	0.0119	30.35
			H: MLDA		DGP		
Ever Drinks	0.0271	3.81	0.981	0.0591	0.96	0.0347	2.50
Drinks Regularly	0.0340	3.96	0.983	0.0750	0.96	0.0425	2.89
Proportion of Days Drinks	0.0154	3.96	0.983	0.0339	0.96	0.0189	3.15

Monte Carlo Results (4)

	KS			CCT		IK	
	RMSE	mean BW	order (%)	RMSE	mean BW	RMSE	mean BW
			F: Head Start	DGP			
Mortality	0.6794	13.86	0.999	1.4265	7.93	0.9853	14.87
			G: Political incumbency	DGP			
Wins	0.0111	21.59	0.988	0.0125	22.76	0.0119	30.35
			H: MLDA	DGP			
Ever Drinks	0.0271	3.81	0.981	0.0591	0.96	0.0347	2.50
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Proportion of Days Drinks	0.0154	3.96	0.983	0.0339	0.96	0.0189	3.15

Monte Carlo Results Summary

- Our approach always beats CCT and IK using DGPs based on real data
- Our approach beats CCT in most of the stylized simulations as well, including those with highly unstable DGPs (Sine and Cosine)
 - Performance against IK more mixed (we usually win with simpler DGPs (linear, quadratic), particularly with more error variance, but usually lose with sine/cosine. However, RMSEs and selected BWs similar.

Outline

- Key Issues and our Contribution
- **Motivating application – 2 policy changes affecting Learner drivers in NSW**
- Theory – show our approach is asymptotically optimal, under restrictive conditions
- Simulations – our approach performs favourably compared to other procedures using stylised and realistic DGPs

Candidate models for our application

1. Conventional (fully-interacted) linear RDD.
2. RDD with a linear fit on the right side of the threshold, and a quadratic on the left.
3. Conventional (fully-interacted) quadratic RDD.
4. Conventional (fully-interacted) linear RPJKD.
5. Quadratic RPJKD, in which quadratic term is not interacted with the threshold indicator.
6. RPJKD model with a linear fit on the right side of the threshold, and a quadratic on the left.
7. Fully-interacted quadratic RPJKD.
8. Conventional (fully-interacted) linear RKD.
9. Quadratic RKD, in which the quadratic term is not interacted with the threshold indicator.
10. RKD with a linear fit on the right side of the threshold, and a quadratic on the left.
11. Fully-interacted quadratic RKD.
12. Month-of-birth cohort IV, with linear DOB control
13. Month-of-birth cohort IV, with quadratic DOB control
14. Month-of-birth cohort IV, with cubic DOB control

Summary of Model Performance in Placebo Zone

Model	Description	RMSE	Optimal BW	Coverage	Bias
1	RDD - linear	0.0083	365	0.962	-0.0004
2	RDD - mixed polynomial	0.0199	365	0.921	0.0014
3	RDD - quadratic	0.0230	365	0.927	0.0015
4	RPJKD - linear	0.0060	365	0.936	0.0010
5	RPJKD - quadratic	0.0073	365	0.980	-0.0005
6	RPJKD - mixed polynomial	0.0052	365	0.992	0.0000
7	RPJKD - interacted quadratic	0.0132	365	0.938	0.0005
8	RKD - linear	0.0096	355	0.910	0.0028
9	RKD - quadratic	0.0179	365	0.953	0.0019
10	RKD - mixed polynomial	0.0057	365	0.984	0.0002
11	RKD - interacted quadratic	0.0177	365	0.950	0.0019
12	birth cohort-IV - linear	0.0051	365	0.946	0.0006
13	birth cohort-IV - quadratic	0.0070	365	0.987	-0.0007
14	birth cohort-IV - cubic	0.0124	365	0.937	0.0001

Key Estimates

Table 6: Estimated effects of minimum supervised driving hours

	Best estimator	estimator	Best cohort-IV	sym. RPJKD	Best sym. RKD	sym. RDD
	(1)	(2)	(3)	(4)	(5)	
A: 2000 Reform (0 → 50 hours)						
MVA	-0.0144***	-0.0132***	-0.0147***	-0.0144**	-0.0168***	
1-year						
SE	0.0041	0.0049	0.0050	0.0058	0.0058	
p-value	0.0005	0.0073	0.0032	0.0129	0.0038	
alt. SE	0.0039	0.0047	0.0049	0.0054	0.0080	
alt. p-value	0.0096	0.0203	0.0157	0.0381	0.0552	
Model	6	12	6	10	1	
BW	365 / 550	365	365	365	365	
B: 2007 Reform (50 → 120 hours)						
MVA	0.0021	0.0003	0.0006	-0.0024	-0.0007	
1-year						
SE	0.0030	0.0033	0.0033	0.0046	0.0035	
p-value	0.4790	0.9259	0.8477	0.6069	0.8422	
alt. SE	0.0042	0.0047	0.0049	0.0054	0.0080	
alt. p-value	0.6310	0.9496	0.9003	0.6774	0.9328	
Model	12	12	6	10	1	
BW	560 / 365	365	365	365	365	

Conclusions

- We propose a new approach for model selection in RDD and related designs using placebo zone data
- Can compare across and within model types, on any number of dimensions. Also offers a new approach for inference
- We recommend its use whenever the DGP is `stable' across the range of the running variable, and the placebo zone is not small

Policy conclusions

- Going from 0 → 50 MSDH reduced MVAs. Relatively large effect sizes (21% reduction in 1st year). No effect after 1-2 years (evidence against habit formation).
- No effect from further increase to 120 MSDH.

Our 2 Stata Programs

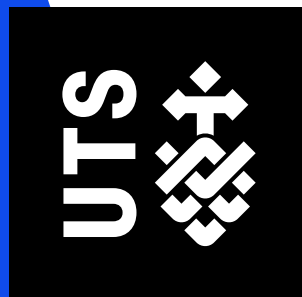
ssc install pzms

<https://econpapers.repec.org/software/bocbocode/s459073.htm>

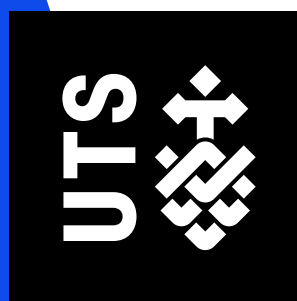
-pzms- implements our approach. Very easy to use. Only required option is the maximum bandwidth

[pzms sim](#) uses simulations based on the data from any application, to examine likely performance of our approach, compared to alternative approaches <https://sites.google.com/site/nrkettlewell/research>

Thank
you



Extra Slides



Data

- Administrative records on licence history of drivers in NSW linked with records on crashes between 1996 and 2016. Only crashes where at least one car is towed away and/or someone was injured are recorded.
- Because of important discontinuous policy changes we only consider BWs of up to 365 days for our baseline analysis.
- In main analysis $N = 154,524$ drivers born within 1 year of 1 July 1984 (2000 reform), $N = 160,301$ within 1 year of 1 July 1991 (2007 reform).

What are Researchers Doing?

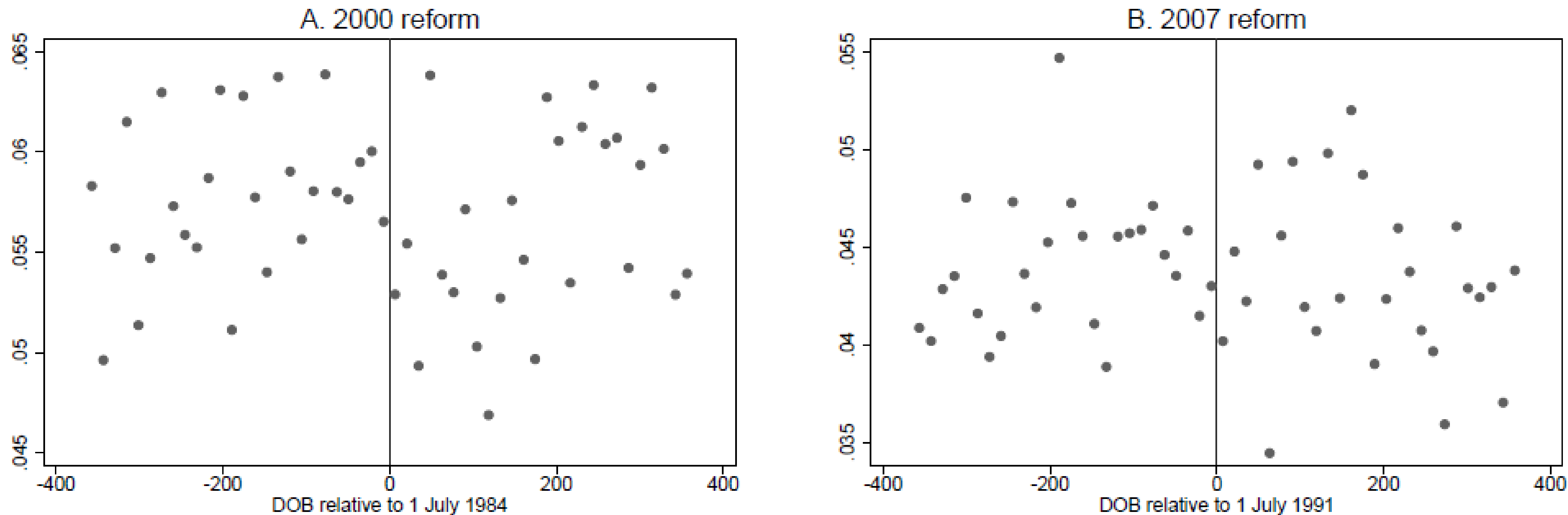
Table: Discontinuity studies published in leading journals in 2019

	Sharp RDD	Fuzzy RDD	Cohort-IV
Papers using this model	15	10	2
<u>Method for bandwidth choice</u>			
No stated method	6	4	2
IK/CCT	8	6	0
<u>Method for polynomial choice</u>			
No stated method	10	9	2
Local linear polynomial as baseline	11	8	-
<u>Robustness tests</u>			
Varied bandwidth	15	9	1
Varied polynomial	13	5	1

Our Approach

- Estimate many placebo treatment effects for each candidate estimator by moving the placebo threshold across the placebo zone where we know the treatment effect = 0.
- Choose the estimator with lowest root mean squared error (RMSE) of treatment effect estimates across all candidate models. Can also assess bias and coverage.
- Can compare models on any dimension and of different model types.
- Extendable to nonlinear models (e.g. logit)
- Combine with randomization inference similar to Ganong & Jager (2018)

Figure: Reduced-Form Relationships between DOB and MVA 1-year



Notes: Scatter plots use 14-day bin size.

Decisions to make

- Which empirical technique should we use (for estimation)?
- Which order of polynomial should we use?
- What bandwidth should we use?
- Should we use a different approach on each side of the threshold?

Monte Carlo Simulations (cont.)

- **Stylised DGPs**

- Sample size = 500 observations (following IK, CCT and others)
- x is uniformly distributed across the range $(-100, 400)$
- Outcome variable $y = 0.3(x > 0) + f(x) + e$,
- where 0.3 is the discontinuity at $x = 0$, and $\sigma^2 = 0.1^2$ (representing 'large' error variance), or 0.3^2 ('small' error variance)
- $f(x)$ is either: Linear, Quadratic, Sine, or Cosine

Proof of Asymptotic Optimality (1)

Let τ be the true treatment effect. The MSE of any estimator is

$$MSE(\hat{\tau}) = E(\hat{\tau} - \tau)^2 = Var(\hat{\tau}) + Bias(\hat{\tau})^2 \quad (1)$$

Consider local linear RD estimators with bandwidth b . Assume n_b observations within this bandwidth, on each side of the threshold.

$\hat{\tau}_b = \hat{\alpha}_{2b} - \hat{\alpha}_{1b}$, where $\hat{\alpha}_{2b}$ and $\hat{\alpha}_{1b}$ are estimated using independent linear regressions on each side of the threshold ($x = 0$)

$$y = \alpha_1 + \beta_1 x + \varepsilon, \quad -b < x < 0 \quad (2a)$$

$$y = \alpha_2 + \beta_2 x + \varepsilon, \quad 0 < x < b \quad (2b)$$

$$MSE(\hat{\tau}_b) = E(\hat{\alpha}_{2b} - \hat{\alpha}_{1b} - \tau)^2 = Var(\hat{\alpha}_{1b}) + Var(\hat{\alpha}_{2b}) + Bias(\hat{\tau}_b)^2 \quad (3)$$

Proof of Asymptotic Optimality (2)

Consider a true DGP that is cubic: $y = \alpha + \tau I(x > 0) + \theta_1 x + \theta_2 x^2 + \theta_3 x^3 + \varepsilon$, (4)

$\theta_2 x^2$ and $\theta_3 x^3$ are omitted variables from (2a) and (2b), $\alpha_1 = \alpha$ and $\alpha_2 = \alpha + \tau$. Using conventional OVB formulas:

$$E(\hat{\alpha}_{1b}) = \alpha + \hat{\delta}_{L1} + \hat{\delta}_{L2} \quad (5)$$

Where $\hat{\delta}_{L1}$ is the estimated constant in the regression of $\theta_2 x^2$ on x : $\theta_2 x^2 = \delta_1 + \pi_1 x + \epsilon_1$, $-b < x < 0$ (6A)

And $\hat{\delta}_{L2}$ is the estimated constant in the regression of $\theta_3 x^3$ on x , $\theta_3 x^3 = \delta_2 + \pi_2 x + \epsilon_2$, $-b < x < 0$ (6B)

Similarly for $E(\hat{\alpha}_{2b})$ using data on the RHS:

$$E(\hat{\alpha}_{2b}) = (\alpha + \tau) + \hat{\delta}_{R1} + \hat{\delta}_{R2} \quad (5B)$$

Proof of Asymptotic Optimality (3)

The expected value of the RD estimate is hence:

$$E(\hat{\tau}_b) = E(\hat{\alpha}_{2b}) - E(\hat{\alpha}_{1b}) = (\alpha + \tau) + \hat{\delta}_{R1} + \hat{\delta}_{R2} - (\alpha + \hat{\delta}_{L1} + \hat{\delta}_{L2}) \quad (7)$$

However, $\hat{\delta}_{R1} = \hat{\delta}_{L1}$ and $\hat{\delta}_{R2} = -\hat{\delta}_{L2}$, and so:

$$E(\hat{\tau}) = \tau + 2\hat{\delta}_{L2}$$

(to see this, replace x with $-x$ in (6A) and (6B), noting the assumed uniform distribution of x . (6A) becomes $\theta_2 x^2 = \delta_1 - \pi_1 x + \epsilon_1$. This regression yields exactly the same estimate of δ_1 . (6B) becomes $-\theta_3 x^3 = \delta_2 - \pi_2 x + \epsilon_2$. This regression yields an estimated constant exactly equal to $-\hat{\delta}_{L2}$.)

(6B) implies that $\hat{\delta}_{L2}$ is proportional to θ_3 , and unrelated to any other parameters of the cubic DGP. The bias of $\hat{\tau}_b$ is hence proportional to the third derivative of the DGP CEF.

Proof of Asymptotic Optimality (4)

The discontinuity estimates at any placebo threshold at $x = k$, have the same bias, assuming the same global DGP, for any $|k| > b$

The linear RD estimate with bandwidth b at a placebo discontinuity at $x = k$ is $\hat{\tau}_{kb} = \hat{\alpha}_{k2b} - \hat{\alpha}_{k1b}$, where $\hat{\alpha}_{k1b}$ and $\hat{\alpha}_{k2b}$ are the estimates from these regressions:

$$y = \alpha_{k1} + \beta_1(x - k) + \varepsilon, \quad (k - b) < x < k \quad \text{and} \quad y = \alpha_{k2} + \beta_2(x - k) + \varepsilon, \quad k < x < (k + b)$$

Substituting $x_k = x - k$, these regressions are equivalent to

$$y = \alpha_{k1} + \beta_1 x_k + \varepsilon, \quad -b < x_k < 0 \quad (9a), \quad \text{and}$$

$$y = \alpha_{k2} + \beta_2 x_k + \varepsilon, \quad 0 < x_k < b \quad (9b)$$

The DGP can be expressed as $y = \alpha + \theta_1(x_k + k) + \theta_2(x_k + k)^2 + \theta_3(x_k + k)^3 + \varepsilon$ (10)

if $k > b$, and similarly if $k < -b$

Proof of Asymptotic Optimality (5)

$$\text{Equivalently, } y = \pi_0 + \pi_1 x_k + \pi_2 x_k^2 + \theta_3 x_k^3, \quad (11)$$

Where $\pi_0 = \alpha + k\theta_1 + k^2\theta_2 + k^3\theta_3$, $\pi_1 = \theta_1 + 2k\theta_2 + 3k^2\theta_3$, $\pi_2 = \theta_2 + 3k\theta_3$

(9a), (9b) and (11) are equivalent to equations (2a), (2b) and (4), respectively, with $\tau = 0$ and the threshold at $x_k = 0$. As shown, the bias of the RDD estimate is proportional only to the third derivative of the true DGP's CEF. The third derivative ($6\theta_3$) is the same in (11) as in (4), and so $Bias(\hat{\tau}_{kb}) = Bias(\hat{\tau}_b)$, for $|k| > b$.

It is trivial to show that $Var(\hat{\tau}_{kb}) = Var(\hat{\tau}_b)$. Therefore $MSE(\hat{\tau}_{kb}) = MSE(\hat{\tau}_b)$.

Proof of Asymptotic Optimality (6)

For any given estimator, as the placebo zone gets large, the mean of the squared placebo estimates approaches the MSE of the treatment effect estimator:

$$\text{Since } \tau_{kb} = 0, \quad \lim_{m \rightarrow \infty} \frac{1}{m} \sum_{k=b+1}^{b+m} \hat{t}_{kb}^2 = \text{MSE}(\hat{t}_b)$$

If the DGP has a non-zero fourth derivative, our approach is no longer asymptotically optimal. To see this, assume a fourth-order polynomial DGP and follow the same steps. The equivalent of equation (11) would have a different coefficient of x_k^3 for each k . Therefore the bias of placebo estimates would be different at each placebo threshold.

Monte Carlo Results (1)

	KS		CCT		IK		
	RMSE	mean	linear	RMSE	mean	RMSE	mean
		BW	(%)		BW		BW
A: Linear DGP							
Baseline DGP	0.0241	250.19	1.000	0.0413	66.42	0.0295	149.85
Small error variance	0.0072	250.19	1.000	0.0092	83.01	0.0079	201.92
Small placebo zone	0.0292	151.53	0.974	0.0601	34.70	0.0319	120.94
Small placebo zone and error variance	0.0088	151.53	0.974	0.0180	34.69	0.0087	149.19
B: Quadratic DGP							
Baseline DGP	0.0285	114.30	0.999	0.0413	66.37	0.0331	137.67
Small error variance	0.0088	98.63	0.995	0.0124	65.87	0.0166	149.54
Small placebo zone	0.0318	118.92	0.969	0.0601	34.70	0.0332	115.63
Small placebo zone and error variance	0.0096	98.54	0.953	0.0180	34.70	0.0116	117.95

Monte Carlo Results (2)

	KS			CCT		IK	
	RMSE	mean BW	order (%)	RMSE	mean BW	RMSE	mean BW
	C: Cubic DGP						
Baseline DGP	0.0301	93.59	1.000	0.0414	66.38	0.0325	114.81
Small error variance	0.0104	79.49	1.000	0.0125	65.94	0.0109	89.25
Small placebo zone	0.0321	100.39	0.953	0.0601	34.69	0.0340	106.87
Small placebo zone and error variance	0.0110	82.66	0.919	0.0180	34.66	0.0111	83.62
	D: Sine DGP						
Baseline DGP	0.0481	59.0	0.857	0.0605	34.1	0.0443	60.8
Small error variance	0.0172	41.6	0.886	0.0187	30.8	0.0174	56.5
Small placebo zone	0.0633	79.9	0.729	0.0605	34.1	0.0443	60.8
Small placebo zone and error variance	0.0207	58.6	0.673	0.0187	30.8	0.0174	56.5

Monte Carlo Results (3)

	KS		CCT		IK		
	RMSE	mean BW	order (%)	RMSE	mean BW	RMSE	mean BW
E: Cosine DGP							
Baseline DGP	0.0446	83.9	0.640	0.0601	34.7	0.0377	75.1
Small error variance	0.0147	42.1	0.876	0.0180	34.7	0.0114	71.0
Small placebo zone	0.0437	82.5	0.579	0.0601	34.7	0.0377	75.1
Small placebo zone and error variance	0.0153	41.8	0.778	0.0180	34.7	0.0114	71.0

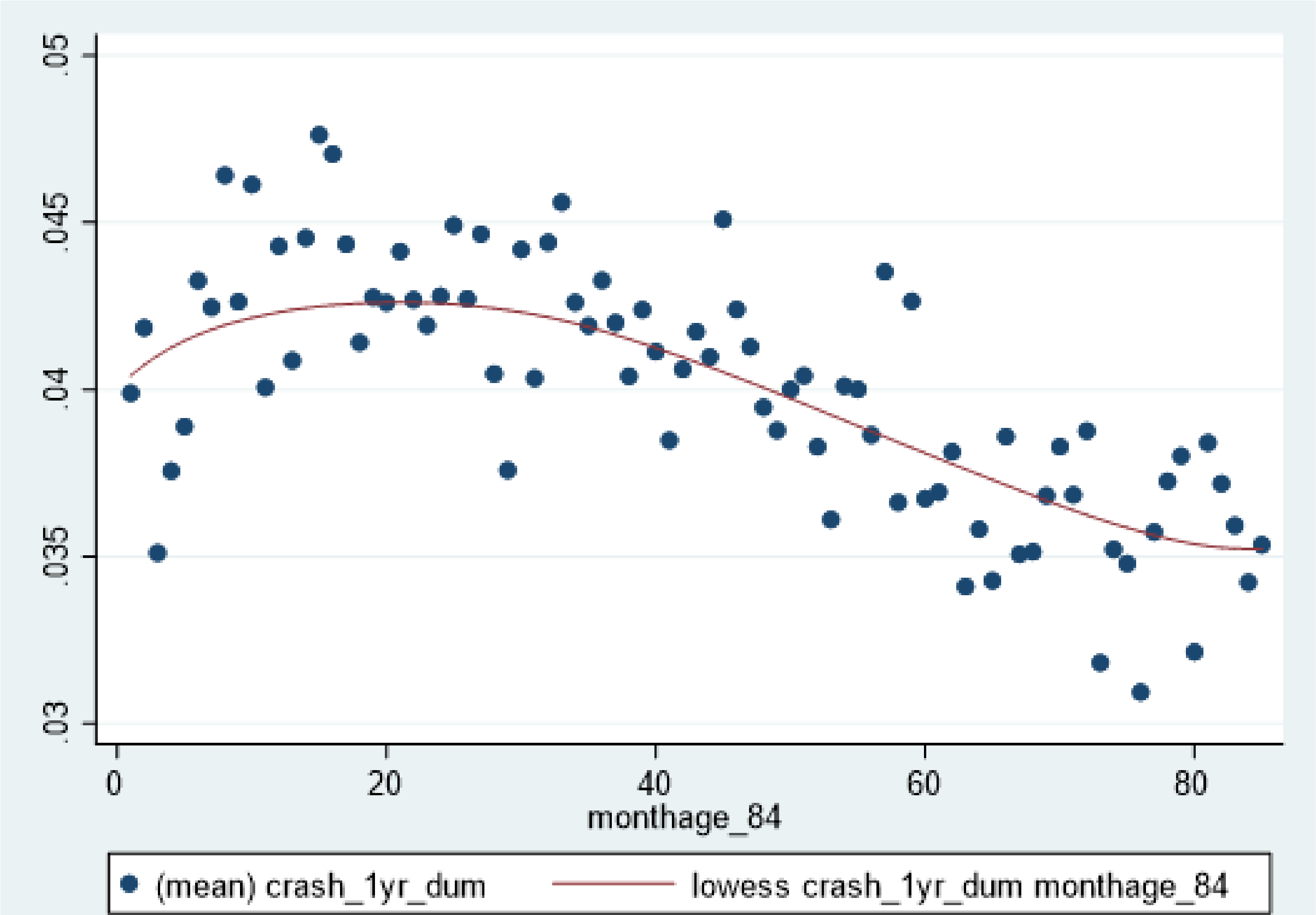
Our procedure

- There are total of $14 \times 331 = 4,634$ candidate models – i.e. 14 models with BW ranging from 35 to 365 days.
- We estimate the placebo treatment effect (which we know to be zero and constant across entities) using each candidate model. We repeat this for all **1826** placebo treatment thresholds, and assess the performance of each candidate model.
- Key stat is the **Root Mean Squared Error (RMSE)** of the estimated treatment effect = square root of the sum of the 1826 estimates
- Other stats are Coverage Rate (% of estimated CIs which include the true effect)

Extensions

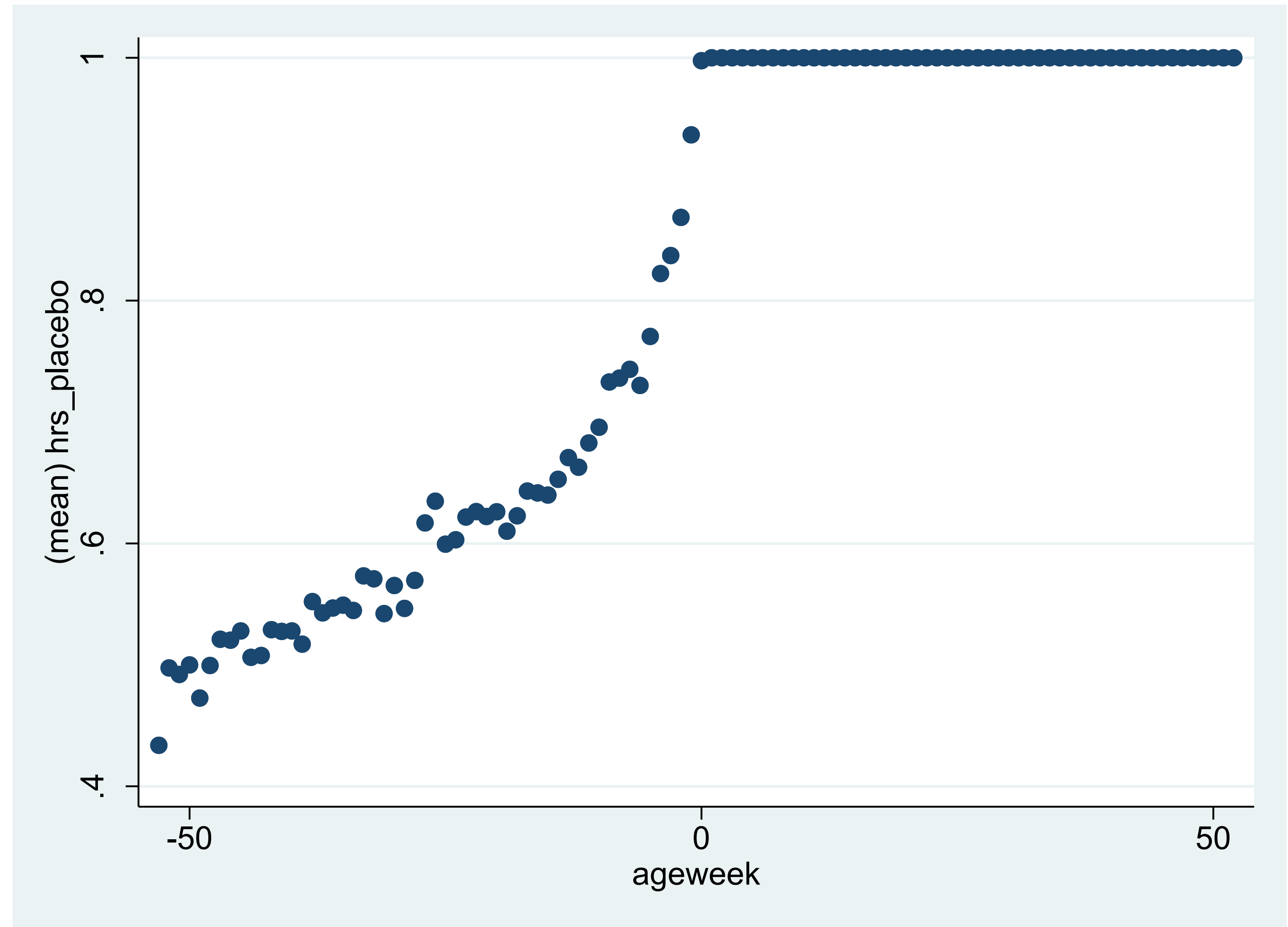
- Allow asymmetric bandwidths (to the left and right of the threshold) – these actually do better!
- *Impose* treatment effect heterogeneity into the placebo zone
 - Random perturbation of the real data for ‘treated’ observations
 - Marginal Treatment Effect which is linear in ‘resistance’

Outcome data in our placebo zone



Placebo treatments

Within this zone, we create placebo treatments in a way which mimics the true treatment selection process. For example, in the first placebo, persons are deemed treated if they obtained their license on or after 1 July 2001. The first stage relationship is shown.



Timing of Treatment Effect

		Best esti- mator	Best sym. cohort-IV	Best sym. RPJKD	Best sym. RKD	Best sym. RDD
B: Timing of Treatment Effect						
MVA	6	-0.0094***	-0.0074**	-0.0078**	-0.0072	-0.0093**
months						
SE		0.0031	0.0035	0.0036	0.0044	0.0044
MVA	6-12	-0.0048*	-0.0059*	-0.0070**	-0.0069*	-0.0077*
months						
SE		0.0028	0.0034	0.0034	0.0040	0.0040
MVA	1-2	0.0035	0.0026	0.0027	0.0019	0.0033
years						
SE		0.0036	0.0045	0.0046	0.0053	0.0053

Seriousness and Heterogeneity

	Best esti- mator	Best sym. cohort-IV	Best sym. RPJKD	Best sym. RKD	Best sym. RDD
C: Serious MVAs					
Injury	-0.0084***	-0.0093***	-0.0100***	-0.0102***	-0.0110***
SE	0.0026	0.0032	0.0032	0.0038	0.0036
Fatality	-0.0002	-0.0001	-0.0002	-0.0002	-0.0002
SE	0.0003	0.0004	0.0004	0.0005	0.0005
D: Heterogeneity by Sex					
MVA 1-year males	-0.0132**	-0.0146**	-0.0139*	-0.0114	-0.0163*
SE	0.0059	0.0072	0.0073	0.0086	0.0085
MVA 1-year females	-0.0164***	-0.0111*	-0.0159**	-0.0181**	-0.0177**
SE	0.0056	0.0066	0.0068	0.0080	0.0083

Other Applications

- We use our method to re-evaluate evidence on Head Start (Ludwig & Miller, 2007) and minimum legal drinking age (Lindo et al, 2016).
- In both cases, our method selects much larger BWs than the original studies.
- However, the conclusions of those studies are unchanged.